



# Psychiatric comorbidity as a risk factor for mortality in people with anorexia nervosa

Hubertus Himmerich<sup>1,2</sup> · Matthew Hotopf<sup>1,2</sup> · Hitesh Shetty<sup>2</sup> · Ulrike Schmidt<sup>1,2</sup> · Janet Treasure<sup>1,2</sup> · Richard D. Hayes<sup>1</sup> · Robert Stewart<sup>1,2</sup> · Chin-Kuo Chang<sup>1,2,3</sup> 

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## Abstract

Anorexia nervosa (AN) is found associated with increased mortality. Frequent comorbidities of AN include substance use disorders (SUD), affective disorders (AD) and personality disorders (PD). We investigated the influence of these psychiatric comorbidities on all-cause mortality with demographic and socioeconomic factors considered as confounders in the observation window between January 2007 and March 2016 for 1970 people with AN, using data from the case register of the South London and Maudsley (SLaM) NHS Foundation Trust, an almost monopoly-secondary mental healthcare service provider in southeast London. We retrieved data from its Clinical Records Interactive Search (CRIS) system as data source. Mortality was ascertained through nationwide tracing by the UK Office for National Statistics (ONS) linked to CRIS database on a monthly basis. A total of 43 people with AN died during the observation period. Standardized Mortality Ratio (SMR) with England and Wales population in 2012 as standard population for our study cohort was 5.21 (95% CI 3.77, 7.02). In univariate analyses, the comorbidity of SUD or PD was found to significantly increase the relative risks of mortality (HRs = 3.10, 95% CI 1.21, 7.92; and 2.58, 95% CI 1.23, 5.40, respectively). After adjustment for demographic and socioeconomic covariates as confounders, moderately but not significantly elevated risks were identified for SUD (adjusted HR = 1.39, 95% CI 0.53, 3.65) and PD (adjusted HR = 1.58, 95% CI 0.70, 3.56). These results suggest an elevated mortality in people with AN, which might be, at least partially, explained by the existence of the comorbidities SUD or PD.

**Keywords** Anorexia nervosa · Comorbidities · Affective disorder · Substance use disorder · Personality disorder

## Introduction

### Anorexia nervosa

Anorexia nervosa (AN) is defined by a significantly low body weight in the context of age, sex, and physical health,

intense fear of weight gain and disturbed body perception [1]. The number of people with AN is rising, and the illness is affecting people at an increasingly younger age [2–4]. The prevalence of AN is between less than 1% and 4% among women, with a male:female ratio of approximately 1:10 [5]. The peak age of incidence is between 14 and 17 years old [6].

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✉ Chin-Kuo Chang  
ckchang@utapei.edu.tw

<sup>1</sup> Department of Psychological Medicine, King's College London, London, UK

<sup>2</sup> South London and Maudsley NHS Foundation Trust, London, UK

<sup>3</sup> Department of Health and Welfare, University of Taipei, No. 101, Sec. 2, Zhongcheng Rd, Shilin District, Taipei 111, Taiwan

### Mortality for people with AN

Dietary deficit is often accompanied by significant physical health issues, including growth retardation, osteopenia, osteoporosis, amenorrhoea, renal insufficiency, and changes in laboratory parameters, cardiac arrhythmia and disturbances of the thyroid function [7]. The most common causes of death in people with AN are sudden cardiac death associated with ventricular arrhythmias, and suicide [8–10]. Mortality rates are five to six times higher for people with AN than in the general population, and in people aged 15–24 years old,

the mortality risk from AN is higher than for other serious diseases in adolescence, such as asthma or type 1 diabetes [3, 4].

### **Mental health comorbidities in AN as risk factors for mortality**

AN is often accompanied by other mental disorders, including substance use disorders (SUD) [11], affective disorders (AD) [12, 13] and personality disorders (PD) [14]. In the Swedish Twin Registry, people with AN exhibited a significantly increased risk for the use of alcohol and illicit drugs, specifically diet pills and stimulants [11]. Another Swedish study which evaluated data from patients presenting to specialist eating disorder clinics between 2008 and 2012 found that 521 (38.8%) of 1343 patients with AN suffered from AD, principally unipolar depression [12]. A German investigation in female inpatients with AN reported a prevalence of around 30% for one or more comorbid PDs [14]. In this study, comorbid PD was found related to a worse treatment outcome. Potential comorbidities of AN including addiction to alcohol [15, 16] or illicit drugs [17], unipolar depression [18] and PD [19] have been found to be associated with increased mortality. In a recent retrospective Swedish register study investigating mortality in women with AN, SUDs and depression significantly contributed to mortality, but not PDs [20].

Other important risk factors contributing to mortality in people with mental disorders include age, gender, marital status, socioeconomic deprivation and ethnicity [21–27]. However, these risk factor influence mortality in people with physical diseases, too [28–30]. Therefore, they seem to be general factors which modify mortality.

Comparable studies, e.g., Kask et al. [20, 31] only covered patients aged 10–40 years. Therefore, results of these studies may not be generalizable to all age groups. We decided not to set an age limit to achieve a comprehensive epidemiologic view on mortality of patients with AN. Moreover, there are case reports of patients newly diagnosed with AN at the age of 65 or older [32], and a recent study demonstrated that active eating disorders are common amongst women in mid-life, both due to new onset and chronic disorders [33].

### **General aspects of mortality in patients with mental disorders**

When thinking about psychiatric comorbidity as a risk factor for mortality in people with AN, it seems sensible to take some important general aspects of mortality in patients with mental disorders into account. According to a recent systematic review and meta-analysis including 203 eligible articles and representing 29 countries in six continents, mortality is significantly higher among people with

mental disorders than among the comparison population. About two-thirds of deaths among people with mental disorders are due to natural, about 18% to unnatural, and the remainder to other or unknown causes. The median years of potential life lost was found to be 10 years, and it was estimated that around 14% of deaths worldwide (8 million deaths each year) are attributable to mental disorders [34]. Important overall risk factors contributing to mortality in patients with mental disorders include sociodemographic variables such as age, gender, ethnicity, unemployment and low educational attainment, physical diseases such as cancer and coronary heart disease, the specific psychiatric diagnosis and additional other psychological and behavioural problems like substance abuse and dependence; unnatural and alcohol-related causes of death seem to play a specific role for increased mortality in younger age groups of people with mental disorders [35–38].

### **Aim of the study**

In this study, we sought to investigate the influence of major psychiatric comorbidities on all-cause mortality in people with AN in a case register for a large secondary mental healthcare service provider in southeast London, UK, with confounding from demographic and socioeconomic factors considered. Our hypothesis for this investigation was that AD, SUD and PD had an independent effect on increased mortality in people with AN.

## **Methods**

### **Setting and study design**

A retrospective cohort study was performed in people with AN. The AN cohort was obtained from the South London and Maudsley NHS Foundation Trust (SLaM) which provides near-monopoly-secondary mental healthcare services to a geographic catchment of about 1.36 million residents in southeast London. SLaM's Eating Disorders Service additionally offers a specialist tertiary service for service users across the UK. Patients included in this study mainly resided in the London boroughs of Lambeth, Croydon, Lewisham, Southwark, Bromley, Bexley and Greenwich. The SLaM National Institute for Health Research Biomedical Research Centre (BRC) supports the infrastructure for research of anonymised electronic clinical records from SLaM's case register, using the BRC's Clinical Record Interactive Search (CRIS) system as data source, established in 2008 [39, 40].

## Inclusion criteria and primary exposure

The analyzed cohort was extracted via CRIS, consisting of individuals who had received an International Classification of Diseases (ICD-10) diagnosis [41] of AN (F50.0/F50.1) within or before the observation window from 1 January 2007 to 31 March 2016. Structured information on diagnosis from drop-down fields in the source record was supplemented by information extracted from open-text fields using a bespoke algorithm generated using the Generalised Architecture for Text Engineering (GATE) software [39]. The exposure of interest was a further diagnosis of SUD (F10–F19), AD (F31–33) or PD (F60) at any stage during or before the observation window extracted from CRIS using identical procedures.

## Primary outcome

The outcome of this analysis was all-cause mortality occurring from January 2007 to March 2016. Information about each death was collected through nationwide tracing using data from the UK Office for National Statistics (ONS) linked to the SLaM database on a monthly basis. In the UK, all death certifications are linked by NHS number (a unique identifier for each UK NHS service user) to all healthcare providers.

## Covariates

Date of birth (for year and month), gender, ethnicity, marital status and residing borough were retrieved from the CRIS database. Age in years was calculated from the individual's first AN diagnosis date in the observation period or on 1 January 2007 if the AN diagnosis preceded the observation period.

Recorded ethnicity was classified into three categories: 'white' (including 'British', 'any other white background' and 'Irish'), 'black' (including 'African', 'any other black background' and 'Caribbean'), and 'others/unknown/mixed'. Recorded marital status was classified into four groups: 'married' (including 'cohabitation', 'married' and 'married/civil partnership'), 'single', 'divorced' (including 'divorced/civil partnership dissolved', 'separated'), and 'others'. Borough was classified by the subject's residence: Lambeth, Croydon, Lewisham, Southwark, Bromley, Bexley and Greenwich, defined by their primary care trust. Multiple deprivation score is a small-area-level measure of socioeconomic status derived from UK Census data, based on the individual's address closest to the diagnosis of AN in the observation window (or 1 January 2007 for those with an earlier diagnosis of AN), covering the seven components, 'employment', 'income', 'education', 'health', 'barriers to housing and services', 'crime' and 'living environment',

with specific weightings. The Index of Multiple Deprivation score is a well-established measure which has been widely used as a regional indicator for socioeconomic status in previous studies [42–44]. The score is then transformed into percentiles (from 1 to 100) with a higher score indicating greater deprivation (and thus lower socioeconomic status in the neighbourhood). In the analysis, deprivation scores were regrouped by tertiles (33.3 percentiles).

## Ethical approval

Ethical approval was obtained from the Oxfordshire Research Ethics Committee C for the use of CRIS as a source of data for secondary analysis (reference 08/H0606/71 + 5) including the mortality linkage. The study has been performed in accordance with the ethical standards laid down in the 1964 Declaration of Helsinki and its later amendments.

## Statistical analysis

Standardized Mortality Ratio (SMR) was used for the adjustment of age and gender structure, comparing our study cohort to a standard population for mortality. Observed number of death in the window was treated as the numerator of SMR in the calculation. Age spanning in 5-year age and sex bands were then utilized to define age- and gender-specific groups in our study cohort. Denominator of SMR was the sum of expected numbers for age-and-sex bands derived by the number of people in individual band in our study cohort multiplying its corresponding age- and gender-specific mortality rate in the general population of England and Wales in 2012 according to ONS data. Since the study cohort was followed up for up to 9.25 years but mortality rates for the general population were given for a year only, specific weightings based on the average follow-up period for corresponding age-and-sex bands for the calculation of SMR's denominator were applied.

Descriptive statistics on demographic characteristics and psychiatric comorbidities were first carried out for the study cohort. We then used Cox proportional hazards regression to estimate the effect for each of the variable by hazard ratio (HR) with a 95% confidence interval (CI) on death as the 'event' of interest within the observation window. Beginning of the follow-up period was defined as the start of the observation window (1 January 2007) if the first diagnosis was given before the window, otherwise the beginning was taken as the date of first AN diagnosis given in the window. Multivariate analyses were lastly performed in parallel to estimate the effect of the considered psychiatric comorbidities as major exposure of interest, controlling for confounding of age, sex, ethnicity, borough, marital status and deprivation score. All the statistical analyses were performed

using Stata 12.1 (StataCorp, College Station, Texas, USA) and the statistical significance (alpha level) was set at 0.05.

## Results

### Descriptive statistics

A total of 1970 individuals diagnosed with AN were identified from the CRIS system. Among them, 1834 were female and 136 were male. The mean age was 24.5 years (standard deviation: 11.3 years; range 8–88 years). Twelve (0.6%) patients with AN were 65 years old or older. More detailed information about age distribution was put in Table 1. During the observation window, 43 of the 1,970 study subjects died. Descriptive statistics are shown in detail in Table 2.

### Standardized mortality ratio (SMR)

Age- and gender-SMR for our study cohort was 5.21 (95% CI 3.77, 7.02), with the population of England and Wales in 2012 as standard population according to ONS data on its website.

### Univariate analysis by Cox regressions

Cox regressions revealed a significant effect of age, marital status, deprivation score, SUD and PD. Specifically, older age, being separated or divorced, being deprived and suffering from comorbidity of SUD and PD were associated with a higher mortality risk (details shown in Table 2). The

mean age of the patients who died in the observation window was 41.03 years old (standard deviation = 18.36 years; ranging from 15 to 86 years old), whereas the mean age of patients who survived was 24.11 years old (standard deviation = 10.79 years; ranging from 8 to 88 years old).

### Multivariate analyses by Cox regressions

Controlling for age, sex, ethnicity, the residing borough, marital status and deprivation score as confounders, multivariate analysis outcomes by Cox regression is shown in Table 3. Significance of the comorbidities to mortality in people with AN disappeared, when controlling for potential confounders. However, moderately elevated risks were identified for SUD (adjusted HR = 1.39, 95% CI 0.53, 3.65) and PD (adjusted HR = 1.58, 95% CI 0.70, 3.56). All detailed information regarding multivariate analyses outcomes for AD, SUD and PD as the major exposure of interest was separately presented in Supplementary Tables 1–3.

## Discussion

### Main findings

In summary, 43 of the 1970 people with AN died during the observation period of more than 9 years. A more than five folds of general mortality was found by age and gender standardization, compared to the general population in England and Wales. Moderately elevated mortality risks were estimated for comorbid SUD and PD with confounders controlled, but not significant because of power issues.

### Comparison to previous research

The SMR in our study cohort is comparable to previous studies, reporting that patients with AN have five to six times higher mortality than the general population [3, 4]. However, there are also studies reporting even higher all-cause mortality rates [20, 31], but these studies included inpatients only. One can assume that inpatient status reflects a higher degree of severity.

In a recent meta-analysis of mortality studies in AN [45], age was—apart from the applied diagnostic criteria—the most significant factor influencing mortality. The finding that age and social deprivation significantly influence mortality is also in accordance with Franko et al. [46] who investigated 246 treatment-seeking women with AN or bulimia nervosa longitudinally, among them 136 patients with AN. However, in contrast to Franko et al., we could not replicate their finding of a significant contribution of alcohol abuse to all-cause mortality in AN.

**Table 1** Age distribution of study participants ( $N=1970$ )

Age group (years)	Frequency	Percent
5–10	4	0.20
10–15	284	14.42
15–20	602	30.56
20–25	375	19.04
25–30	255	12.94
30–35	152	7.72
35–40	100	5.08
40–45	64	3.25
45–50	47	2.39
50–55	33	1.68
55–60	29	1.47
60–65	13	0.66
65–70	4	0.25
70–75	2	0.10
75–80	3	0.15
80–85	0	0.00
85–90	2	0.10

**Table 2** Univariate analysis on demographics and clinical characteristics for mortality among people with anorexia nervosa by Cox regressions ( $N=1970$ )

Variables	Mean $\pm$ SD/numbers (% in column)	Number of deaths (% in row)	Hazard ratio (95% CI)	<i>p</i> value
Age (years old)	24.48 $\pm$ 11.27	–	1.08 (1.06, 1.10)*	<0.01
Gender				
Female	1,834 (93.10%)	39 (2.13%)	Ref	
Male	136 (6.90%)	4 (2.94%)	1.33 (0.47, 3.71)	0.59
Marital status				
Single	1,532 (77.77%)	24 (1.57%)	Ref	
Married/partner	175 (8.88%)	6 (3.43%)	2.12 (0.87, 5.19)	0.10
Divorced/separated	43 (2.18%)	3 (6.98%)	3.75 (1.13, 12.48)*	<0.05
Others	220 (11.17%)	10 (4.55%)	2.35 (1.12, 4.93)*	<0.05
Borough				
Lambeth	298 (15.13%)	3 (1.01%)	Ref	
Croydon	237 (12.03%)	8 (3.38%)	2.61 (0.69, 9.85)	0.16
Lewisham	217 (11.02%)	4 (1.84%)	1.57 (0.35, 7.03)	0.55
Southwark	219 (11.12%)	7 (3.20%)	2.89 (0.75, 11.17)	0.12
Bromley	253 (12.84)	5 (1.98%)	1.87 (0.45, 7.81)	0.39
Bexley	114 (5.79%)	4 (3.51%)	3.15 (0.71, 14.09)	0.13
Greenwich	103 (5.23%)	2 (1.94%)	1.83 (0.30, 10.93)	0.51
Others	529 (26.85%)	10 (1.89%)	1.72 (0.47, 6.24)	0.41
Ethnicity				
White	1,639 (83.20%)	34 (2.07%)	Ref	
Black	63 (3.20%)	2 (3.17%)	1.27 (0.30, 5.28)	0.74
Others/mixed	268 (13.60%)	7 (2.61%)	1.27 (0.56, 2.87)	0.56
Deprivation Score <sup>a</sup>				
1.14–17.60	644 (32.69%)	10 (1.55%)	Ref	
17.61–29.17	643 (32.64%)	10 (1.56%)	0.92 (0.38, 2.22)	0.86
29.18–100	645 (32.74%)	22 (3.41%)	2.12 (1.01, 4.50)*	<0.05
Missing/homeless	38 (1.93%)	1 (2.63%)	1.37 (0.17, 10.69)	0.77
SUD				
No	1909 (96.90%)	38 (1.99%)	Ref	<0.05
Yes	61 (3.10%)	5 (8.20%)	3.10 (1.21, 7.92)*	
AD				
No	1,680 (85.28%)	34 (2.02%)	Ref	0.37
Yes	290 (14.72%)	9 (3.10%)	1.40 (0.67, 2.93)	
PD				
No	1,798 (91.27%)	34 (1.89%)	Ref	<0.05
Yes	172 (8.73)	9 (5.23%)	2.58 (1.23, 5.40)*	

Higher deprivation scores indicate greater deprivation and thus lower socioeconomic status

SUD substance use disorder, AD affective disorder, PD personality disorder

\*Statistically significant according to a level of significance of 0.05

<sup>a</sup>Further information

Considering sample size, data characteristics and the link between a clinical register and a death register, our study seems to be most comparable to the Swedish register studies where the Swedish Hospital Discharge Register is linked to the Swedish Cause-of-Death Register. The latest comparable Swedish studies investigated mortality in 8,069 female inpatients [20] and 609 males who received hospitalized care for AN [31]. Both studies found that SUD significantly

contributed to mortality. These two studies included only patients aged 10–40 years to minimize the risk of incorrect diagnoses [20, 31]. In contrast, our study included patients with AN at all ages. If we look at SUD across the age range, SUD are most prevalent and most important between 10 and 40 years of age [47, 48]. Therefore, results of these two studies by Kask et al. may not be generalizable to all age groups of patients with AN [20, 31]. We decided not to set an age

**Table 3** Multivariate analysis on demographics and clinical characteristics for mortality among people with anorexia nervosa by Cox regressions, controlling age, sex, ethnicity, borough, marital status and deprivation score ( $N=1970$ )

Comorbidity	Adjusted hazard ratio (95% CI)	<i>p</i> value
AD	0.80 (0.36, 1.81)	0.60
SUD	1.39 (0.53, 3.65)	0.50
PD	1.58 (0.70, 3.56)	0.27

SUD substance use disorder, AD affective disorder, PD personality disorder

limit for inclusion of people in our study for three reasons. First, we sought to achieve a comprehensive epidemiologic view on mortality of patients with AN. One should keep in mind though, while interpreting the results of our study, that a diagnosis of AN did neither require that AN was the leading current diagnosis nor that it was the first episode of AN. Second, current research has highlighted the occurrence of AN in mid-life and older adults [32, 33, 49]. Third, studies in AN have used different upper age limits, for example, 40 years of age [20, 31], even though other studies have shown that in a relevant proportion of people with AN is first diagnosed or first treated after the age of 40 [32, 33]. For that reason, imposing a specific age limit seemed arbitrary.

Moderately elevated mortality risks were identified for SUD and PD in people with AN in our analysis, although not significant. Besides the issue of potentially insufficient statistical power, marital status might work as a critical mediator within the causal pathway between AN and mortality, taking away the effect from AN to mortality in the effect estimation by multivariate analysis. However, when we put marital status away from the final model, the outcomes for PD and SUD comorbidity did not change dramatically. Thus, the possibility of a mediating role for marital status seems limited.

### Strengths and limitations

The main strengths of this study were the study design and size of cohort using CRIS as our data source. CRIS is one of the largest collections of secondary and tertiary mental healthcare services. Using this data in a retrospective cohort study, we identified a large number of 1970 patients with AN. However, in the observation period, only 43 patients with AN died, resulting in limited statistical power. The research outcomes of this study might be only generalizable to secondary and tertiary health care settings, but not to people with AN in community. In terms of generalizability, we also should be cautious that diagnoses were coded according to ICD-10 [41] which is still the currently used classification in the UK. Nevertheless, it needs to be mentioned that criteria have slightly changed in DSM-5 [1].

In principle, using data derived from secondary and tertiary mental healthcare services may potentially lead to an overestimation of the effect of a certain disorder on mortality and might compromise generalizability. The potential reasons for this assumption are that some patients with the disorder in question may not be under the care of any health service, and some of those cared for in primary care might not be referred to a specialist service. Thus, those who are referred to a specialist service might be more severely affected than the whole group of patients with this specific disorder.

However, studies which have successfully ascertained prevalence and service use rates in AN are scarce due to the relatively low prevalence of AN compared to other eating disorders. For example, a study by Solmi et al. [50] which examined 1698 individuals between 16 and 90 years of age in southeast London found a point prevalence of eating disorders of 4.4%, but no cases of AN at all. Thus, this study was unable to gauge service use among patient with AN. A US-American study found prevalence rates of AN between 0.03 and 0.39% and service utilization rates of patients with AN between 30 and 100% depending on the ethnicity of patients [51]. For non-Latino white patients with AN service use rates were around 76% [51]. White people are currently the biggest ethnic group in London [52] and in our study sample (see Table 2). Therefore, we can assume a similarly high service use rate for our patients. Nevertheless, we are aware of the ongoing debate whether general practitioners are good or poor at identifying the eating disorders [53].

It is likely that most patients presenting with AN to primary care in the UK are referred on to secondary care, including both general mental health services and specialist eating disorder services [54], because previous [54] and current [55] guidelines of The National Institute for Health and Care Excellence (NICE) in the UK advise general practitioners to refer anyone with a suspected eating disorder to specialist services without delay.

As information about each death was obtained through nationwide tracing using ONS, data on mortality are very reliable. A failure to obtain death data would appear, if people left the country during the observation period and died abroad. Such an event could have led to a failure to register the death of a patient with AN treated in SLAM.

Overall, we think that CRIS data on the diagnosis of AN and on mortality are generalizable because we assume—based on the limited evidence available—that most patients with AN seek medical help and are referred to secondary mental healthcare services [51, 54], and SLAM is an almost monopoly-secondary mental healthcare service provider in southeast London. As we can assume that our data on the diagnosis of AN and on the cases of death are reliable, we think that the obtained SMR for our study cohort of 5.21 is accurate.

Compared to other studies [11–14, 56, 57] we obtained relatively low rates of comorbid SUD, AD and PD in patients with AN. Thus, these disorders may have been underdiagnosed in our study sample, and underdiagnosis could be a source of error. When drawing further conclusions from our data, one must keep in mind that our data on SUD, AD and PD were not collected whilst specifically searching for these diagnoses. Instead, these diagnoses were documented by clinicians when clinically relevant. Therefore, our results cannot be easily compared with studies in which researchers used questionnaires and diagnostic tools to specifically detect SUDs, ADs or PDs.

An important limitation of our investigation is that we were not able to consider the subtypes of AN (binge/purge vs. restrictive), illness severity and duration, the number of hospital admissions during the observation period and the body mass index (BMI) which have all been found to contribute to mortality of people with AN in previous studies [46, 58–60]. As what already mentioned before, we do not know, when the diagnosis was first established in participants of this study.

### Future research directions

For this particular project, the date of death, but not the cause of death was available from the ONS. The causes of death could be physical consequences of AN such as ventricular arrhythmias or suicide [8–10]. Future research could shed more light on the specific causes of death as the consequence of AN.

Further studies should also focus on modifiable risk factors for premature death prevention. Given a significant contribution of social deprivation to all-cause mortality, the question arises whether a reduction of mortality can be achieved by interventions intending to tackle social deprivation, such as family therapy, social work, social inclusion therapy and debt advice.

We have found an alarmingly increased mortality in patients with AN compared to the general population in accordance with previous studies [3, 4, 20, 31]. Thus, the question arises, whether the increased mortality can, in principle, be influenced by the available therapies. However, scientific knowledge on long-term treatment outcomes of most of the various psychotherapeutic, refeeding and psychopharmacological approaches in AN is scarce [61, 62], and we know even less about whether these therapies can reduce mortality.

Psychotherapy can have side effects such as the emergence of new symptoms, the deterioration of existing symptoms, dependence and regression, traumatic stress, strains or changes in family or work relations and missing the chance to seek suitable medical treatment [63, 64]. This is interesting considering our findings that being separated or divorced

and being deprived in the sense of a lower socioeconomic status contribute to mortality in AN. Thus, future research could investigate whether specific therapies influence the occurrence of additional psychiatric symptoms and risk factors in a favourable or in an unfavourable way.

We can assume a widespread use of medications such as olanzapine during inpatient treatment for AN (for example [65]). Moreover, most patient with AN want and request medication to help with anxiety, sleep problems and pain [66, 67]. However, there is almost no evidence for the effects of pharmacological treatment in AN due to a lack of randomized controlled trials [68] apart from a few studies with promising results such as [69]. The effect of psychopharmacological treatment on mortality is also unknown in AN. Due to potential side effects such as electrolyte imbalances and abnormalities of the heart rhythm [70], mortality might even increase in patients with AN during treatment with these medications.

Taken together, potential future research directions to further explore and complement the findings of the present study include the investigation of the specific causes of death, physical risk factors of mortality in patients with AN and the effects of psychotherapeutic, nutritional and psychopharmacological therapies on mortality.

**Author contributions** HH, MH, UD, JT, RDH, RS and CKC designed the study. HH, MH, HS, RDH, RS and CKC contributed to the statistical analysis; HH and CKC drafted the manuscript; all authors contributed to the interpretation of the results, revised the first draft of the manuscript and approved its final version.

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### Compliance with ethical standards

**Conflict of interest** The authors declare no conflict of interest relevant to this paper.

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