



Research article

Repeatability of diffusion-weighted MRI of the prostate using whole lesion ADC values, skew and histogram analysis



Tristan Barrett^{a,b,c,*,1}, Edward M. Lawrence^{a,d,1,2}, Andrew N. Priest^b, Anne Y. Warren^{c,e}, Vincent J. Gnanapragasam^{c,f}, Ferdia A. Gallagher^{a,b}, Evis Sala^{a,b}

^a Department of Radiology, University of Cambridge, Cambridge, UK

^b Department of Radiology, Addenbrooke's Hospital, Cambridge, UK

^c CamPARI Clinic, Addenbrooke's Hospital and University of Cambridge, Cambridge, UK

^d Department of Radiology, Memorial Sloan Kettering Cancer Center, 1275 York Ave, New York, NY, United States of America

^e Department of Histopathology, Addenbrooke's Hospital and University of Cambridge, Cambridge, UK

^f Department of Urology, Addenbrooke's Hospital and University of Cambridge, Cambridge, UK

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ABSTRACT

Purpose: To investigate the repeatability of diffusion-weighted imaging parameter including ADC-derived histogram values in prostate cancer.

Methods: 10 patients with prostate cancer were prospectively recruited to a retest cohort. 3 T diffusion-weighted MRI of the prostate was acquired consecutively with patient getting off the scanner between studies. Prostatectomy-histopathology defined tumour regions-of-interest were outlined on ADC maps and diffusion-weighted metrics including histograms were calculated. The coefficient of reproducibility (CoR) and Bland-Altman plots were used to assess repeatability.

Results: 10th centile, 90th centile, and median ADC showed good repeatability with mean difference ranging from -0.005 to $-0.025 \times 10^3 \text{ mm}^2\text{s}^{-1}$, and CoR ranging from 0.271–0.294 $\times 10^3 \text{ mm}^2\text{s}^{-1}$ of scan 1 mean). Two measures of heterogeneity and simplified texture, IQR and mean local range, had only moderate repeatability. IQR had a mean difference of $-0.032 \times 10^3 \text{ mm}^2\text{s}^{-1}$ between scans with CoR 0.181 $\times 10^3 \text{ mm}^2\text{s}^{-1}$ (56% of scan 1 mean). Mean local range had a mean difference $-0.008 \times 10^3 \text{ mm}^2\text{s}^{-1}$ between scans (37% of scan 1 mean). Bland-Altman plots showed good repeatability for test and re-test analysis for median, percentile and mean range values. All ADC values had good reliability regardless of whether the tumour border was included in quantitative analysis. ADC histogram skew had poor repeatability, CoR 0.78 $\times 10^3 \text{ mm}^2\text{s}^{-1}$ (373% of scan 1 mean).

Conclusion: 10th and 90th centile ADC demonstrated sufficient repeatability for clinical use. However, more advanced measures of heterogeneity such as histogram skew, IQR, or mean local range may be limited by their repeatability.

1. Introduction

Prostate cancer is the commonest cancer in men, with the incidence expected to double by 2030 mainly due to the ageing population [1,2]. The traditional work-up of prostate cancer with transrectal ultrasound (TRUS) biopsy is limited by random and systematic errors in sampling [3,4]. However, this practice is beginning to change, driven by multiparametric (mp) MRI which offers the potential to overcome many of these disadvantages [5,6].

The improved ability of mpMRI to detect lesions has mainly been due to the addition of functional sequences such as diffusion-weighted imaging (DWI), and dynamic contrast-enhanced (DCE) MRI. DWI images the diffusion of water molecules and provides information related to tumour cellularity and tissue composition. DWI-derived apparent diffusion coefficient (ADC) maps can provide a quantitative measure of the degree of restricted diffusion, with a number of studies showing that these values correlate inversely to Gleason grade [7–9]. Whilst potentially attractive as a surrogate marker of tumour

* Corresponding author at: Department of Radiology, Addenbrooke's Hospital and University of Cambridge, Cambridge, CB2 0QQ, UK.

E-mail address: tb507@medschl.cam.ac.uk (T. Barrett).

¹ These authors contributed equally to this work.

² Current affiliation: Department of Radiology, Memorial Sloan Kettering Cancer Center, 1275 York Ave, New York, NY, United States of America

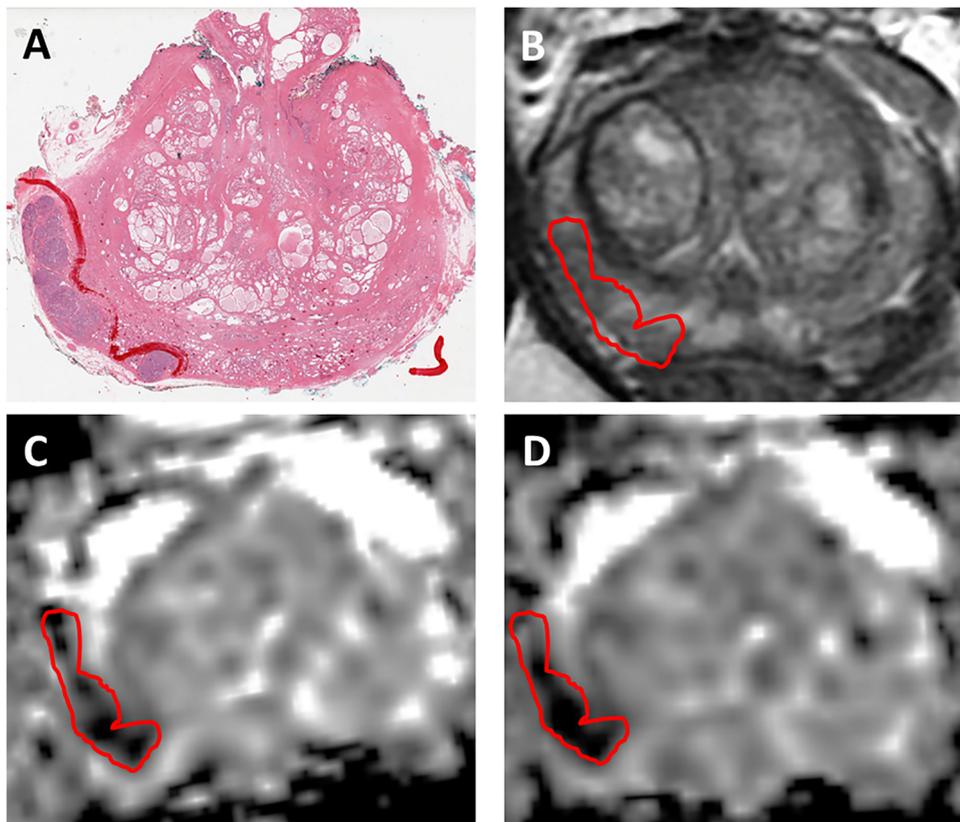


Fig. 1. ROI Outlining.

64 yr-old, PSA 6.7 ng/ml. A: Right base PZ Gleason 4 + 5 tumour outlined on histopathology slide. B: matching tumour ROI outlined on T2 axial image, and transposed to initial (C) and repeat ADC maps (D).

aggressiveness, absolute ADC values can vary depending on the choice and number of b values selected and thus current guidelines caution against the use of quantitative ADC measurements [10–12]. Another potential source of error is the reproducibility of the test itself, which may cause particular problems in the assessment of response to treatment and determining meaningful change in patients on active surveillance.

ADC values have shown reasonable retest reproducibility in the body, with a variation of around 20% [13–15]. However, these evaluations have been primarily focused on an assessment of the mean or median ADC values. Similar studies have been performed in the prostate looking at the reproducibility of mean ADC values in the same imaging session [16] or within 2 weeks [17], with reproducibility variation ranging from 10–40%. ADC histogram-derived values have shown promising results in detecting and characterizing disease along with evaluating treatment response [18–21]. Histogram-derived values attempt to assess the spatial variation of ADC values and may provide additional information on tumour heterogeneity, which may be increased within tumours, rather than averaging out these differences within a region-of-interest. Indeed, the current version of the Prostate Imaging Reporting and Data System (PI-RADS 2) guidelines strongly supports the continued development of further novel MRI sequences and analysis methods such as multiple b-value assessment of fractional ADC, and measures of cellular level heterogeneity such as diffusion kurtosis imaging [10,22]. However, for either quantitative ADC metrics or histogram analysis to be used clinically, their repeatability and reliability needs to be evaluated and quantified. Therefore, the purpose of this study was to evaluate, in the setting of prostate cancer, the reliability of ADC histogram-derived parameters and simplified textural analysis for selected regions of interest.

2. Methods and materials

The local institutional review board and ethics committee granted approval for this prospective study, with all participants signing written informed consent. 11 patients with biopsy-proven intermediate or high risk prostate cancer underwent a dedicated research prostate MRI scan prior to treatment with radical prostatectomy. Exclusion criteria included previous treatment for prostate cancer, or clinical contraindication to MRI. One patient was excluded due to significant susceptibility artefact on DW imaging, with 10 patients completing the study.

2.1. MR imaging technique

All patients underwent 3-T MRI (Signa HDx, GE-Healthcare, WI, USA) using an 8-channel cardiac phased-array coil. The protocol included axial T1-weighted images of the pelvis and T2-weighted images in axial, sagittal and coronal planes. Multislice diffusion-weighted (DW)-MRI was performed (TE/TR = 78/4400 ms; FOV $30 \times 30 \text{ cm}^2$; acquisition matrix 128×128 (reconstruction matrix 256×256); parallel imaging (ASSET) factor = 2; 8 signal averages; b-values: 150,1000s/mm²) using a dual-spin-echo (DSE) EPI acquisition sequence to minimise eddy-current-related distortion. Slice thickness for the axial T2-weighted sequences was 3.5 mm, with 0.5 mm gap; the axial DW imaging was matched using a 4 mm slice thickness, with 0 mm gap. Water-selective excitation was used for fat saturation. An identical DWI acquisition was performed twice with the patient getting off the scanner bed in between the scans. Slice locations were matched to the initial scan using anatomical landmarks. ADC maps were generated using in-house software (b-values- 150, 1000s/mm²) programmed with

Matlab (MathWorks, Natick, Mass).

2.2. Pathologic assessment and comparison with MR images

Following surgery each ex-vivo prostate was measured in three dimensions, oriented by the location of the seminal vesicles, flat posterior surface, and by the position of the urethra. The apical end and basal cone were amputated, sliced left-right into 3–4 mm thick pieces, and placed in small cassettes preserving their order. The remaining gland was sliced into 5 mm sections in the horizontal plane from inferior-superior. Slices from each section were annotated by an experienced specialist uro-pathologist (XXAYW).

2.3. Image analysis

Regions of interest (ROIs) were drawn on the ADC maps by two authors (EML, TB) in consensus within the same session, for each patient, with reference to T2-weighted MRI and whole-mount histopathology annotated for index lesion tumour location (Fig. 1). ROIs drawn on the first acquisition were loaded onto the matching ADC maps. The authors then reviewed all ROI transfer results and manually adjusted the ROI locations when necessary. ADC values were recorded from each pixel with the ROI and after combination of values from all relevant slices into a volume of interest. An ADC histogram was generated using in-house software programmed in Matlab (Fig. 2), with the following parameters derived: (a) median ADC; (b) 10th and 90th centiles (defined as the ADC value below which the corresponding percentage of all ADC voxel values lie); (c) interquartile range (IQR; a measure of the spread of the distribution) and (d) skewness (a measure of the asymmetry of ADC distribution).

2.4. ROI erosion

ROI erosion was performed to assess whether partial volume will affect the measured results, and may be of particular relevance when transposing an identical ROI between separately acquired studies. The morphological operation of erosion reduces the shape of an ROI by removing voxels on the image boundary. The original ROI is eroded by a 3×3 circle structuring neighbourhood resulting in a smaller ROI after removal of the border voxels whose neighbourhood is not entirely within the ROI (Supplemental Fig. 1). For the histogram-derived quantitative parameters an additional set of central ROIs were obtained, in Matlab, by using this automated image erosion tool to remove the border voxels.

2.5. Local neighbourhood ROI filtering

Local range filtering, a texture filtering function available for Matlab, was used to assess local variability (Supplemental Fig. 2), to produce the “mean range”, a simplified texture feature. Two parameters are set: the neighbourhood to be considered, and the filter to be used. Each pixel is filtered using a 3×3 structuring neighbourhood, with the range filter calculating the difference between the maximum and minimum local values and the standard deviation filter calculating the standard deviation of the local neighbourhood. The generated ROI is also eroded after filtering and using the above-described method in order to avoid inclusion of voxels outside the ROI in subsequent analysis. After filtering of the entire MR image the mean value for each ROI was determined.

2.6. Statistical analysis

In order to evaluate the short-term repeatability of the quantitative DWI parameters the difference between the two baseline mean quantitative measurements (d) for the cancerous regions for each patient were determined along with the mean difference for the study cohort.

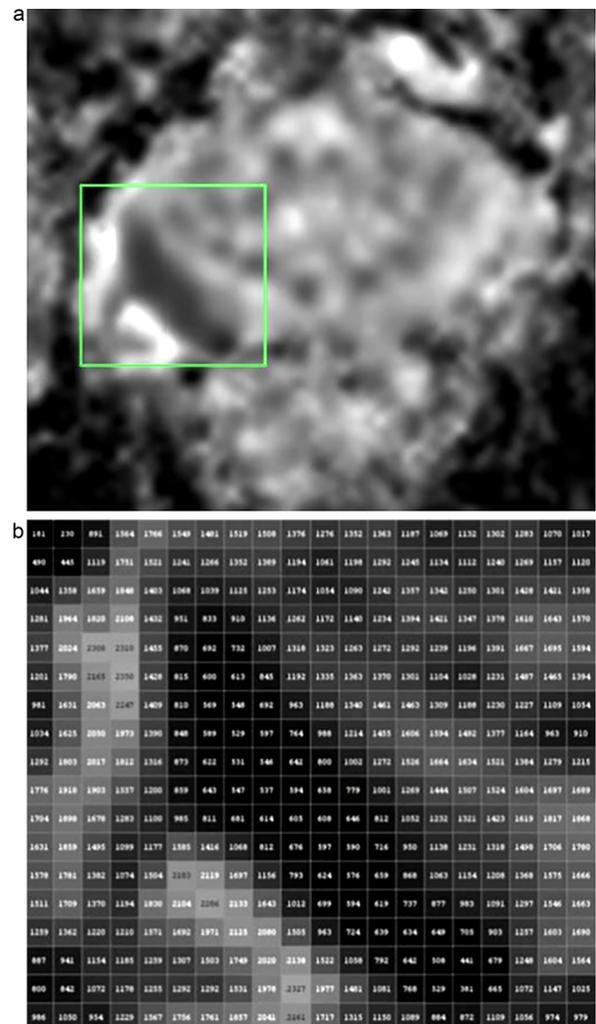


Fig. 2. ADC voxelwise data analysis.

(A) ADC map ($b = 150$, $1000\text{s}/\text{mm}^2$) shows a large right PZ tumour. The FOV was $300\text{ mm} \times 300\text{ mm}$ with a matrix of 256×256 . Thus each voxel has in-plane dimensions of $1.17 \times 1.17\text{ mm}$ and an area of 1.37 mm^2 . (B) A magnified image grid with dimensions of 20×20 voxels, corresponding to the square in A, and displaying the calculated ADC values ($\times 10^{-3}\text{ mm}^2\text{s}^{-1}$) for each individual voxel.

The mean squared difference (msd) was calculated as:

$$\text{msd} = \sqrt{\frac{\sum d^2}{n-1}} \quad (1)$$

The 95% confidence intervals (CI) for changes in the study cohort were determined using

$$\text{CI} = \pm 2.228 \times \text{msd}/n \quad (2)$$

The coefficient of repeatability calculates the maximum expected absolute difference that may be observed between any future measurements on 95% of occasions. The coefficient of repeatability (CoR) was determined using the equation:

$$\text{CoR} = 1.96 \times \text{msd} \quad (3)$$

Mean difference and 95% CI were also calculated to analyse the reliability of quantitative values depending on inclusion of the border voxels using the full ROI values and central ROI values from scan 1. All statistical analyses were performed in Matlab. Repeatability was also assessed on a lesion-by-lesion basis by the use of a Bland-Altman analysis [23].

Table 1

Repeatability for ADC heterogeneity and texture parameters using the full tumour region. *Data in units, (as % of scan 1 mean, in parentheses). SD = Standard deviation; CI = Confidence interval; IQR = Interquartile range.

	Scan 1, Mean (± SD)	Scan 2, Mean (± SD)	Mean Difference (Group 95% CI)	Coefficient of repeatability* (%)
10th centile (x 10 ⁻³ mm ² s ⁻¹)	0.839 (± 0.263)	0.834 (± 0.221)	-0.005 (± 0.106)	0.294 (35%)
Median (x 10 ⁻³ mm ² s ⁻¹)	1.086 (± 0.267)	1.078 (± 0.222)	-0.008 (± 0.098)	0.271 (25%)
90th centile (x 10 ⁻³ mm ² s ⁻¹)	1.390 (± 0.255)	1.365 (± 0.222)	-0.025 (± 0.105)	0.291 (21%)
IQR (x 10 ⁻³ mm ² s ⁻¹)	0.320 (± 0.115)	0.288 (± 0.081)	-0.032 (± 0.065)	0.181 (56%)
Skew	0.21 (± 0.33)	0.29 (± 0.40)	0.08 (± 0.28)	0.78 (373%)
Mean Range (x 10 ⁻³ mm ² s ⁻¹)	0.302 (± 0.086)	0.294 (± 0.062)	-0.009 (± 0.004)	0.112 (37%)

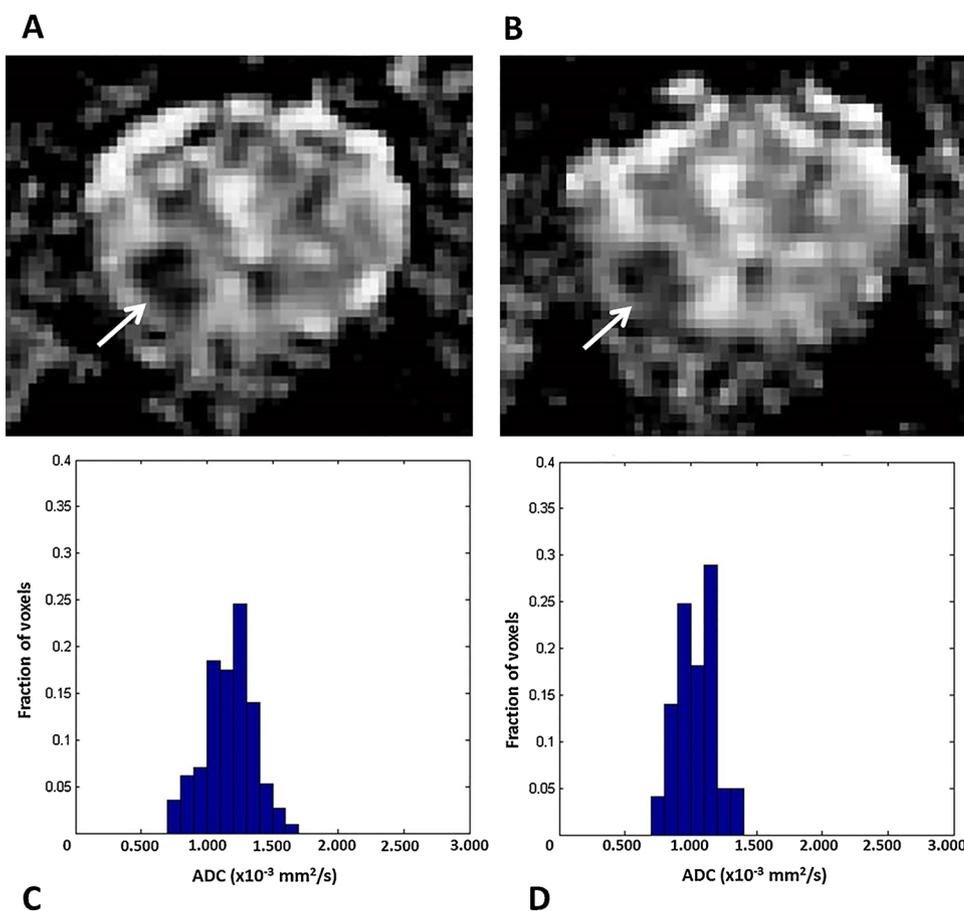


Fig. 3. Retest comparison of ADC maps and histograms in a 66-year-old man with right posterior peripheral zone prostate cancer.

(A) Study 1 and (B) study 2 ADC maps with anterior prostate cancer ROI (arrows). (C) Study 1 and (D) study 2 ADC histograms. The median ADC for study was $1.186 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ (study 2 $1.032 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$), 10th centile $0.913 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ ($0.849 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$), 90th centile $1.381 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ ($1.199 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$), IQR was $0.241 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ ($0.223 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$), and the skew -0.15 for study 1 and +0.12 for study 2. The mean local range, which was not derived from the histogram, was $0.350 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ for study 1 and $0.248 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ for study 2.

Table 2

Reliability for ADC heterogeneity parameters depending on tumour region evaluated. SD = Standard deviation; CI = Confidence interval; IQR = Interquartile range.

	Full Region, Mean (± SD)	Central Region, Mean (± SD)	Mean Difference (Group 95% CI)
10th centile (x 10 ⁻³ mm ² s ⁻¹)	0.839 (± 0.263)	0.858 (± 0.287)	0.019 (± 0.021)
Median (x 10 ⁻³ mm ² s ⁻¹)	1.086 (± 0.267)	1.046 (± 0.302)	-0.040 (± 0.042)
90th centile (x 10 ⁻³ mm ² s ⁻¹)	1.390 (± 0.255)	1.306 (± 0.303)	-0.084 (± 0.068)
IQR (x 10 ⁻³ mm ² s ⁻¹)	0.320 (± 0.115)	0.252 (± 0.110)	-0.068 (± 0.050)
Skew	0.21 (± 0.33)	0.21 (± 0.48)	-0.0001 (± 0.18)

3. Results

Ten imaged patients had a median age of 64 years (range 42–73 years), a median PSA level of 10.3 ng/mL (range 6.1–11.7 ng/mL), and Gleason grade of 3 + 3 (n = 1), 3 + 4 (n = 5), 4 + 3 (n = 3), and 4 + 5 (n = 1). Seven index lesions were located in the peripheral zone (PZ), 2 in the transition zone (TZ) and 1 tumour in both the PZ and TZ. The mean longest axial tumour diameter was 1.55 cm (median 1.45,

range 0.5–3 cm) and the mean tumour volume was 1.73 cm³ (median 0.67, range 0.05–6.92 cm³).

The results of repeatability assessment for ADC 10th centile, median, 90th centile, IQR, skew, kurtosis, and mean local range are summarized in Table 1 for the full tumour ROIs and the ADC maps and histograms for one patient are presented in Fig. 3. 10th centile, median, and 90th centile ADC were found to exhibit reasonable repeatability. They all had a small mean difference between the two scans, with the

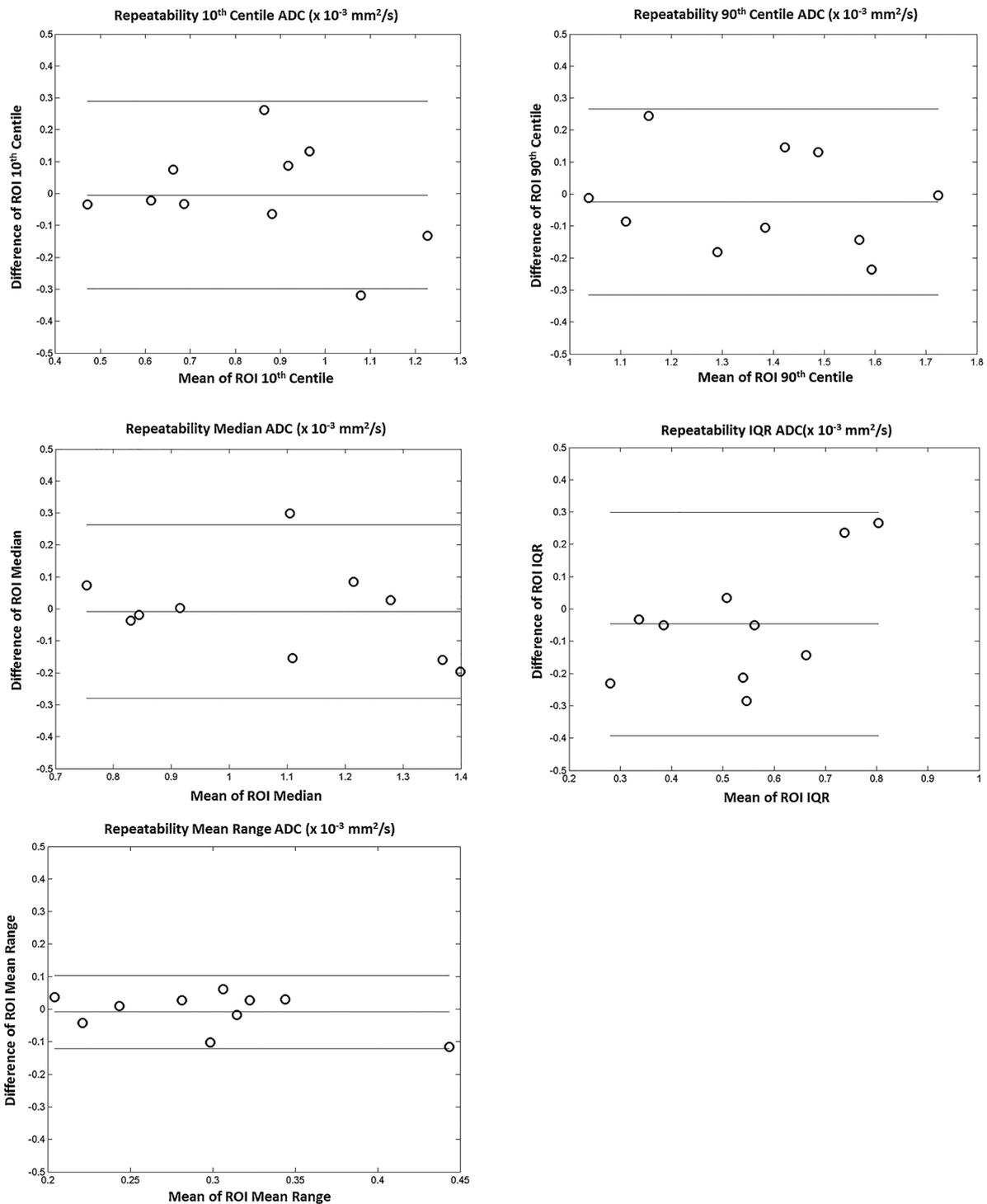


Fig. 4. Bland–Altman plots summarizing repeatability of the measurements of 10th and 90th centiles, median, interquartile range (IQR), and mean range values. The difference between the test and re-test study on the y-axis and the mean parameter value on the x-axis.

mean difference ranging from -0.005 to $-0.025 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$, and the coefficient of repeatability from $0.271 - 0.294 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$. The skew of the ADC histogram was found to have poor repeatability, with the coefficient of repeatability only $0.78 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$.

IQR of the ADC histogram had a group mean of $0.32 \pm 0.115 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ for scan 1 and a mean difference (95% CI) of $-0.032 (\pm 0.065) \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ between the two consecutive scans. The coefficient of repeatability was $0.181 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$, which is 56% of

the scan 1 mean. The mean local range of ADC had a group mean of $0.302 \pm 0.086 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$ for scan 1 and a mean difference (95% CI) of $-0.008 (\pm 0.04 \times 10^{-3} \text{ mm}^2\text{s}^{-1})$ between the two consecutive scans. The coefficient of repeatability was $0.112 \times 10^{-3} \text{ mm}^2\text{s}^{-1}$, which is 37% of the scan 1 mean.

When the values obtained from scan 1 for the full tumour ROIs and the central tumour ROIs are compared only a minimal difference is found (Table 2). 10th centile ADC had a mean difference of

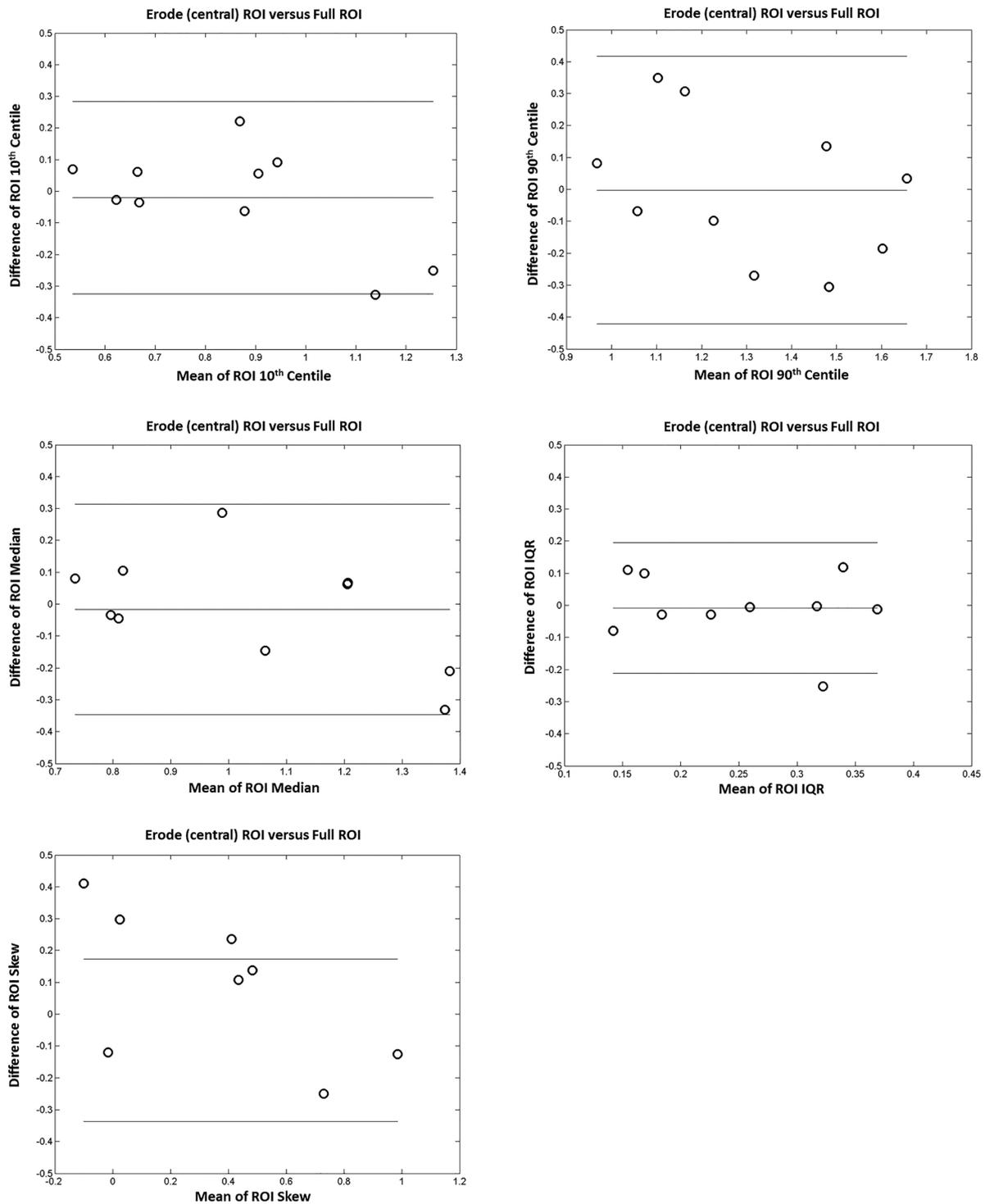


Fig. 5. Bland–Altman difference plots comparing the distribution of erode regions-of-interest (ROIs) to full ROIs for the various measured parameters. The difference between the test and re-test study on the y-axis and the mean parameter value on the x-axis.

+0.019 × 10⁻³ mm²s⁻¹ when the border was excluded while median, 90th centile and IQR had mean differences of -0.04 × 10⁻³ mm²s⁻¹, -0.084 × 10⁻³ mm²s⁻¹, and -0.068 × 10⁻³ mm²s⁻¹, respectively. The mean difference for ADC skew between the full and central ROIs was very small at -0.0001, but it had a large group 95% CI of 0.18 due to the large standard deviation of the values.

Bland–Altman plots showed good repeatability for test and re-test analysis for median, percentile and mean range values (Fig. 4) and in comparing eroded to full ROIs for all variables (Fig. 5), with the majority of points lying within the two error bars (1.96 standard

deviations).

4. Discussion

Diffusion-weighted MRI has been proposed as a non-invasive tool for prostate cancer characterization and treatment response assessment [24], with some researchers proposing that the adoption of quantitative ADC measurements can help improve the PI-RADS scoring system [25,26]. However, for a quantitative parameter to be used as a reliable measurement tool it must demonstrate adequate repeatability. ADC

mean and median values have shown good repeatability in previous studies [14,27–29], however, the reliability of ADC histogram-derived values along with texture analysis of ADC maps is largely untested. In this study we quantified the repeatability of different ADC-derived metrics, including histogram analyses in tumour regions-of-interest, demonstrating reasonable short-term repeatability for median, 10th and 90th centile ADC values.

In our retest cohort the coefficient of repeatability was $0.294 \times 10^3 \text{ mm}^2\text{s}^{-1}$ and $0.291 \times 10^3 \text{ mm}^2\text{s}^{-1}$ for the 10th and 90th centile respectively, which is the maximum absolute difference that would be expected between any two future measurements on 95% of occasions. A study of ADC centile values repeatability in 18 patients with colorectal liver metastases reported a coefficient of repeatability of $0.260 \times 10^3 \text{ mm}^2\text{s}^{-1}$ for 10th centile and $0.280 \times 10^3 \text{ mm}^2\text{s}^{-1}$ for 90th centile ADC [30]. Histogram analyses have shown promise for the characterisation of prostate tumours, with the 10th centile demonstrating better correlation to Gleason grade than mean and median ADC values [9] and 90th centile K_{app} values of diffusion kurtosis imaging outperforming other diffusion-weighted imaging metrics for differentiating lower from higher grader tumours [31]. Although no studies have used histogram analysis for the assessment of prostate cancer treatment response, Kyriazi et al demonstrated the 25th ADC centile to be the best predictor of chemotherapy response in patients with metastatic ovarian and primary peritoneal cancer [20]. Despite this later study not quantifying histogram repeatability, they showed that the coefficient of repeatability for mean ADC was 9.5%. MRI is increasingly playing a role in the follow-up of patients on active surveillance (AS) for prostate cancer [32]. The Response Evaluation Criteria in Solid Tumours (RECIST) criteria used in oncology to define progression can only be applied for lesions $\geq 10 \text{ mm}$ [33]. However, many of the low volume low-grade tumours suitable for AS do not meet this threshold, and therefore evaluation of functional measures such as DWI might be considered as surrogate markers of progression [34,35]. In order for DWI to be implemented as a measure of response to supplement the more established size criteria of RECIST, the reproducibility error of the test needs to be quantified in order to determine what constitutes meaningful change outside the range of normal variation. In patients on AS, Morgan et al showed that a whole gland reduction of $> 10\%$ was associated with disease progression, however, non-progressors additionally showed a less marked decrease in ADC [35]. These differences are within the repeatability error demonstrated here and within other studies. Although it should be noted that the error we report is for small tumours rather than whole gland where the difference might be much smaller, this is the more typical clinical situation in active surveillance where small lesions are compared over time. This highlights the limitations of measurement reliability, where a significant difference can often be established for large patient populations, but the difference is insufficient to allow for prediction on an individual basis.

In contrast to the results for the centile values, an analysis of the repeatability of ADC heterogeneity and simplified texture analysis was more equivocal. First, we found that histogram skew, which evaluates its asymmetry, had poor repeatability. More complex forms of heterogeneity and texture analysis are available, but two basic underlying characteristics are dispersion of the voxel values present, as measured by the interquartile range, and the spatial variation of those values, measured through the mean local range. We found that the repeatability for mean local range was better than IQR, having a coefficient of repeatability of $0.112 \times 10^3 \text{ mm}^2\text{s}^{-1}$ compared to $0.181 \times 10^3 \text{ mm}^2\text{s}^{-1}$ for IQR. Indeed, the coefficient of repeatability for mean local range of tumour ADC was 37% expressed as a percentage of scan 1 mean compared to 373% for IQR. Our results suggest that an evaluation of ADC heterogeneity and basic texture features might provide additional benefit in the analysis of prostate tumours, however careful consideration of the repeatability and reliability of the measurements will be essential. This is important given the increasing interest in developing more advanced means of texture analysis to characterise disease

at baseline and assess response in the follow-up setting [36]. Finally, we found that inclusion of the tumour border had only a minimal impact on the ADC heterogeneity values we assessed, suggesting that partial volume effect and minor errors in ROI transposition did not substantially affect the results.

Our study has several limitations. First, although prostatectomy provides more definitive histology, there is a potential selection bias in only including patients fit to undergo surgery. Second, this study had a small sample size and the evaluation in a larger study is warranted, in particular for differences relating to tumour size, origin or grade. We did not assess additional sources of error such as inter-observer differences or repeatability by testing intra-observer variation. Intra-observer repeatability has been shown to be around 10% for DWI regardless of whether 2D or 3D ROIs are used [37], and inter-observer variation brings an additional subjective element of assessing lesion presence and conspicuity, which would be better assessed from blinded clinical reads, rather than direct pathological correlation.

In summary, we found that 10th and 90th centile ADC had sufficient repeatability to be considered for clinical use. While more advanced heterogeneity assessment might be clinically beneficial, the use of values such as histogram IQR or mean local range could be limited by their repeatability and further evaluation is warranted.

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Appendix A. Supplementary data

Supplementary material related to this article can be found, in the online version, at doi:<https://doi.org/10.1016/j.ejrad.2018.11.014>.

References

- [1] R.L. Siegel, et al., Cancer Statistics, 2017, *CA Cancer J. Clin.* 67 (1) (2017) 7–30.
- [2] J. Maddams, et al., Projections of cancer prevalence in the United Kingdom, 2010–2040, *Br. J. Cancer* (107) (2012) 1195–1202.
- [3] R. Kvåle, et al., Concordance between Gleason scores of needle biopsies and radical prostatectomy specimens: a population-based study, *BJU Int.* 103 (12) (2009) 1647–1654.
- [4] Y. Hu, et al., A biopsy simulation study to assess the accuracy of several transrectal ultrasonography (TRUS)-biopsy strategies compared with template prostate mapping biopsies in patients who have undergone radical prostatectomy, *BJU Int.* 110 (2012) 812–820.
- [5] H.U. Ahmed, et al., Diagnostic accuracy of multi-parametric MRI and TRUS biopsy in prostate cancer (PROMIS): a paired validating confirmatory study, *Lancet* 389 (10071) (2017) 815–822.
- [6] V. Kasivisvanathan, A.S. Rannikko, M. Borghi, et al., MRI-targeted or standard biopsy for prostate-cancer diagnosis, *N. Engl. J. Med.* 378 (19) (2018) 1767–1777.
- [7] R. Nagarajan, D. Margolis, S. Raman, T. McClure, S. Raman, M.A. Thomas, Correlation of Gleason scores with diffusion-weighted imaging findings of prostate cancer, *Adv. Urol.* 2012 (2012) 374805.
- [8] T. Barrett, A.N. Priest, E.M. Lawrence, D.A. Goldman, A.Y. Warren, V.J. Gnanapragasam, E. Sala, F.A. Gallagher, Ratio of tumor to normal prostate tissue apparent diffusion coefficient as a method for quantifying DWI of the prostate, *AJR Am. J. Roentgenol.* 205 (6) (2015) W585–93.
- [9] O.F. Donati, Y. Mazaheri, A. Afaq, et al., Prostate cancer aggressiveness: assessment with whole-lesion histogram analysis of the apparent diffusion coefficient, *Radiology* 271 (2014) 143–152.
- [10] J.C. Weinreb, et al., PI-RADS prostate imaging - reporting and data system: 2015, version 2, *Eur. Urol.* 69 (1) (2016) 16–40.
- [11] T. Barrett, B. Turkbey, P.L. Choyke, PI-RADS version 2: what you need to know, *Clin. Radiol.* 70 (11) (2015) 1165–1176.
- [12] M. Brizmohun Appayya, J. Adshad, H.U. Ahmed, et al., National implementation of multi-parametric magnetic resonance imaging for prostate cancer detection - recommendations from a UK consensus meeting, *BJU Int.* 122 (1) (2018) 13–25.
- [13] A. Braithwaite, B. Dale, D. Boll, E. Merkle, Short- and midterm reproducibility of apparent diffusion coefficient measurements at 3.0-T diffusion-weighted imaging of the abdomen, *Radiology* 250 (2) (2009) 459–465.
- [14] A.B. Rosenkrantz, M. Oei, J.S. Babb, B.E. Niver, B. Taouli, Diffusion-weighted

- imaging of the abdomen at 3.0 Tesla: image quality and apparent diffusion coefficient reproducibility compared with 1.5 Tesla, *J. Magn. Reson. Imaging* 33 (1) (2011) 128–135.
- [15] M.E. Miquel, A.D. Scott, N.D. Macdougall, R. Boubertakh, N. Bharwani, A.G. Rockall, In vitro and in vivo repeatability of abdominal diffusion-weighted MRI, *Br. J. Radiol.* 85 (1019) (2012) 1507–1512.
- [16] M. Sadinski, M. Medved, I. Karademir, S. Wang, Y. Peng, Y. Jiang, S. Sammet, G. Karczmar, A. Oto, Short-term reproducibility of apparent diffusion coefficient estimated from diffusion-weighted MRI of the prostate, *Abdom. Imaging* 40 (7) (2015) 2523–2528.
- [17] A. Fedorov, M.G. Vangel, C.M. Tempany, F.M. Fennessy, Multiparametric magnetic resonance imaging of the prostate: repeatability of volume and apparent diffusion coefficient quantification, *Invest. Radiol.* 52 (9) (2017) 538–546.
- [18] K. Downey, S.F. Riches, V.A. Morgan, S.L. Giles, A.D. Attygalle, T.E. Ind, et al., Relationship between imaging biomarkers of stage I cervical cancer and poor-prognosis histologic features: quantitative histogram analysis of diffusion-weighted MR images, *AJR Am. J. Roentgenol.* 200 (2) (2013) 314–320.
- [19] A.B. Rosenkrantz, M.J. Triolo, J. Melamed, H. Rusinek, S.S. Taneja, F.M. Deng, Whole-lesion apparent diffusion coefficient metrics as a marker of percentage Gleason 4 component within Gleason 7 prostate cancer at radical prostatectomy, *J. Magn. Reson. Imaging* 41 (2015) 708–714.
- [20] S. Kyriazi, D.J. Collins, C. Messiou, K. Pennert, R.L. Davidson, S.L. Giles, et al., Metastatic ovarian and primary peritoneal cancer: assessing chemotherapy response with diffusion-weighted MR imaging—value of histogram analysis of apparent diffusion coefficients, *Radiology* 261 (1) (2011) 182–192.
- [21] Y. Peng, Y. Jiang, T. Antic, M.L. Giger, S.E. Eggener, A. Oto, Validation of quantitative analysis of multiparametric prostate MR images for prostate cancer detection and aggressiveness assessment: a cross-imager study, *Radiology* 271 (2014) 461–471.
- [22] T. Barrett, M. McLean, A.N. Priest, et al., Diagnostic evaluation of magnetization transfer and diffusion kurtosis imaging for prostate cancer detection in a re-biopsy population, *Eur. Radiol.* 28 (8) (2018) 3141–3150.
- [23] J.M. Bland, D.G. Altman, Statistical methods for assessing agreement between two methods of clinical measurement, *Lancet* 1 (8476) (1986) 307–310.
- [24] A. Padhani, G. Liu, D. Koh, T. Chenevert, H. Thoeny, T. Takahara, et al., Diffusion-weighted magnetic resonance imaging as a cancer biomarker: consensus and recommendations, *Neoplasia* 11 (2) (2009) 102–125.
- [25] A.B. Rosenkrantz, J.S. Babb, S.S. Taneja, J.M. Ream, Proposed adjustments to PI-RADS version 2 decision rules: impact on prostate cancer detection, *Radiology* 283 (1) (2017) 119–129.
- [26] N.L. Hansen, B.C. Koo, A.Y. Warren, C. Kastner, T. Barrett, Sub-differentiating equivocal PI-RADS-3 lesions in multiparametric magnetic resonance imaging of the prostate to improve cancer detection, *Eur. J. Radiol.* 95 (2017) 307–313.
- [27] A. Braithwaite, B. Dale, D. Boll, E. Merkle, Short- and midterm reproducibility of apparent diffusion coefficient measurements at 3.0-T diffusion-weighted imaging of the abdomen, *Radiology* 250 (2) (2009) 0459–0465.
- [28] P. Gibbs, M. Pickles, L. Turnbull, Repeatability of echo-planar-based diffusion measurements of the human prostate at 3 T, *Magn. Reson. Imaging* 25 (10) (2007) 1423–1429.
- [29] D.M. Koh, M. Blackledge, D.J. Collins, A.R. Padhani, T. Wallace, B. Wilton, et al., Reproducibility and changes in the apparent diffusion coefficients of solid tumours treated with combretastatin A4 phosphate and bevacizumab in a two-centre phase I clinical trial, *Eur. Radiol.* 19 (11) (2009) 2728–2738.
- [30] L. Heijmen, E.E. Ter Voert, I.D. Nagtegaal, P. Span, J. Bussink, C.J. Punt, et al., Diffusion-weighted MR imaging in liver metastases of colorectal cancer: reproducibility and biological validation, *Eur. Radiol.* 23 (3) (2013) 748–756.
- [31] Q. Wang, H. Li, X. Yan, et al., Histogram analysis of diffusion kurtosis magnetic resonance imaging in differentiation of pathologic Gleason grade of prostate cancer, *Urol. Oncol.* 33 (8) (2015) 337.e15–337.e24.
- [32] National Institute for Health and Care Excellence, Prostate Cancer: Diagnosis and Treatment. Clinical Guideline 175 (Update of Clinical Guideline 58), Available at: (2014) [Accessed 13.11.18] <https://www.nice.org.uk/guidance/cg175>.
- [33] E.A. Eisenhauer, P. Therasse, J. Bogaerts, et al., New response evaluation criteria in solid tumors: revised RECIST guideline (version 1.1), *Eur. J. Cancer* 45 (2009) 228–247.
- [34] T. Barrett, M.A. Haider, The emerging role of MRI in prostate Cancer Active surveillance and ongoing challenges, *AJR Am. J. Roentgenol.* 208 (1) (2017) 131–139.
- [35] V.A. Morgan, S.F. Riches, K. Thomas, N. Vanas, C. Parker, S. Giles, et al., Diffusion-weighted magnetic resonance imaging for monitoring prostate cancer progression in patients managed by active surveillance, *Br. J. Radiol.* 84 (997) (2011) 31–37.
- [36] E. Sala, E. Mema, Y. Himoto, et al., Unravelling tumour heterogeneity using next-generation imaging: radiomics, radiogenomics, and habitat imaging, *Clin. Radiol.* 72 (1) (2017) 3–10.
- [37] T. Tamada, C. Huang, J.M. Ream, M. Taffel, S.S. Taneja, A.B. Rosenkrantz, Apparent diffusion coefficient values of prostate Cancer: comparison of 2D and 3D ROIs, *AJR Am. J. Roentgenol.* 210 (1) (2018) 113–117.