



# The impact of robotic colorectal surgery in obese patients: a systematic review, meta-analysis, and meta-regression

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## Abstract

**Background** Robotic surgery (RS) may overcome the limitations of laparoscopic colorectal surgeries (LS) in obese patients, but remains less well studied. This systematic review and meta-analysis aims to evaluate the outcomes of obese patients who have undergone robotic colorectal surgery.

**Methods** This study was performed according to the PRISMA guidelines. A search was performed on Medline, EMBASE, and the Cochrane Library to identify relevant articles. Dichotomous and continuous outcomes were analyzed as risk ratio (RR) and mean difference (MD), respectively. All post-operative outcomes were within 30 days after surgery. The quality of studies was assessed using the Newcastle–Ottawa Scale. Meta-regression analysis was conducted to identify sources of heterogeneity.

**Results** Three studies totaling 262 subjects compared LS (45.0%) against RS (55.0%) in obese patients. The RS group had a significantly reduced length of hospital stay (LOS) (MD –2.55 days, 95%CI –3.13 to –1.97 days,  $P < 0.00001$ ,  $I^2 = 26%$ ) and lower risk of re-admission (RR 0.42, 95%CI 0.19–0.92,  $P = 0.030$ ,  $I^2 = 0%$ ), however, the length of operative time was longer (MD 40.54 min, 95%CI 32.72–48.36 min,  $P < 0.00001$ ,  $I^2 = 37%$ ). Six studies totaling 761 subjects compared obese (40.5%) against non-obese (59.5%) patients who underwent RS. An increased operative time (MD 20.72 min, 95%CI 7.39–34.04 min,  $P = 0.002$ ,  $I^2 = 0%$ ) and risk of wound infection (RR 2.59, 95%CI 1.12–6.02,  $P = 0.030$ ,  $I^2 = 0%$ ) were noted in the former, with no differences in other intra- and post-operative outcomes. Meta-regression revealed that the pathology (rectal, colon, both) ( $P = 0.255$ ), age ( $P = 0.530$ ), gender ( $P = 0.279$ ), and continent that the study originated from ( $P = 0.583$ ) were not significant sources of heterogeneity for the risk of wound infection.

**Conclusion** Compared to laparoscopy, robotic surgery provides earlier recovery with shorter LOS and reduced re-admission rates for obese patients, without compromising on other operative outcomes. Among patients undergoing robotic colorectal surgery, obesity is associated with a longer operative duration and greater risk of wound infection.

**Keywords** Robotics · Laparoscopy · Obesity · Colorectal surgery

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Minimally invasive colorectal surgery has gained widespread acceptance due to its superior short-term outcomes compared to open surgery [1–4]. However, the technical challenges of conventional laparoscopic surgery (LS) have been acknowledged, with obesity being one of the contributory factors [5–7]. Visceral obesity limits working space during laparoscopy due to increased tissue volume, making it more difficult to achieve adequate exposure. Retraction is also more challenging because fatty mesentery tends to tear and bleed more easily. This is aggravated by the greater amount of force required to support the weight of these bulky tissues. Such problems become compounded when operating in the confined space of a narrow pelvis during rectal surgery.

Robotics was introduced to address the limitations of conventional laparoscopy, and the feasibility of robotic surgery (RS) in obese patients has been inferred from the results of earlier publications [8–10]. However, studies specifically investigating the benefits of robotics in this group of patients remain limited. Our objective was therefore to provide a contemporary review of the literature, and a meta-analysis to evaluate the impact of robotic colorectal surgery in obese patients.

## Methods

This systematic review and meta-analysis was conducted in strict accordance to the preferred reporting items for systematic reviews and meta-analyses (PRISMA) statement guidelines [11]. As this was a meta-analysis of published studies, there were no human subjects involved. Hence, IRB approval and written consent were not required.

## Literature search

An electronic search was performed on MEDLINE, EMBASE, and the Cochrane Library from date of inception to 30th August, 2018 to identify all published studies evaluating the safety and efficacy of robotic colorectal surgery in obese patients. A combination of ‘MeSH’ and non-‘MeSH’ search terms using Boolean operators was used in Medline: (robotic.m.p. OR ROBOTICS/OR robot-assisted.mp. OR robotic-assisted.mp.) AND (obesity.mp. OR OBESITY/OR obese.mp. OR body mass index/OR BMI.mp.) AND (colorectal surgery.mp. OR colorectal surgery/OR rectal cancer.mp. OR rectal neoplasms/OR colorectal cancer.mp. OR colorectal neoplasms/). A manual search of the reference lists of included studies was performed to identify additional relevant studies.

## Study selection and inclusion criteria

Two reviewers (IW, JN) screened the studies in an independent fashion for potential inclusion. In the first stage, the abstracts and titles were screened to include potentially relevant studies. Full-texts of these studies were subsequently retrieved and reviewed in their entirety to confirm inclusion. Conflicts were resolved by consensus.

Randomized (RCT) and non-randomized (non-RCT, cohort study, case–control study) studies that evaluated robotic colorectal surgery outcomes were screened. Only studies that exclusively compared obese and non-obese patients were included. The other group of studies reviewed included those that compared the outcomes between RS to LS or open surgery solely in obese patients. Studies of the

following designs were excluded: those that evaluated the impact of obesity as secondary outcomes or subgroup analyses; single-arm; non-English language; case reports and case series; animal and laboratory studies; literature reviews; conference abstracts with no extractable data.

## Comparisons

Two main comparisons were made (1) Robotic versus laparoscopic colorectal surgery for obese patients; (2) Obese versus non-obese patients undergoing robotic colorectal surgery.

## Obesity definitions and data extraction

For simplification of statistical analysis, the obese arm included both obese and morbidly obese patients, while the non-obese arm included normal weight, and overweight patients. Most studies utilized the National Institutes of Health criteria: underweight (Body Mass Index, BMI  $\leq 18.5$  kg/m<sup>2</sup>); normal weight (BMI of 18.6–24.9 kg/m<sup>2</sup>); overweight (BMI of 25–29.9 kg/m<sup>2</sup>); obese I (BMI of 30–34.9 kg/m<sup>2</sup>); obese II (BMI of 35–39.9 kg/m<sup>2</sup>); obese III (BMI of  $> 40$  kg/m<sup>2</sup>). One study measured visceral obesity (VFA) and defined obesity as having a VFA of equal to or greater than 130 cm<sup>2</sup> [12].

Primary outcomes of interest included intra-operative data: conversion, blood loss, and operative time. Secondary outcomes included 30-day post-operative data: length of hospital stay (LOS), anastomotic leak, re-operation, re-admission, wound infection, ileus, and pulmonary complications. Two authors (IW, JN) extracted the following data from each study independently: first author, year, type of publication, country, mean age, proportion of male gender, BMI, and oncologic parameters (use of neo-adjuvant chemotherapy, positive margins, mean lymph node harvested, and tumour size).

## Quality assessment

For cohort studies, the Newcastle–Ottawa Scale (NOS) was employed to assess the quality of included studies against the maximum score of 9 points. It comprises aspects of patient selection, comparability of study groups, and outcome assessment. Publication bias was not assessed due to the lack of studies included in each outcome (less than 10) [13].

## Statistical analysis

The risk ratio (RR) was calculated from each study and pooled using the Mantel–Haenszel method with 95% confidence intervals using Review Manager software (Revman 5.3, The Cochrane Collaboration). For continuous outcomes,

mean differences (MD) were calculated and pooled using the inverse-variance method. Statistical heterogeneity was assessed using the  $I^2$  statistic, where a value above 50% was deemed to be of substantial heterogeneity [14]. A random-effects meta-analysis was employed in such instances. To ensure that findings prevailed with better quality data, leave-one-out sensitivity analyses of exclusion were performed. Although bias and heterogeneity may arise from including both colon and rectal surgery in the same analysis, separate meta-analysis was not possible since the included studies did not stratify these results accordingly.

As such, mixed-effects meta-regression using the *metaprop* command in STATA was conducted to explore potential sources of heterogeneity, such as the type of pathology (rectal, colon, both), patient age, proportion of male gender, and continent. This was only performed when there was a sufficient number of studies (5 or more) in the specific outcome.

## Results

### Systematic search

The screening process is depicted in the PRISMA flow diagram (Fig. 1). A total of 9 retrospective cohort studies were included, of which one was conducted in Japan [12], one in Turkey [15], one in the United Kingdom [16], and six in the United States of America [9, 10, 17–20]. Three studies totaling 262 patients compared LS (118/262; 45.0%) against RS (144/262; 55.0%) in obese patients with rectal cancer [12, 16, 20]. The remaining 6 studies comprising 761 patients compared obese (308/761; 40.5%) against non-obese (453/761; 59.5%) patients who underwent RS for colon and rectal pathologies. Baseline and oncologic parameters are presented in Table 1. The overall risk of bias is low, since all studies scored 8 and above, out of 9 points on the NOS scale. A summary of baseline and oncologic parameters can be found in Table 1.

### First comparison: RS versus LS in obese patients

Obese patients who underwent RS had a significantly reduced LOS (MD  $-2.55$  days, 95%CI  $-3.13$  to  $-1.97$  days,  $P < 0.00001$ ,  $I^2 = 26\%$ ) (Fig. 2), and lower risk of re-admission (RR 0.42, 95%CI 0.19–0.92,  $P = 0.030$ ,  $I^2 = 0\%$ ) (Fig. 3). Of note, 2 of 3 studies employed an enhanced recovery after surgery (ERAS) protocol [16, 20]. The length of operative time was longer (MD 40.54 min, 95%CI 32.72–48.36 min,  $P < 0.00001$ ,  $I^2 = 37\%$ ) (Fig. 4).

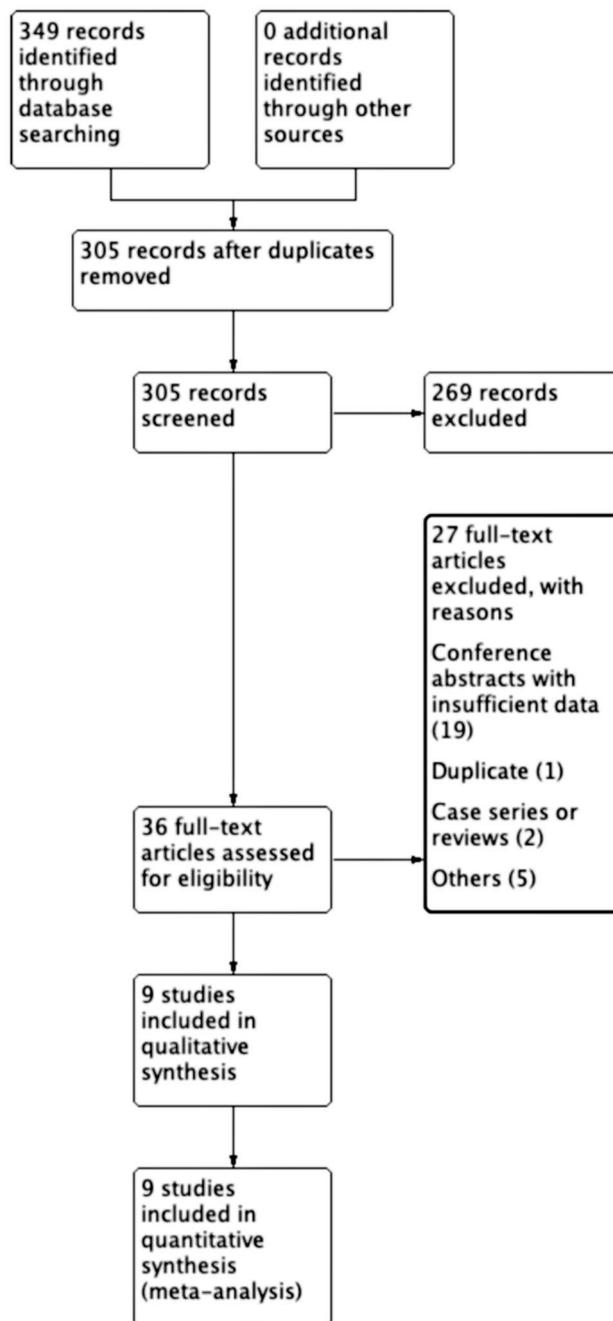


Fig. 1 PRISMA flow diagram

In terms of secondary outcomes, although there was a trend towards a reduced conversion rate in the RS group as compared to the LS group, this was not statistically significant (RR 0.19, 95%CI 0.03–1.05,  $P = 0.060$ ,  $I^2 = 0\%$ ) (Supplementary Fig. 1). Other outcomes including the amount of blood loss (MD  $-30.71$  mL, 95%CI  $-120.66$  to 59.25 mL,  $P = 0.500$ ,  $I^2 = 87\%$ ), risk of anastomotic leak (RR 1.65, 95%CI 0.26–10.55,  $P = 0.590$ ,  $I^2 = 0\%$ ), and re-operation (RR 0.52, 95%CI 0.12–2.39,  $P = 0.380$ ,

**Table 1** Baseline characteristics of included studies

First author, year	Country	da Vinci model	BMI cut-off (kg/m <sup>2</sup> )	Pathological indication	Comparison	No.	Mean age ± S.D (years)	Male (%)	BMI (kg/m <sup>2</sup> )	Chemo (%)	Positive margins (%)	Mean LN harvested ± S.D (n)	Tumour size (cm)
Gogun, 2016	USA	NR	≥30	Rectal Ca	All OB, LAP versus ROB	27/29	60.3 ± 9.8/58.8 ± 10.7	59.3/75.9	35.2 ± 5.0/34.9 ± 7.2	51.9/65.5	NR	21.8 ± 9.6/25.5 ± 14.0	2.8 ± 2.2/2.1 ± 2.0
Panteleimonitis, 2018	UK and Portugal	Si and Xi	≥30	Rectal Ca	All OB, LAP versus ROB	61/63	67.3 ± NR/65.8 ± NR	67.2/63.5	32.0 ± 1.2/32.4 ± 1.7	23.0/38.1	NR	16.0 ± NR/17.0 ± NR	NR
Shiomi, 2016	Japan	NR	VFA ≥ 130 cm <sup>2</sup>	Rectal Ca	All OB, LAP versus ROB	30/52	66.8 ± 9.8/64.0 ± 11.5	80.0/86.5	27.3 ± 4.0/27.6 ± 5.3	0.0/1.9	0.0/0.0	26.0 ± 11.3/32.3 ± 13.0	NR
Keller, 2016	USA	NR	≥30	Benign and malignant colorectal lesions	All OB, versus OB versus NW	45/45	51.7 ± 11.4/51.9 ± 11.6	57.8/57.8	34.5 ± 4.2/25.1 ± 3.3	28.9/24.4	4.4/4.4	15.6 ± 9.5/17.9 ± 9.5	NR
Lagares-Garcia, 2016	USA	Si	≥30	Benign and malignant colorectal lesions	All OB, versus OB versus NW	34/36	56.1 ± 11.6/64.6 ± 14.2	25.0/35.3	NR	NR	NR	NR	NR
Bayraktar, 2017	Turkey	Xi	≥30	Rectal Ca	All OB, versus OB versus NW	30/71	61.0 ± 9.0/60.0 ± 11.0	63.0/65.0	32.0 ± 1.5/25.0 ± 3.0	23.0/39.0	NR	29.0 ± 15/26.0 ± 11.0	4.2 ± 2.5/4.2 ± 1.9
Harr, 2017	USA	NR	≥30	Benign and malignant colorectal lesions	All OB, versus OB versus NW	108/108	57.1 ± 12.4/59.2 ± 11.3	50.0/50.0	36.2 ± 5.7/24.6 ± 3.2	NR	NR	NR	NR
Pai, 2017	USA	NR	≥30	Rectal Ca	All OB, versus OB versus NW	33/68	57.7 ± 11.3/63.4 ± 10.5	54.5/66.2	33.8 ± 3.4/25.2 ± 3.2	87.9/67.6	9.1/2.9	15.9 ± 7.2/14.5 ± 7.5	NR
Duchalais, 2018	USA	Si and Xi	≥30	Benign and malignant colorectal lesions	All OB, versus OB versus NW	58/125	59.0 ± 10.0/58.0 ± 13.0	69.0/69.0	NR	59.0/58.0	2.0/2.0	26.0 ± 11.0/27.0 ± 14.0	NR

PC prospective cohort, LAP laparoscopic, ROB robotic/robotic-assisted, OB obese, NW normal weight, BMI body mass index, LN lymph node, NR not reported. All parameters are reported as LAP/ROB or OB/NW

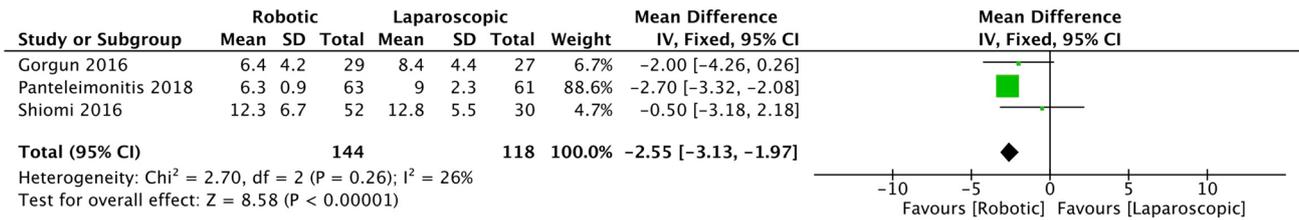


Fig. 2 Forest plot of length of hospital stay in comparison 1



Fig. 3 Forest plot of re-admission rates in comparison 1

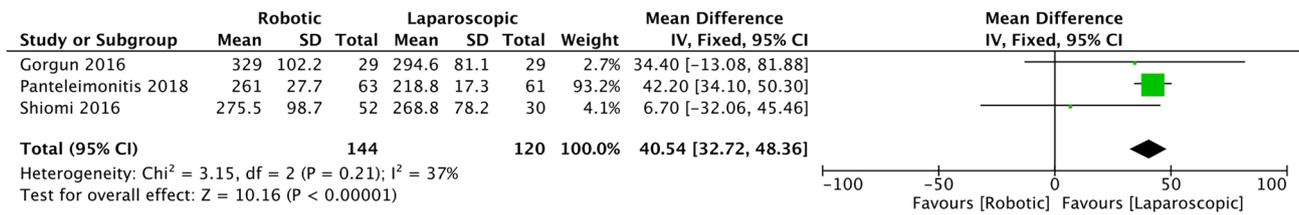


Fig. 4 Forest plot of length of operative time in comparison 1

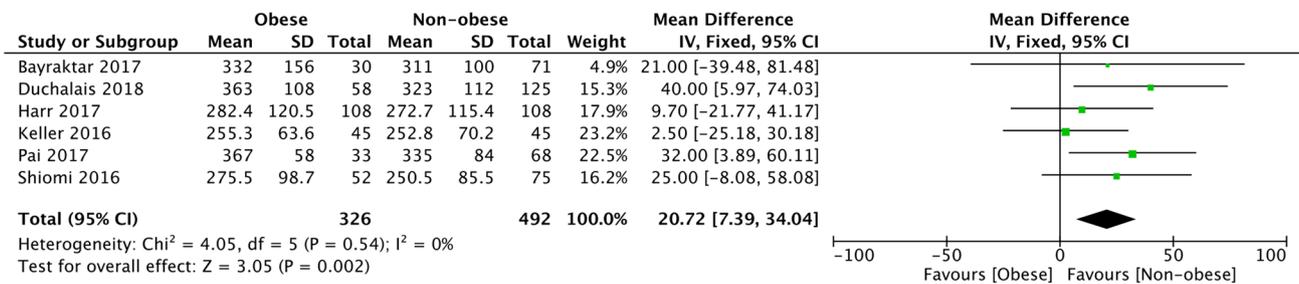


Fig. 5 Forest plot of length of operative time in comparison 2

$I^2 = 0\%$ ) were similar between both groups (Supplementary Figs. 2–4). Consistent findings were demonstrated in the sensitivity analysis.

### Second comparison: obese versus non-obese for participants who underwent RS

An increased operative time was noted in the obese group as compared to the non-obese group (MD 20.72 min, 95%CI 7.39–34.04 min,  $P = 0.002$ ,  $I^2 = 0\%$ ) (Fig. 5). In addition, the obese group had a significantly increased risk of wound infections (RR 2.59, 95%CI 1.12–6.02,

$P = 0.030$ ,  $I^2 = 0\%$ ) (Fig. 6). However, there were no significant differences in other intra- and post-operative outcomes including the risk of conversion (RR 1.12, 95%CI 0.56–2.27,  $P = 0.740$ ,  $I^2 = 0\%$ ), amount of blood loss (MD  $-0.76$  mL, 95%CI  $-13.57$  to  $12.05$  mL,  $P = 0.910$ ,  $I^2 = 0\%$ ), LOS (MD 0.48 days, 95%CI  $-0.33$  to  $-1.30$  days,  $P = 0.240$ ,  $I^2 = 74\%$ ), risk of ileus (RR 1.43, 95%CI 0.90–2.26,  $P = 0.130$ ,  $I^2 = 41\%$ ), pulmonary complications (RR 1.57, 95%CI 0.36–6.84,  $P = 0.550$ ,  $I^2 = 18\%$ ), anastomotic leak (RR 0.85, 95%CI 0.43–1.67,  $P = 0.630$ ,  $I^2 = 0\%$ ), and re-admission rates (RR 1.20, 95%CI 0.71–2.05,  $P = 0.500$ ,  $I^2 = 0\%$ ) (Supplementary Figs. 5–11). Consistent findings were demonstrated in the

sensitivity analysis. A summary of outcomes can be found in Table 2.

### Meta-regression analysis

A meta-regression analysis was performed to explore sources of heterogeneity in the risk of wound infection. The analysis revealed that the type of pathology (rectal, colon, both) ( $P = 0.255$ ), patient age ( $P = 0.530$ ), proportion of male gender ( $P = 0.279$ ), and continent ( $P = 0.583$ ) were not significant sources of heterogeneity.

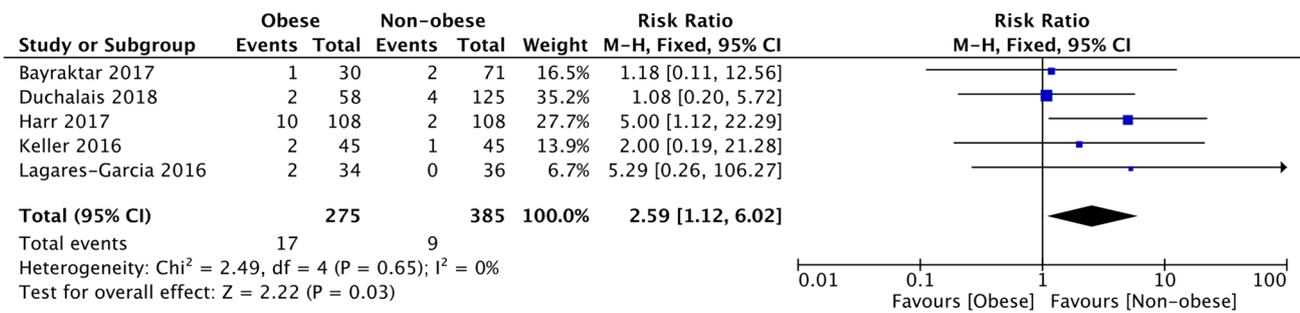


Fig. 6 Forest plot of wound infection rates in comparison 2

Table 2 Summary of outcomes

No.	Outcome	Studies/participants (n)	Effect estimate	P value	I <sup>2</sup> value (%)
<b>Robotic versus laparoscopic for all obese patients</b>					
1.	Conversion	3/262	RR 0.19, 95%CI 0.02–1.05	0.06	0
2.	Blood loss	2/262	MD $-30.71$ , 95%CI $-120.66$ to $59.25$	0.50	87
3.	Operative time	3/264	MD 40.54, 95%CI 32.72 to 48.36	<b>&lt;0.00001</b>	37
4.	Length of hospital stay	3/262	MD $-2.55$ , 95%CI $-3.13$ to $-1.97$	<b>&lt;0.00001</b>	26
5.	Anastomotic leak	3/262	RR 1.65, 95%CI 0.26–10.55	0.59	0
6.	Re-operation	2/180	RR 0.52, 95%CI 0.12–2.39	0.38	0
7.	Re-admission	2/180	RR 0.42, 95%CI 0.19–0.92	<b>0.03</b>	0
<b>Obese versus non-obese for all robotic surgery</b>					
1.	Conversion	7/888	RR 1.12, 95%CI 0.56–2.27	0.74	0
2.	Blood loss	6/287	MD $-0.76$ , 95%CI $-13.57$ to $12.05$	0.91	0
3.	Operative time	6/818	MD 20.72, 95%CI 7.39–34.04	<b>0.002</b>	0
4.	Length of hospital stay	7/888	MD 0.48, 95%CI $-0.33$ to $-1.30$	0.24	74
5.	Wound infection	5/660	RR 2.59, 95%CI 1.12–6.02	<b>0.03</b>	0
6.	Ileus	6/761	RR 1.43, 95%CI 0.90–2.26	0.13	41
7.	Pulmonary complications	2/284	RR 1.57, 95%CI 0.36–6.84	0.55	18
8.	Anastomotic leak	6/761	RR 0.85, 95%CI 0.43–1.67	0.63	0
9.	Re-admission	4/559	RR 1.20, 95%CI 0.71–2.05	0.50	0

Bold P values indicate statistical significance

RR relative risk, MD mean difference

## Discussion

Obesity has been shown to be a significant risk factor for increased morbidity, conversion rate, and blood loss during conventional laparoscopic colorectal surgery [21, 22]. Given the rising prevalence of obesity throughout the world [23, 24], operative techniques need to be developed to ensure that surgical outcomes remain consistent with those reported in non-obese patients. In the landmark ROLARR trial, the odds of conversion in obese patients were significantly higher compared to underweight or normal weight patients (adjusted OR = 4.69 [95%CI 2.08–10.58];  $P < 0.001$ ), with RS showing a trend towards a lower conversion rate compared to LS in the obese group of patients (18.9% vs. 27.8%, adjusted OR = 0.583 [95%CI 0.212–1.602];  $P = 0.29$ ) [25]. A similar finding was demonstrated in the current meta-analysis. In addition, the results suggest that RS provides earlier recovery with shorter LOS and reduced re-admission rates, without compromising on other operative outcomes. This could be due to the earlier recovery of gastrointestinal function after RS, as well as the incorporation of ERAS pathways in 2 of 3 studies [16, 20, 26, 27]. Other outcomes including blood loss, anastomotic leak, and re-operation rates were comparable between the two minimally invasive approaches. Nonetheless, the small number of studies included in this systematic review highlights the paucity of evidence available [16, 20].

While RS was associated with a longer operative time, some authors have attributed this to the protracted docking time associated with the older robotic systems [16], which often involved complicated docking procedures and port placement configurations [28]. With the newer da Vinci Xi system, however, docking can be completed faster with the boom-mounted robotic arms and laser targeting system [28]. In addition, the multi-quadrant capabilities of the Xi also help to shorten the operative duration by reducing the need for re-docking or hybrid procedures [29]. Another argument in favour of RS would be that most surgeons in the included studies were experienced in LS but were still in their learning phase for RS [12, 16, 20], potentially introducing bias in favour of the LS group. The actual benefit of RS could therefore be greater than what has been reported in the current literature, especially when performed by experienced robotic surgeons.

Although the main focus of this systematic review was intra- and post-operative surgical outcomes, cost remains an important factor, especially since the length of operative time was longer with RS, and robotic instruments and accessories may cost more [30–32]. However, emerging evidence is demonstrating that robotics may be more cost-effective than laparoscopy and open resection if certain

thresholds in quality of life, instrument costs, operating room time, and complication rates are achieved [32]. The other paramount consideration is oncological outcome. While current evidence suggests that RS is comparable to LS in terms of long-term oncological outcomes [18, 33–36], evidence pertaining to the obese population remains limited [16]. Since obesity was previously shown to be a risk factor for poorer oncological outcomes after surgery in other cancers such as prostate [37, 38], further research is warranted for colorectal cancer surgery in this group of patients.

Nonetheless, the results from this meta-analysis must be interpreted in the context of known limitations of retrospective cohort studies, such as selection and confounding bias. While the quality of studies were graded using the NOS tool, other residual confounders such as pathology, obesity-associated comorbidities, age, and gender, could not be accounted for using meta-regression analysis given the limited number of studies. Next, studies with pre-specified subgroups of obese patients [25] were excluded since outcomes reported may not be stratified according to the obese cohort. However, this may potentially result in a smaller sample size. Given the large number of meta-analyses performed in this study, there may also be an issue of multiple testing, which we have attempted to address by classifying complications into primary or secondary. While BMI is a commonly used method for defining obesity, other indices such as visceral fat measurements (VFA), waist and hip circumference, and waist-to-hip ratio may be more relevant when comparing surgical outcomes. Given the difference in fat distribution between Asian and Western populations, it might be more accurate to use the VFA for studies on the former [39]. In addition, as obesity was assessed on a dichotomous scale ( $30 \text{ kg/m}^2$ ), it was not possible to ascertain the extent of obesity at which surgery becomes more challenging. Conversely, patients who were underweight could also have been associated with poorer surgical outcomes [40], but given the lack of data, we were not able to stratify between normal and underweight patients in the non-obese group. Another potential bias would be heterogeneity in the types (benign vs. malignant) and location (rectal vs. colon) of disease conditions being studied, which may confound the overall outcomes and complication risks for obese patients. Unfortunately, subgroup analyses were not feasible as the data reported were not stratified according to these subgroups. Lastly, the lack of long-term follow-up data precluded the evaluation of oncologic outcomes.

## Compliance with ethical standards

**Disclosures** Ian Jun Yan Wee, Li-Jen Kuo, and James Chi-Yong Ngu have no conflicts of interest or financial ties to disclose.

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