



# Hyperhomocysteinemia and risk of incident cognitive outcomes: An updated dose-response meta-analysis of prospective cohort studies

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## ABSTRACT

**Objective:** This study aimed to comprehensively assess the dose-response relationship between blood homocysteine levels and risk of all cause, Alzheimer and vascular dementia, as well as cognitive impairment without dementia (CIND).

**Method:** We searched for all related prospective cohort studies reporting homocysteine as an exposure from patients with cognitive disorders as a result in the PubMed and EMBASE databases up to June 18, 2018. Pooled relative risks (RRs) and corresponding 95% confidence intervals (CIs) were extracted. The dose-response meta-analyses were conducted to assess potential linear and non-linear dose-response relations. Summary RRs and 95% CIs were calculated using a random- or fixed-effects model.

**Results:** Twenty-eight prospective cohort studies were eligible in this meta-analysis. During average follow-up periods ranging from 2.7 to 35 years there were 2557 cases (1035 all-cause dementia, 530 Alzheimer's disease, 92 vascular dementia and > 900 CIND) among 28,257 participants. There was a clear linear dose-response relationship between blood homocysteine concentration and risk of Alzheimer-type dementia ( $P > 0.05$  for non-linearity). The pooled RR of Alzheimer-type dementia was 1.15 (95% CI: 1.04 to 1.26;  $I^2 = 56.6\%$ ,  $n = 5$ ) for every 5  $\mu\text{mol/L}$  increase in blood homocysteine. Sensitivity analysis showed similar results, and there was no clear evidence of publication bias with Begg's and Egger's tests for Alzheimer dementia ( $P = 0.806$ ,  $0.084$ , respectively), strengthening the linear relationship between blood homocysteine levels and risk of Alzheimer dementia. Due to the presence of publication bias and low statistical power, elevated levels of blood homocysteine were not appreciably associated with risk of all-cause, vascular dementia and CIND.

**Conclusions:** Every 5  $\mu\text{mol/L}$  increase in blood homocysteine is linearly associated with a 15% increase in relative risk of Alzheimer-type dementia. This meta-analysis provides further evidence that a higher concentration of blood homocysteine is associated with a higher risk of Alzheimer-type dementia.

## 1. Introduction

Age-related cognitive decline or impairment is a major public health problem, affecting about 20% of people aged 70 years and older in the United State (Plassman et al., 2008). The prevalence of dementia, as a severe cognitive problem, increases with age, so that at the age of 80 years, about one in eight people are affected (Wald et al., 2011). To search for effective prevention and treatment strategies, it is important to identify the causes of cognitive deterioration (Scarmeas et al., 2018), and critical to develop an approach to treat or delay cognitive decline (Dolgin, 2016). Furthermore, novel preventive approaches focused on modifying risk factors (Zhou and Haina, 2017) for cognitive disorders are urgently needed to combat this growing epidemic. As a promising molecule for treating or preventing nervous diseases, high blood levels

of homocysteine (Hcy) have been closely associated with several diseases that affect the central nervous system, such as epilepsy (Elliott et al., 2007) and stroke (Lehotsky et al., 2016).

It has been 20 years since two case-control studies (Clarke et al., 1998; McCaddon et al., 1998) found that elevated blood total Hcy levels were associated with Alzheimer's disease (AD). Observational data suggested a link between hyperhomocysteinemia (HHcy) and increased risk of cognitive disorders such as Alzheimer's dementia (Ravaglia et al., 2005; Zylberstein et al., 2011), vascular dementia (Miwa et al., 2016) and cognitive impairment/decline (Haan et al., 2007; Nurk et al., 2005). Individual studies, however, have provided conflicting estimates (Ford et al., 2012; Hooshmand et al., 2010; Luchsinger et al., 2004; Miwa et al., 2016) of the strength of the association between Hcy and a range of cognitive disorders and have not agreed on whether there are

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relevant associations at all, possibly because of the small sample size examined. Furthermore, the knowledge and understanding of the clinical importance and implications of these associations are limited, and the broad range of literature needs to be reviewed comprehensively to characterize the associations of HHcy with different cognitive outcomes. In addition, previous meta-analyses of case-control and cohort studies on this topic found a significant positive relation between blood Hcy levels and risk of cognitive disorders (Van Dam and Van Gool, 2009; Wald et al., 2011). However, no dose-response analyses were conducted, thus questions about the strength and shape of the dose-response relationship between blood Hcy levels and risk of cognitive disorders remain to be addressed.

Given the inconsistency in the literature regarding the role of Hcy in risk of cognitive disorders, we conducted a meta-analysis to review current evidence on the associations of blood Hcy levels with incident risk of all-cause, Alzheimer-type, and vascular dementia, as well as cognitive impairment without dementia (CIND). The present meta-analysis was undertaken to provide an updated, more comprehensive and dose-response review about the relation between blood Hcy and risk of cognitive problems ranging from slight decline to dementia.

## 2. Methods

### 2.1. Search strategy

Following the guidelines by the MOOSE (Meta-analysis Of Observational Studies in Epidemiology) statement (Stroup et al., 2000), we searched the electronic databases (PubMed and EMBASE) from inception to June 18, 2018 using the following terms: homocysteine, hyperhomocysteinemia; blood, plasma, serum, circulat\*; dementia, Alzheimer\*, cognit\*; prospective, cohort, follow up, inciden\*, longitudinal, “nested case” (Details of search strategies are shown in Table S1, in Appendix 1). A list of the excluded studies is provided in table S2 in appendix 1. No language restrictions were imposed. Bibliographies of eligible studies and relevant meta-analyses were hand-searched for potential missing studies (Fig. 1).

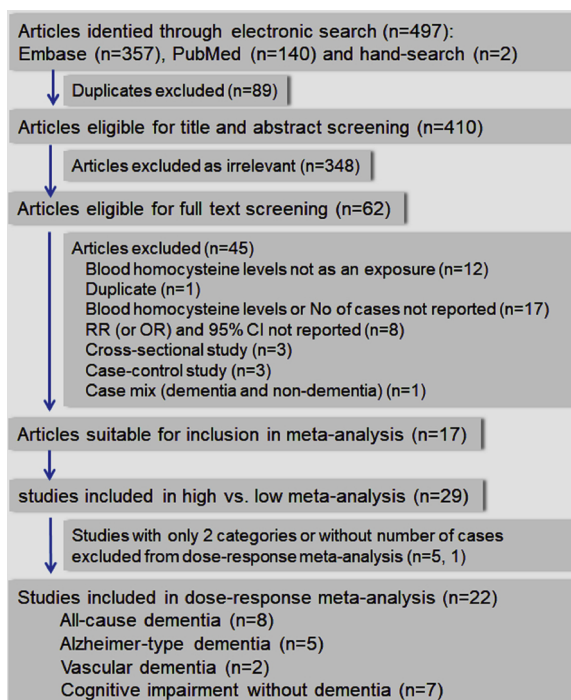


Fig. 1. Screening and selection process of studies investigating effects of blood Hcy concentration on risk of cognitive diseases.

### 2.2. Study selection

Studies were included if they were prospective cohort or prospective nested case-control studies, investigated an association between blood Hcy levels and cognitive disorders (All-cause dementia, or Alzheimer's disease, or vascular dementia, or cognitive impairment, or cognitive decline or cognitive deficit), classified blood Hcy concentrations into two or more categories, and reported adjusted risk estimates. For the dose-response analysis, the level-specific case numbers and person-years or sufficient data for deriving these numbers were required. The inclusion decisions were made independently by two reviewers (Zhou FT and Chen SR) and any disagreements were resolved by consensus after discussion.

### 2.3. Data extraction and quality evaluation

For each study included, we extracted the first author's last name, publication year, region (or country), cohort name, gender distribution (% female), mean age or age range, mean follow-up duration, sample size, number of cases, and person-years stratified by blood Hcy dose, cognitive outcomes, diagnosis criteria of cognitive disorders, sample source, method of measuring blood Hcy concentration, categories of blood Hcy, adjusted covariates, and multivariable-adjusted effects (RR and 95% CI) for each exposure category. The study quality was evaluated with the Newcastle-Ottawa Quality Assessment Scale (NOS), the quality score ranged from 0 to 9. Details of how the criteria were applied are shown in Table S3, in Appendix 1.

### 2.4. Statistical methods

In this meta-analysis, all associations were estimated as RRs and 95% CIs; HRs were considered equivalent to the RR (Xu et al., 2015). The ORs were transformed into RRs using the formula  $RR = OR / [(1 - P_0) + (P_0 \times OR)]$  where  $P_0$  is the incidence of the outcome of interest in the non-exposed group. Some studies reported the odds ratio (OR) or hazard ratio (HR) in each category, and the OR (or HR) was considered equivalent to the RR in cohort studies if the value of  $P_0$  was small (Zhang and Yu, 1998). For each of the included studies, we assigned the reported median or mean blood Hcy concentration of each category as the category of blood Hcy concentration. When a study reported only the range of blood Hcy levels for a category, we used the average value of the lower and upper bounds. When the highest category was open-ended, we assigned the lower end value of the category multiplied by 1.5. When the lowest category was open-ended, we set the lower boundary concentration to a fixed value of 3.0 because the lower limit of blood Hcy is normally around 3.0  $\mu\text{mol/L}$  (Zylberstein et al., 2011). The risk estimate from the most fully adjusted models in the analysis of the pooled RR was used. If the number of cases/non-cases in each category was not available and the authors did not give their reply, a method (Bekkering et al., 2008) was used to provide approximate data based on the total number and RRs of each category. We excluded the studies without the number of participants and/or cases in the whole cohort, also not providing RRs (ORs) and 95% CI, or without sufficient data to calculate the values required for the dose-response analyses.

We first summarized the RRs for the highest versus lowest category of blood Hcy levels in the included studies using the random effects ( $I^2 > 50\%$ , the DerSimonian-Laird method) or fixed effects ( $I^2 \leq 50\%$ , the Mantel-Haenszel method) meta-analysis (high v low meta-analysis). For a dose-response meta-analysis, we used the 2-stage generalized least-squares trend estimation method to estimate the study-specific slope lines first and then derive an overall average slope using the method described by Greenland and Longnecker (Greenland and Longnecker, 1992). We performed a dose-response meta-analysis to examine a potential nonlinear relationship between blood Hcy levels and risk of cognitive disorders by using restricted cubic splines. Restricted cubic spline models with 3 knots were fitted in each study

**Table 1**  
Characteristics of the identified prospective studies of blood Hcy levels and risk of cognitive disorders.

Author; year; country	Cohort name	baseline age (mean $\pm$ SD or range)	Female (%)	No of participants	Mean follow-up time	Diagnosis criteria	Adjusted confounders	Sample source	Quantification method of Hcy	Disease type
Ford et al., 2012; Australia	Health in Men Study cohort	67.4 $\pm$ 4.7 (62.7– 72.1)	0	2959	5.8	ICD-10	Age, history of ischaemic heart disease and of stroke	Plasma	HPLC	All-cause dementia
Hendrie et al., 2013 <sup>a</sup> ;USA	Indianapolis-Ibadan dementia project (African Americans)	76.75 $\pm$ 5.08	70.4	912	4.7	DSM-III-R/CD-10	Age, education, APOE- $\epsilon$ 4, smoking, time of enrollment	Plasma	NR	
Hendrie et al., 2013 <sup>b</sup> ; USA	Indianapolis-Ibadan dementia project (Yoruba)	75.88 $\pm$ 4.87	64.1	819	5.1	DSM-III-R/ICD-10	Age, education, APOE- $\epsilon$ 4, smoking, time of enrollment	Plasma	NR	
Kivipelto et al., 2009; Sweden	Kungsholmen Project	81 $\pm$ 4.6	74.6	213	6.7	DSM-III-R	Age, sex, education, BMI, albumin, Hb, creatinine, APOE- $\epsilon$ 4, MMSE score, holo-TC, B12 and folate	Plasma	IMx assay	
Miwa et al., 2016; Japan	Japanese cohort of participants with vascular risk factors	67.2 $\pm$ 8.4	41	643	7.3	DSM-III-R	Age, sex, education, APOE- $\epsilon$ 4, BMI, MMSE, hypertension, previous cerebrovascular events, eGFR and MRI-findings	Plasma	HPLC	
Ravaglia et al., 2005; Italian	CSBA	73.6 $\pm$ 6.3	53.2	816	3.8	DSM-IV	Age, sex, education, APOE- $\epsilon$ 4, serum creatinine, serum folate, serum Vit B12, history of stroke	Plasma	IMx assay	
Ravaglia et al., 2007; Italian	CSBA	73.6 $\pm$ 6.3	53.1	804	3.7	DSM-IV	Age, gender, education, APOE- $\epsilon$ 4, history of cardiovascular disease, history of stroke, physical activity, BMI, serum creatinine, serum folate, serum Vitamin B12	Plasma	IMx assay	
Seshadri et al., 2002; UK	Framingham Study cohort	76	61.1	1092	8	NINCDS	Age, sex, APOE- $\epsilon$ 4, educational status, history of stroke, smoking status, alcohol intake, diabetes mellitus, BMI, SDP	Plasma	HPLC	
Whalley et al., 2014; UK	Aberdeen 1921 Birth Cohort study	78 (77–79)	44.3	173	5	ICD-10	Sex, IQ, education, deprivation, plasma vita B12, folate, heart diseases, hypertension and plasma micronutrients	Plasma	HPLC	
Zylberstein et al., 2011; Sweden	Prospective Population Study of Women in Gothenburg	46.8	100	1368	35	DSM-III-R	Age, education, BMI, cholesterol, triglycerides, SBP/DBP, smoking, creatinine, vit B12	Serum	IMx assay	
Luchsinger et al., 2004; USA	Washington Heights- Inwood Columbia Aging Project	76.2 $\pm$ 5.7	70.7	679	4.72	NINCDS	Age, sex, education, APOE- $\epsilon$ 4, stroke	Plasma	HPLC	Alzheimer-type dementia
Miwa et al., 2016; Japan	Japanese cohort of participants with vascular risk factors	67.2 $\pm$ 8.4	41	643	7.3	DSM-IV	Age, sex, education, APOE- $\epsilon$ 4, BMI, MMSE	Plasma	HPLC	
Ravaglia et al., 2005; Italian	CSBA	73.6 $\pm$ 6.3	53.2	816	3.8	NINCDS-ADRDA	Age, sex, education, APOE- $\epsilon$ 4, serum creatinine, serum folate, serum vitamin B-12, history of stroke	Plasma	IMx assay	
Zylberstein et al., 2011; Sweden	Prospective Population Study of Women in Gothenburg	46.8	100	1368	35	DSM-III-R	Age, education, BMI, cholesterol, triglycerides, SBP/DBP, smoking, creatinine and vita B12, without cerebrovascular disease	Serum	IMx assay	
Hooshmand et al., 2010; Sweden/ Finland	CAIDE study	70.7 $\pm$ 3.6	62	271	7	NINCDSADRDA	Age, sex, education, and duration of follow- up, APOE- $\epsilon$ 4, BMI, MMSE, SBP/DBP, smoking, history of stroke.	Serum	Chemi-luminescent microparticle immunoassay	
Kivipelto et al., 2009; Sweden	Kungsholmen Project	81 $\pm$ 4.6	74.6	213	6.7	DSM-III-R	Age, sex, education, BMI, albumin, haemoglobin, creatinine, APOE- $\epsilon$ 4, MMSE, holo-TC, B12, folate	Plasma	IMx assay	
Ravaglia et al., 2007; Italian	CSBA	73.6 $\pm$ 6.3	53.1	804	3.7	NINCDSADRDA	Age, sex, education, APOE- $\epsilon$ 4, history of cardiovascular disease, history of stroke, physical activity, BMI, serum creatinine, serum folate, serum Vita B12	Plasma	IMx assay	

(continued on next page)

Table 1 (continued)

Author; year; country	Cohort name	baseline age (mean $\pm$ SD or range)	Female (%)	No of participants	Mean follow-up time	Diagnosis criteria	Adjusted confounders	Sample source	Quantification method of Hcy	Disease type
Seshadri et al., 2002; USA	Framingham Study cohort	76	61.1	1092	8	NINCDS	Age, sex, APOE- $\epsilon$ 4, education, history of stroke, smoking status, alcohol intake, diabetes mellitus, BMI, SDP	Plasma	HPLC	
Miwa et al., 2016; Japan	Japanese cohort of participants with vascular risk factors	67.2 $\pm$ 8.4	41	643	7.3	DSM-IV	Age, sex, education level, APOE- $\epsilon$ 4, BMI, MMSE	Plasma	HPLC	Vascular dementia
Zylberstein et al., 2011; Sweden	The Prospective Population Study of Women in Gothenburg	46.8	100	1368	35	DSM-III-R	Age, education, BMI, cholesterol, triglycerides, SBP/DBP, smoking, creatinine, Vit B12	Serum	IMx assay	
Ravaglia et al., 2007; Italian	CSBA	73.6 $\pm$ 6.3	53.1	804	3.7	NINCDS-AIREN	Age, sex, education, APOE- $\epsilon$ 4, history of cardiovascular disease, history of stroke, physical activity, BMI, serum creatinine, serum folate, serum Vit B12	Plasma	IMx assay	
Kalmijn et al., 1999; Netherlands	Prospective Rotterdam Study	67.7 $\pm$ 7.1	60	819	2.7	MMSE; the MMSE score of > 1 point /year	Age, sex, education, and baseline MMSE	Serum	HPLC	Cognitive impairment/ decline
Reitz et al., 2009; USA	all-cause MCI	77.4 $\pm$ 5.8	70.4	516	5.2	DSM-IV	Age, gender, ethnic group, APOE- $\epsilon$ 4	Plasma	HPLC	
Reitz et al., 2009; USA	amnestic MCI	77.4 $\pm$ 5.8	70.4	432	5.2	DSM-IV	Age, gender, ethnic group, APOE- $\epsilon$ 4	Plasma	HPLC	
Reitz et al., 2009; USA	non-amnestic MCI	77.4 $\pm$ 5.8	70.4	467	5.2	DSM-IV	Age, gender, ethnic group, APOE- $\epsilon$ 4	Plasma	HPLC	
Mendonca et al., 2017; UK	Newcastle 85+ Study	> 85	61	763	5	sMMSE	Alcohol intake, smoking status, APOE- $\epsilon$ 4 (rs429358/rs7412), sex, education, BMI, depression, hypertension, diabetes type 1/2, history of cardiovascular diseases, physical activity.	Plasma	IMx assay	
Nurk et al., 2005; Norway	Hordaland Homocysteine Study	66	0	2189	6	ICD-9	Sex, APOE- $\epsilon$ 4, education, history of cardiovascular disease and hypertension, depression score	Plasma	HPLC	
Kado et al., 2005; USA	MacArthur Studies of Successful Aging	74 $\pm$ 2.7	58	370	7	Five standardized cognitive performance tests	Age, sex, and baseline cognitive function	Plasma	HPLC	
Dufouil et al., 2003; France	Epidemiology of Vascular Ageing (EVA) study	67 $\pm$ 3	58.6	1107	2	MMSE	Age, gender, education level, baseline cognition, BMI, alcohol consumption, smoking status, hypertension, hypercholesterolemia, glycemic status, history of vascular disease, folate and vitamin B12 concentrations	Plasma	HPLC	

NR = not reported; HPLC = high-performance liquid chromatography assay; ICD-9/10 = the International International Classification of Diseases, Ninth/Tenth Revision; DSM-III = the Diagnostic and Statistical Manual of Mental Disorders 3rd edition; NINCDS-ADRDA = National Institute of Neurologic and Cognitive Disorders and Stroke-AD and Related Disorders Association criteria; NINCDS-AIREN = National Institute of Neurological and Communicative Disease and Stroke-Alzheimer's Disease and Related Disorders Association; BMI = body mass index; APOE- $\epsilon$ 4 = apolipoprotein E- $\epsilon$ 4; MMSE = Mini-mental State Examination; MCI = mild cognitive impairment; CSBA = Conselice Study of Brain Aging; SBP = systolic blood pressure; DBP = diastolic blood pressure.

taking into account the covariance among log RR, and the regression coefficients were then combined using multivariate meta-analysis. A test for a non-linear relation was calculated by making the coefficient of the second spline equal to zero, as described previously (Orsini et al., 2012). When the number of studies reporting a specific outcome was small, we did not carry out a dose-response meta-analysis. To generate a linear dose-response curve, data on the blood Hcy levels, the distribution of cases and person-years, and RRs plus 95% CIs for 3 or more categories were extracted. First, specific linear trends and 95% CIs were estimated from the natural logs of RRs across categories of Hcy by the generalized least-square models method. Then, the estimated linear trends were pooled with fixed- or random-effects meta-analysis, depending on the absence or presence of statistical heterogeneity. If the nonlinearity was not statistically significant, the linear dose-response outcomes were presented in Hcy levels per 3- or 5-unit ( $\mu\text{mol/L}$ ) increase in the forest plots.

In addition to evaluating the entire cohort, we also performed stratified analyses according to study location, gender distribution (% female), diagnosis criteria, sample source (plasma or serum), method used to quantify blood Hcy levels, mean follow-up duration, study quality score, and adjustment for confounders (age, sex, BMI, APOE- $\epsilon 4$  status, education levels, B vitamins and history of cardiovascular diseases). Study heterogeneity was assessed using the Q-test and  $I^2$  statistic,  $p < 0.10$  and  $I^2 > 50\%$  indicated evidence of heterogeneity. If the  $I^2$  statistic was 50% or less, a meta-analysis based on a fixed-effect model was conducted, otherwise the random-effects model was used. Sensitivity analyses excluding one study at a time were conducted to explore whether the results were strongly influenced by a specific study. Potential publication bias was assessed by the application of contour-enhanced funnel plots (Peters et al., 2008), Egger's linear regression test, and Begg's rank correlation test. If publication bias was present, we further evaluated the number of missing studies using the trim-fill method and re-calculated the pooled risk estimates with the addition of those missing studies. All statistical analyses were conducted with two-tailed test at the  $P < 0.05$  level for statistical significance using STATA v14.0 (Stata Corp, College Station, TX, USA).

### 3. Results

#### 3.1. Literature search

We identified 410 articles for review of title and abstract. After the initial screening, full text of potentially eligible articles was retrieved for detailed assessment. After full text reviews, 45 articles were excluded (see Fig. 1 and Table S2 in appendix 1), and 29 eligible cohort studies from 17 eligible articles were included for meta-analysis, with a total of 28,257 participants and 2557 patients with cognitive disorders (1035 cases of all-cause dementia, 530 cases of AD, 92 cases of vascular dementia, and 900 above cases of CIND). A study (Haan et al., 2007) did not independently calculate the values of the relations of dementia and CIND. We chose not to include (neither in the dementia group nor in the CIND group) it and thus made exclusion decision. All the studies included have been published as full manuscripts and are of high quality (see Table S3, appendix 1). Fig. 1 shows a flowchart of the study selection. In addition, the characteristics of the included studies summarize in Table 1, and summary statistics on the exposure and outcome variables are provided in Table 2.

#### 3.2. Study characteristics

As shown in Tables 1 and 2, there are 10 studies for all-cause dementia, 8 for Alzheimer-type dementia, and 3 for vascular dementia, 8 for CIND. Average follow-up periods ranged from 2.7 to 35 years. Of all the studies included, 8 studies were conducted from the United States, 17 from Europe, 3 from Asia, and 1 from Australia. Most of studies adjusted for age, sex and education levels. Some cohorts also controlled

for some conventional risk factors, including body mass index, APOE status, vitamin B status, and cardiovascular disease. One analysis from an article was based on two different nationalities (Yoruba and African Americans), which was considered as two independent studies (Hendrie et al., 2013).

#### 3.3. Blood Hcy levels and risk of all-cause dementia

Ten cohort studies in nine articles (Ford et al., 2012; Hendrie et al., 2013; Kivipelto et al., 2009; Miwa et al., 2016; Ravaglia et al., 2007, 2005; Seshadri et al., 2002; Whalley et al., 2014; Zylberstein et al., 2011) investigated the association between blood Hcy levels and risk of all-cause dementia, with a total of 11,168 participants and 1035 incident patients (Tables 1 and 2). The pooled RR for the highest versus lowest Hcy levels was 1.58 (95% CI: 1.33–1.87,  $I^2 = 17.4\%$ ,  $P_{\text{heterogeneity}} = 0.283$ ) (Fig. 2A). Begg's and Egger's tests indicated publication bias ( $P = 0.007$ ,  $0.011$ , respectively; see Table 3). We used the trim and fill method to recalculate the pooled risk estimate, and the analysis suggested that the imputed risk estimate was 1.471 (95% CI: 1.264–1.712), which is slightly decreased in risk but still identical to our original risk estimate.

For dose-response meta-analysis, we excluded two studies (Ravaglia et al., 2007; Seshadri et al., 2002) that divided Hcy concentration into only two categories, because this requires for at least three quantitative exposure categories. Therefore, this analysis included eight studies with a total of 9272 participants and 815 all-cause dementia cases. Using a restricted cubic splines model, we found no evidence of a curvilinear relationship between blood Hcy levels and risk of all-cause dementia ( $P = 0.443$  for non-linearity; Fig S6 in appendix 2). For linear dose-response analysis, the summary RR per  $3 \mu\text{mol/L}$  (unit) increase in blood Hcy level was 1.07 (95% CI: 1.03–1.12; Fig S2 in appendix 2), with moderate between-study heterogeneity ( $P = 0.034$ ,  $I^2 = 53.7\%$ ). The summary RR per 5 unit increase was elevated to a height of 1.12 (95% CI: 1.05–1.20) with similar heterogeneity ( $P = 0.03$ ,  $I^2 = 54.8\%$ ; Fig S5 in appendix 2). Begg's and Egger's tests indicated publication bias ( $P = 0.035$ ,  $0.012$ , respectively; see Table 3). The trim and fill method was used to re-calculate our pooled risk estimate. The analysis suggested that the imputed risk estimate was 1.064 (95% CI: 0.993–1.142) per 5-unit increase, 1.038 (95% CI: 0.996–1.083) per 3-unit increase, with no significance for the adjusted risk estimates.

#### 3.4. Blood Hcy levels and risk of Alzheimer-type dementia

The association between blood Hcy levels and risk of Alzheimer-type dementia was investigated in eight studies (Hooshmand et al., 2010; Kivipelto et al., 2009; Luchsinger et al., 2004; Miwa et al., 2016; Ravaglia et al., 2007, 2005; Seshadri et al., 2002; Zylberstein et al., 2011) with a total of 5777 participants and 530 patients with Alzheimer's disease (Tables 1 and 2). As shown in Table 3, the pooled RR of Alzheimer-type dementia for the highest versus lowest category of blood Hcy was 1.74 (95% CI 1.32–2.29), with significant heterogeneity ( $I^2 = 58.5\%$ ,  $P = 0.018$ ) (Fig. 2B). Begg's test indicated no publication bias ( $P = 0.063$ ), but Egger's test indicated publication bias ( $P = 0.004$ ; Table 3). We used the trim and fill method to re-calculate the pooled risk estimate. The analysis suggested that the imputed risk estimate was 1.372 (95% CI: 1.033–1.822), which is less than our original risk estimate, but its significance remains clear.

For dose-response meta-analysis, three studies (Hooshmand et al., 2010; Ravaglia et al., 2007; Seshadri et al., 2002) were excluded due to only two categories in Hcy levels. Therefore, in this analysis five studies were included with a total of 3610 participants and 362 Alzheimer-type dementia cases. Similarly, we observed that there was no significant non-linear relationship between blood Hcy levels and risk of Alzheimer-type dementia ( $P = 0.586$  for non-linearity; Fig. 3B). For linear dose-response analysis, the summary RR per 3-unit increase in Hcy was 1.09 (95% CI: 1.02–1.15; Fig S3 in appendix 2) with moderate between-



**Table 2**

Risk relative for cognitive disorders in studies included in systematic review and dose-response meta-analysis on blood homocysteine concentration ( $\mu\text{mol/L}$ ) and risk of cognitive disorders.

Author; year	Sample size	Case No	Person-years	Category	Hcy exposure	Mean	Multivariable-adjusted RR (95%CI)	Disease type
Ford et al., 2012	1033	43	5991.4	4	$\leq 10.3$	6.87	1 (ref)	All-cause dementia
	995	48	5771		NR	12.6	0.88 (0.57–1.36)	
	988	65	5730.4		NR	13.9	1.06 (0.70–1.61)	
	981	74	5689.8		$> 15.1$	22.65	1.2 (0.8–1.79)	
Hendrie et al., 2013 <sup>a</sup>	283	19	1330.1	4	4.44–10.71	7.58	1 (ref)	
	284	22	1334.8		10.72–16.99	13.86	1.16 (0.58–2.28)	
	284	32	1334.8		17.0–23.27	20.14	1.78 (0.94–3.38)	
	284	28	1334.8		23.28–29.52	26.4	1.41 (0.73–2.71)	
Hendrie et al., 2013 <sup>b</sup>	259	10	1320.9	4	3.56–11.28	7.42	1 (ref)	
	259	13	1320.9		11.29–19	15.15	1.27 (0.52–3.09)	
	259	14	1320.9		19.01–26.73	22.87	1.39 (0.57–3.38)	
	259	22	1320.9		26.74–34.46	30.6	2.19 (0.95–5.07)	
Kivipelto et al., 2009	53	19	355.1	4	5–8.7	6.85	1 (ref)	
	53	14	355.1		8.8–12.5	10.65	0.83 (0.39–1.77)	
	53	15	355.1		12.6–16.3	14.45	1.12 (0.51–2.42)	
	54	35	361.8		16.4–20	18.2	1.79 (0.86–3.74)	
Miwa et al., 2016	214	11	1562.2	3	$\leq 8.2$	5.47	1 (ref)	
	214	9	1562.2		8.3–10.7	9.5	0.78 (0.27–1.94)	
	215	27	1569.5		$\geq 10.8$	16.2	2.5 (1.01–6.63)	
Ravaglia et al., 2005	211	13	801.8	4	$< 10.1$	6.73	1 (ref)	
	204	23	775.2		10.1–12.5	11.3	1.7 (0.6–3.5)	
	184	21	699.2		12.6–15.0	13.8	2.1 (0.9–4.1)	
	217	55	824.6		$> 15.0$	22.5	3.5 (1.7–7.5)	
Ravaglia et al., 2007	590	56		2	$< 15$		1 (ref)	
	214	53			$> 15$		1.85 (1.13–3.02)	
Seshadri et al., 2002	321	55		2	$< 14$		1 (ref)	
	771	56			$> 14$		1.4 (1.1–1.9)	
Whalley et al., 2014	46	4	230	3	$< 10.8$	7.2	1 (ref)	
	49	11	245		10.8–14	12.4	2.63 (0.65–10.59)	
	78	17	390		$> 14$	21	3.72 (1.06–13.08)	
Zylberstein et al., 2011	254	39	8890	3	3.1–9.8	6.45	1 (ref)	
	564	49	19740		9.8–12.6	11.2	1.3 (0.84–2.0)	
	441	63	15435		12.6–78.9	45.75	1.67 (1.10–2.57)	
Luchsinger et al., 2004	177	26	835.44	4		10.75	1 (ref)	Alzheimer-type dementia
	184	29	868.48			14.09	1.10 (0.7–2.0)	
	164	24	774.08			17.52	1 (0.6–1.8)	
	154	30	726.88			27.44	1.3 (0.8–2.3)	
Miwa et al., 2016	214	1	1562.2	3	$\leq 8.2$	5.47	1 (ref)	
	214	1	1562.2		8.3–10.7	9.5	0.71 (0.09–4.49)	
	215	22	1569.5		$\geq 10.8$	16.2	3.31 (0.82–17.13)	
Ravaglia et al., 2005	211	8	801.8	4	$< 10.1$	6.73	1 (ref)	
	204	17	775.2		10.1–12.5	11.3	2.5 (0.93–6.0)	
	184	14	699.2		12.6–15.0	13.8	2.5 (0.93–6.1)	
	217	31	824.6		$> 15.0$	22.5	4.2 (1.7–11.0)	
Zylberstein et al., 2011	254	12	8890	3	3.1–9.8	6.45	1 (ref)	
	564	40	19740		9.8–12.6	11.2	1.54 (0.78–3.05)	
	441	46	15435		12.6–78.9	45.75	2.43 (1.25–4.71)	
Hooshmand et al., 2010	271	17		2	$\leq 12.3$		1 (ref)	
					$> 12.3$		1.19 (1.01–1.39)	
Kivipelto et al., 2009	53	13	355.1	4	5–8.7	6.85	1 (ref)	
	53	8	355.1		8.8–12.5	10.65	0.71 (0.26–1.91)	
	53	12	355.1		12.6–16.3	14.45	1.45 (0.57–3.69)	
	54	28	361.8		16.4–20	18.2	2.57 (1.06–6.24)	
Ravaglia et al., 2007	590	38		2	$< 15$		1 (ref)	
	214	30			$> 15$		1.91 (1.02–3.56)	
Seshadri et al., 2002	321	27		2	$< 14$		1 (ref)	
	771	56			$> 14$		1.6 (1.2–2.1)	
Miwa et al., 2016	214	3	1562.2	3	$\leq 8.2$	5.47	1 (ref)	Vascular dementia
	214	4	1562.2		8.3–10.7	9.5	1.29 (0.15–6.48)	
	215	14	1569.5		$\geq 10.8$	16.2	6.29 (1.10–18.42)	
Zylberstein et al., 2011	254	9	8890	3	3.1–9.8	6.45	1 (ref)	
	564	15	19740		9.8–12.6	11.2	0.76 (0.35–1.64)	
	441	13	15435		12.6–78.9	45.75	0.70 (0.28–1.72)	
Ravaglia et al., 2007	590	15	2183	2	$< 15$		1 (ref)	
	214	19	791.8		$> 15$		1.76 (0.70–4.23)	

(continued on next page)

Table 2 (continued)

Author; year	Sample size	Case No	Person-years	Category	Hcy exposure	Mean	Multivariable-adjusted RR (95%CI)	Disease type
Kalmijn et al., 1999	248	49	669.6	3	< 12.9	8.6	1 (ref)	CIND
	262	72	707.4		12.9–15.8	14.35	1.04 (0.73–1.49)	
	309	106	834.3		> 15.8	23.7	1.01 (0.70–1.44)	
Reitz et al., 2009 <sup>a</sup>	172	44	894.4	3	7.5–14.0	10.75	1 (ref)	
	173	49	899.6		14.1–21.4	17.75	1 (0.73–1.36)	
	171	39	889.2		21.5–29.6	25.55	1 (0.72–1.39)	
Reitz et al., 2009 <sup>b</sup>	143	15	743.6	3	7.5–14.0	10.75	1 (ref)	
	140	17	728		14.1–21.4	17.75	1.09 (0.58–2.06)	
	149	17	774.8		21.5–29.6	25.55	1.26 (0.66–2.40)	
Reitz et al., 2009 <sup>c</sup>	157	29	816.4	3	7.5–14.0	10.75	1 (ref)	
	156	32	811.2		14.1–21.4	17.75	1.08 (0.60–1.95)	
	154	22	800.8		21.5–29.6	25.55	0.92 (0.57–1.46)	
Mendonca et al., 2017	189	17	945	4	< 13.5	9	1 (ref)	
	188	15	940		13.5–16.7	15.1	1.71 (0.68–4.30)	
	200	12	1000		16.7–21.4	19.05	0.97 (0.37–2.56)	
	186	13	930		> 21.4	32.1	1.81 (0.63–5.25)	
Kado et al., 2005	370	NR		2	< 8.87		1 (ref)	
					13.38–40		1.53 (1.0–2.4)	
Nurk et al., 2005	452	33	2712	5		8.3	1 (ref)	
	440	37	2640			10	1.05 (0.58–1.89)	
	472	50	2832			11.5	1.70 (1.01–2.88)	
	386	51	2316			13.3	1.66 (0.95–2.91)	
	432	63	2592			16.5	2.34 (1.39–3.91)	
Dufouil et al., 2003	1131	30	2262	4	< 10	6.5	1 (ref)	
	702	29	1404		10–11.9	10.95	1.57 (0.73–3.38)	
	694	24	1388		12–14.9	13.45	1.29 (0.55–3.02)	
	520	35	1040		≥ 15	22.5	2.67 (1.22–5.85)	

Hendrie et al., 2013<sup>a</sup>: cohorts of elderly African Americans; Hendrie et al., 2013<sup>b</sup>: cohorts of elderly Yoruba; Reitz et al., 2009<sup>a</sup>: cases were all-cause mild cognitive impairment (MCI); Reitz et al., 2009<sup>b</sup>: cases were amnesic MCI; Reitz et al., 2009<sup>c</sup>: cases were non-amnesic MCI; NR = not reported; CIND = cognitive impairment without dementia.

study heterogeneity ( $P = 0.056$ ,  $I^2 = 56.6\%$ ). The summary RR per 5  $\mu\text{mol/L}$  increase, pooled RR was elevated to 1.15 (95% CI: 1.04–1.26) with similar heterogeneity ( $P = 0.058$ ,  $I^2 = 56.6\%$ ; Fig. 3A). Begg's and Egger's tests indicated no publication bias ( $P = 0.806$ , 0.084, respectively; Table 3). Fig. 3 shows the results of non-linear dose-response meta-analysis, and every 3 or 5  $\mu\text{mol/L}$  increases in blood Hcy was estimated to be associated with a 9% or 15% higher risk of Alzheimer-type dementia, respectively (Table S4 in appendix 1).

### 3.5. Blood Hcy levels and risk of vascular dementia

For the association between blood Hcy levels and incident risk of vascular dementia were investigated in three studies (Miwa et al., 2016; Ravaglia et al., 2007; Zylberstein et al., 2011) with a total of 2706 participants and 92 patients with vascular dementia (Tables 1–3). As shown in Table 3, the pooled RR of vascular dementia for the highest versus lowest category of blood Hcy levels was 1.78 (95% CI 0.58–5.42; Fig S1 in appendix 2), with significant heterogeneity ( $I^2 = 70.6\%$ ,  $P = 0.033$ ). Although Begg's and Egger's tests indicated no publication bias ( $P = 1$ , 0.377, respectively), the statistical power was too low to draw definitive conclusions regarding the association. For dose-response meta-analysis, one study (Ravaglia et al., 2007) was excluded due to only two categories in Hcy levels. Therefore, this analysis included two studies with a total of 1902 participants and 58 vascular dementia cases. Because the number of the studies of the association between Hcy levels and risk of CIND was small, then we did not perform a nonlinear dose-response analysis (there was no significance for linear dose-response analysis; Table 3).

### 3.6. Blood Hcy levels and risk of cognitive impairment without dementia (CIND)

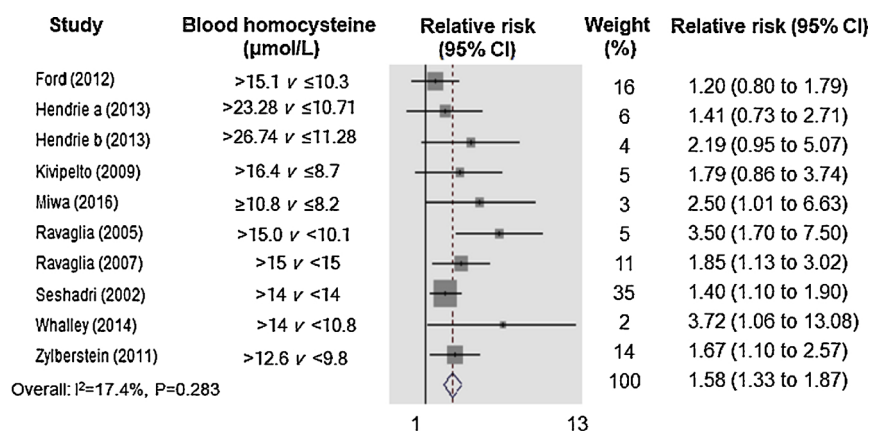
For the association between blood Hcy levels and incident risk of CIND, there were eight studies in six articles (Dufouil et al., 2003; Kado

et al., 2005; Kalmijn et al., 1999; Mendonca et al., 2017; Nurk et al., 2005; Reitz et al., 2009) included, with a total of 8606 participants and 900 and more patients with CIND (see Tables 1–3, a case number not reported in one study [31] included). As shown in Table 3, the pooled RR for the highest versus lowest category was 1.34 (95% CI 1.02–1.74), with small heterogeneity ( $I^2 = 53.7\%$ ,  $P = 0.033$  for heterogeneity; Fig. 3). There was no evidence of publication bias with Begg's and Egger's tests for CIND ( $P = 0.174$ , 0.097, respectively). For dose-response meta-analysis, one study (Kado et al., 2005) not reporting the number of cases could not be included in the dose-response analysis. Therefore, this analysis included seven studies with a total of 8226 participants and 900 cases. Using a restricted cubic splines model, it was shown that there was no significant non-linear relationship between blood Hcy levels and risk of CIND ( $P = 0.0974$  for non-linearity; Fig S8, in appendix 2). For linear dose-response analysis, the summary RR for a 3-unit increase in Hcy was 1.04 (95% CI: 1.00–1.08; Fig S4 in appendix 2), with moderate between-study heterogeneity ( $I^2 = 53.7\%$ ,  $P = 0.033$ ). For a 5-unit increase, the pooled RR was 1.06 (95% CI: 0.99–1.13;  $I^2 = 62\%$ ,  $P = 0.015$ ; Fig S7, in appendix 2). Begg's and Egger's tests indicated publication bias (both of  $P$  values: 0.034). We used the trim-fill method to re-calculate the pooled risk estimate. The analysis suggested, however, that the imputed risk estimate was 1.007 (95% CI: 0.960 to 1.056) for every 3-unit increase, 1.012 (95% CI: 0.935–1.095) for every 5-unit increase, with no significance for the adjusted risk estimates (see Table 3).

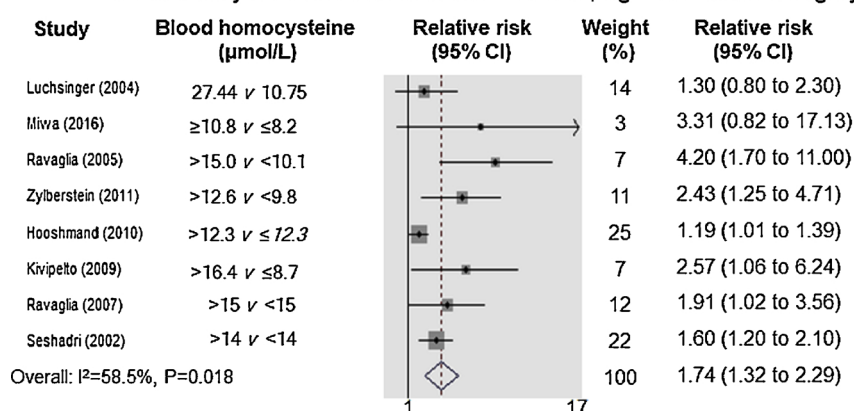
### 3.7. Study quality, subgroup analyses, and sensitivity analyses

Assessment of study quality yielded an average score of 7 (9 representing the highest quality), and 13 studies had a score of  $\geq 7$  (Table S3 in appendix 1). Mean (median) study quality scores were 7.4 for all-cause dementia, 7.1 for Alzheimer-type dementia, 7.3 for vascular dementia, 6.6 for CIND.

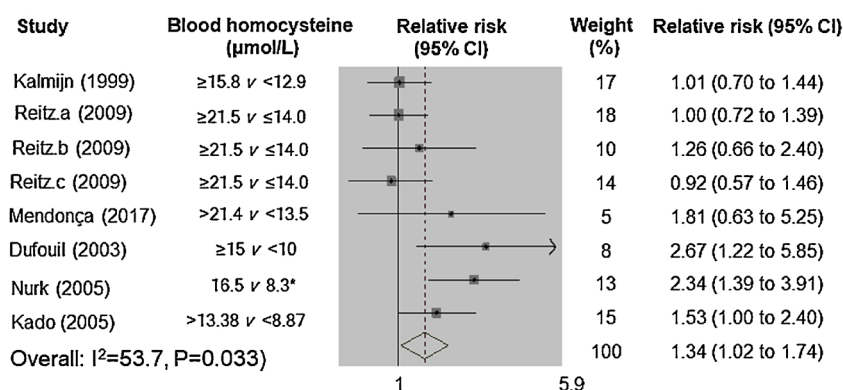
To evaluate the robustness of the risk estimates, several stratified



A. Summary relative risk of all-cause dementia, highest v lowest category



B. Summary relative risk of Alzheimer-type dementia, highest v lowest category



C. Summary relative risk of CIND, highest v lowest category

**Fig. 2.** Summary relative risk of all-cause (A, weights from fixed effects analysis) and Alzheimer-type dementia (B, weights from random effects analysis), as well as CIND (C, weights from fixed effects analysis), highest vs lowest blood Hcy category. \*indicates mean value.

analyses were done based on study location, gender distribution (% female), mean follow-up duration, diagnosis criteria, quantification method of Hcy, study quality and adjustment for critical confounders. Table 4 shows the different subgroup analyses. The positive associations between blood Hcy levels and risk of all-cause and Alzheimer-type dementia persisted in most of subgroup analyses. For all-cause dementia, most of the subgroups (mean follow-up duration, sample source, quantification method, NOS score) follow the overall trend and show statistically significant increases. Subgroup analyses showed no significant associations for studies reporting participants from beyond Europe, following the diagnosis criteria of ICD-10, and with small number of female participants (50% below) (Table 3), as well as with no adjustment for sex and education levels. We noted that in other

subsets associations were detected. For Alzheimer-type dementia, there were significant associations in most of the subgroup analyses, with exception of the serum subgroup, the subgroup with a low NOS score, and the subgroup including the studies (Table 2). For CIND, most of the subgroups demonstrated no significant relationship. Sensitivity analyses demonstrated that the estimates were not substantially altered for all-cause dementia, Alzheimer-type dementia and CIND (Figs S16–S24 in appendix 2).

#### 4. Discussion

To the best of our knowledge, the present meta-analysis is the largest and most comprehensive evaluation of the dose-response



**Table 3**  
Meta-analysis of blood Hcy levels and risk of cognitive disorders.

Comparison	No of studies	Cases/participants	Pooled RR (95% CI)	Heterogeneity ( $I^2$ ), P value	P values for Begg's and Egger's tests	Adjusted RR (95% CI) via the trim-and-fill method
<b>All-cause dementia</b>						
Highest v lowest	10	1,035/11,168	1.58 (1.33 to 1.87)	17.4, 0.263		1.471 (1.264 to 1.712)
per 3-unit increment	8	926/10,364	1.07 (1.03 to 1.12)	53.7, 0.034	0.007, 0.011	1.038 (0.996 to 1.083)
per 5-unit increment	8	926/10,364	1.12 (1.05 to 1.20)	53.7, 0.035	0.035, 0.012	1.064 (0.993 to 1.142)
<b>Alzheimer-type dementia</b>						
Highest v lowest	8	530/5,777	1.74 (1.32 to 2.29)	58.8, 0.018	0.063, 0.004	1.372 (1.033 to 1.822)
per 3-unit increment	5	362/3,610	1.09 (1.02 to 1.15)	56.6, 0.056	0.806, 0.081	
per 5-unit increment	5	362/3,610	1.15 (1.04 to 1.26)	56.6, 0.058	0.806, 0.084	
<b>Vascular dementia</b>						
Highest v lowest	3	92/2,706	1.78 (0.58 to 5.42)	70.6, 0.033	1, 0.377	
per 3-unit increment	2	58/1,902	1.18 (0.79 to 1.76)	90.1, 0.001		
per 5-unit increment	2	58/1,902	1.32 (0.68 to 2.58)	90.1, 0.001		
<b>CIND</b>						
Highest v lowest	8	900*/8,606	1.34 (1.02 to 1.74)	53.7, 0.033	0.174, 0.097	1.007 (0.960 to 1.056)
per 3-unit increment	7	900/8,226	1.04 (1.00 to 1.08)	62.0, 0.015	0.034, 0.034	1.012 (0.935 to 1.095)
per 5-unit increment	7	900/8,226	1.06 (0.99 to 1.13)	62.0, 0.015	0.035, 0.034	

RR = relative risk; CI = confidence interval; Hcy = homocysteine; CIND = cognitive impairment without dementia.  
\* indicates that p-value had a significant level ( $P < 0.05$ ) after the trim-and-fill correction.

relationships between Hcy levels and risks of incident cognitive disorders in the general population. There were positive associations between Hcy levels and risks of these disorders, including all-cause and Alzheimer-type dementia, and cognitive impairment without dementia, suggesting increases of 58%, 74%, and 34%, respectively. Nevertheless, there was a non-significant association for vascular dementia. Particularly importantly, there was a linear, dose-dependent relationship between Hcy levels (per 5 or 3  $\mu\text{mol/L}$  increase) and risk of incident Alzheimer-type dementia (a 15% or 9% increase in risk, respectively). The findings from the current meta-analysis of prospective cohort studies support the notion that an increased level of blood Hcy appears to play a causal role in the development of AD, but neither in the other dementia nor in cognitive impairment.

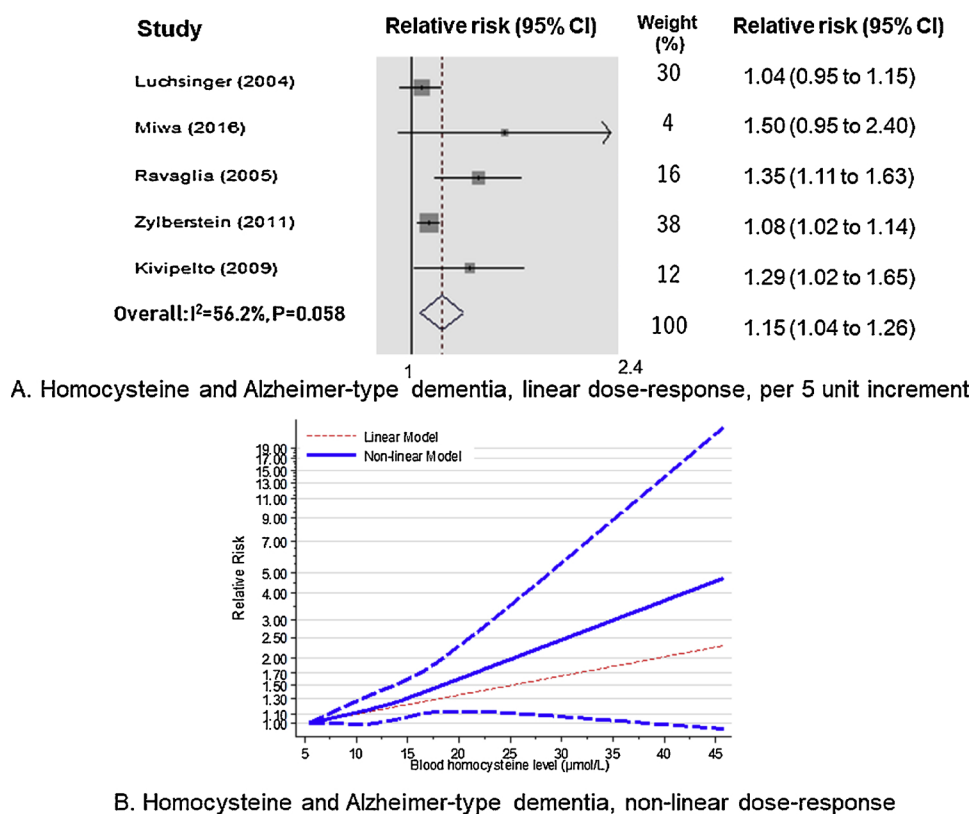
#### 4.1. Exploration of heterogeneity and publication bias

In the current meta-analysis of Alzheimer-type dementia, there was no between-study heterogeneity in the subgroups of HPLC and IMx assay, indicating that quantification method of Hcy contributed to most of the observed heterogeneity. In addition, this heterogeneity might partly be explained by the fact whether the included studies adjusted for B vitamins. Sensitivity analyses showed that exclusion of any single study did not substantially alter the primary overall RRs, which further confirmed in the direction and magnitude of the findings in the present study. There were no missing studies imputed in regions of the contour enhanced funnel plots. In the linear dose-response analysis for Alzheimer-type dementia, Egger's and Begg's tests suggested no evidence of publication bias ( $p > 0.05$ ).

#### 4.2. Comparisons with other studies

The first meta-analysis on this topic was published in 2009 (Van Dam and Van Gool, 2009), and included only 3 cohort studies. Two years later, Wald et al. (Wald et al., 2011) reported the relationship between serum Hcy and dementia risk. Although this meta-analysis conducted by Wald has included 8 cohort studies with 8669 participants, it should be noticed that it was obviously different with the current study. First, the authors included a study (Haan et al., 2007) in which the cases were mixed and not the dementia-only patients (excluded by us, as above described). Thus, this probably affected, to a certain degree, the precision of the pooled results. Secondly, the authors did not distinguish Alzheimer disease from other dementia, and included the data containing the number of all-cause dementia cases. Our meta-analysis of cohort studies for AD aimed to explore the relationship of Alzheimer with Hcy, and included simply AD cases meeting the either NINCDS or DSM criteria. In addition, sensitivity analysis and publication bias test were not performed in this meta-analysis, which potentially decreased the stability of the pooled results (Wald et al., 2011). A report (Beydoun et al., 2014) included 5 studies in AD patients and performed a pooled analysis only in high versus low categories of blood Hcy levels. Another meta-analysis study (Nie et al., 2014) reported the association of HHcy with risk of cognitive decline. In fact, however, the cases from the study contained both dementia and cognitive decline. In other words, the authors defined abnormal cognitive function as cognitive decline, and included dementia and cognitive decline no dementia.

More comprehensively, our current meta-analysis examined the relationships between blood Hcy levels and risks of cognitive disorders including all-cause dementia, Alzheimer-type dementia, CIND, as well as vascular dementia, providing greater statistical power and more precise estimates because of pooling of multiple studies. It was also worth noting that the current dose-response analysis was performed in a broad range of, simple but not composite, cognitive outcomes, from cognitive decline no dementia to dementia. Importantly, there was a clear linear, but no curvilinear dose-response relationship between blood Hcy levels and risk of incident Alzheimer-type dementia. Our



**Fig. 3.** Blood Hcy and risk of Alzheimer-type dementia, linear and nonlinear dose-response analysis. A, linear dose-response analysis (per 5  $\mu\text{mol/L}$  increase); B, non-linear dose-response analysis. The solid line represents the best fitting cubic spline model. The area between two dash lines represents the 95% CI.

results support the notion that HHcy is a linear risk factor for AD.

#### 4.3. Implications

Our finding has important implications for prevention and treatment of AD. It was reported that the overall pooled prevalence of HHcy in China was 27.5%, particularly in northern populations, the inlanders, males and the elderly (Yang et al., 2014). The results of the current study imply that the increased incidence of AD might be attributed to the rapid elevation in blood Hcy. Clinical trial studies found that lowering Hcy levels with folic acid and B-vitamins could interfere with cognitive decline and AD (Cacciapuoti, 2013; Chen et al., 2016; Rommer et al., 2016). A report showed that B-vitamin supplementation could slow the atrophy of specific brain regions associated with AD process (Douaud et al., 2013). Overall, high levels of Hcy can be reduced or even reversed through changes in nutrition, and efforts toward early detection of HHcy in conjunction with implementation of lifestyle changes to improve Hcy metabolism may represent a viable strategy to reduce the risk of incident Alzheimer-type dementia, which is strongly supported by our results.

#### 4.4. Strengths and limitations

Several strengths of the current study include the comprehensive analyses of blood Hcy in relation to a range of mild-to-severe cognitive impairment risks; linear and non-linear dose-response analyses; the detailed subgroups, sensitivity and influence analyses; a large number of cases and participants; a complete quality assessment, and large populations. This meta-analysis was based on some prospective cohort studies from various populations. The sample size was large and the follow-up period was long enough. Compared with previous meta-analyses on this topic, the current dose-response meta-analysis covered broader classifications of cognitive outcomes. The estimates from the

fully adjusted models for each study were used in our analyses to reduce the potential of confounding. This can help to quantify the associations and test the shape of these possible associations.

Despite these strengths, our study also has some limitations. Firstly, as a meta-analysis of observational studies, there was the possibility of remaining residual confounding due either to known but unmeasured or imperfectly measured risk factors, or to factors that are not yet known to influence cognitive function risk. Secondly, most of the studies included have found that raised Hcy remained associated with cognitive impairment even after adjusting for B vitamins (Kivipelto et al., 2009; Ravaglia et al., 2005; Smith and Refsum, 2016; Whalley et al., 2014; Zylberstein et al., 2011), suggesting being independent of B vitamins for the association of cognitive impairment risk with blood Hcy levels. During follow-up, however, participants in the included studies might take in B vitamins, by means of either direct use or alterations induced by diet, resulting in changes in blood Hcy levels. In this sense, it might not be well-controlled for the potential confounder (B vitamins). Thirdly, differential adjustment for confounders across different studies could potentially influence our study findings. However, this was not observed in pooled analyses using HR associated with models with versus without adjustment for risk factors. Fourthly, our meta-analysis was conducted with summary statistics, rather than individual data which allow more precise delineation of the dose-response relation and further control of potential residual confounding. Lastly, as a meta-analysis of published literature, publication bias may have affected our findings. There was no evidence of publication bias in the analyses for Alzheimer-type dementia, but there was some indication of missing negative studies.

#### 4.5. Conclusions

Our dose-response meta-analysis shows that every 5  $\mu\text{mol/L}$  increase in blood Hcy is linearly associated with a 15% increased risk for

**Table 4**  
Stratified meta-Analyses of blood homocysteine levels and risk of cognitive disorders.

Category	All-cause dementia			Alzheimer-type dementia			CIND		
	n	RR (95% CI)	I <sup>2</sup> , %	n	RR (95% CI)	I <sup>2</sup> , %	n	RR (95% CI)	I <sup>2</sup> , %
<b>Location</b>									
Europe	6	1.67 (1.38 to 2.07)	30.4	5	2.04 (1.25 to 3.34) <sup>*</sup>	72	4	1.73 (1.00 to 2.99) <sup>*</sup>	68.9
America	2	1.67 (0.99 to 2.79)	0	2	1.53 (1.19 to 1.96)	0	4	1.12 (0.89 to 1.41)	6.4
Australia-Asia	2	1.34 (0.93 to 1.95)	49.4	1	3.31 (0.82 to 17.13)	0			
<b>Mean follow-up duration</b>									
Below 5 years	3	1.97 (1.39 to 2.79)	40.6	4	1.60 (1.05 to 2.45) <sup>*</sup>	65.1	2	1.54 (0.60 to 3.96) <sup>*</sup>	79.5
5 years or above	7	1.52 (1.26 to 1.82)	0	4	1.79 (1.40 to 2.28)	0	6	1.32 (0.98 to 1.78) <sup>*</sup>	50.8
<b>Female (%)</b>									
50 or above	7	1.64 (1.37 to 1.97)	4	7	1.7 (1.28 to 2.25)	61.6	7	1.15 (0.97 to 1.38)	31.1
50 below	3	1.87 (0.93 to 3.74) <sup>*</sup>	53.4	1	3.31 (0.82 to 17.13)	0	1	2.34 (1.39 to 3.91)	0
<b>Diagnosis criteria</b>									
ICD	2	1.80 (0.62 to 5.19) <sup>*</sup>	64.6	0			1	2.34 (1.39 to 3.91)	0
DSM-III-R/IV	7	1.90 (1.50 to 2.40)	0	3	2.56 (1.55 to 4.22)	0	3	1.01 (0.79 to 1.30)	0
NINCDS-AIREN/ADRD	1	1.40 (1.10 to 1.90)	0	5	1.54 (1.15 to 2.06) <sup>*</sup>	62.5			
MMSE							3	1.53 (0.79 to 2.95) <sup>*</sup>	61.9
other							1	1.53 (1.00 to 2.40)	0
<b>Sample source</b>									
plasma	9	1.59 (1.34 to 1.90)	26.3	6	1.73 (1.40 to 2.15)	22.5	7	1.42 (1.05 to 1.94) <sup>*</sup>	55.9
serum	1	1.67 (1.10 to 2.57)	0	2	1.58 (0.80 to 3.12) <sup>*</sup>	76.2	1	1.01 (0.70 to 1.44)	0
<b>Quantification method of Hcy</b>									
HPLC	4	1.42 (1.14 to 1.77)	30.7	3	1.56 (1.22 to 1.99)	0	7	1.32 (1.00 to 1.75) <sup>*</sup>	59.4
IMx assay	4	1.92 (1.46 to 2.53)	0	4	2.46 (1.70 to 3.56)	0	1	1.69 (0.63 to 5.25)	0
<b>Adjustment for confounding factors</b>									
<b>Age</b>									
Yes	9	1.58 (1.34 to 1.86)	12.5	8	1.74 (1.32 to 2.29) <sup>*</sup>	58.5	6	1.14 (0.95 to 1.37)	39
No	1	3.72 (1.06 to 13.08)	0	0			2	2.20 (1.38 to 3.50)	0
<b>Sex</b>									
Yes	6	1.60 (1.35 to 1.89)	31.5	8	1.74 (1.32 to 2.29) <sup>*</sup>	58.5	8	1.34 (1.02 to 1.74) <sup>*</sup>	53.7
No	4	1.67 (0.99 to 2.79)	0	0					
<b>BMI</b>									
Yes	4	1.81 (1.37 to 2.39)	0		2.56 (1.55 to 4.22)	0	2	1.27 (1.21 to 4.27)	0
No	6	1.51 (1.24 to 1.84)	46		1.54 (1.15 to 2.06) <sup>*</sup>	62.5	6	1.23 (0.95 to 1.61) <sup>*</sup>	54.1
<b>APOE</b>									
Yes	6	1.64 (1.34 to 2.01)	23.7	7	1.65 (1.24 to 2.19) <sup>*</sup>	57.3	5	1.29 (0.89 to 1.85) <sup>*</sup>	56.7
No	4	1.54 (1.17 to 2.01)	28.3	1	2.43 (1.25 to 4.71)	0	3	1.46 (0.90 to 2.36) <sup>*</sup>	64.6
<b>Education</b>									
Yes	9	1.70 (1.42 to 2.02)	6.1	8	1.74 (1.32 to 2.29) <sup>*</sup>	58.5	4	1.73 (1.00 to 2.99) <sup>*</sup>	68.6
No	1	1.20 (0.80 to 1.79)	0				4	1.12 (0.90 to 1.39)	6.4
<b>B vitamins</b>									
Yes	5	1.98 (1.52 to 2.59)	0	4	2.46 (1.70 to 3.56)	0	1	2.67 (1.22 to 5.85)	0
No	5	1.42 (1.16 to 1.74)	0	4	1.29 (1.13 to 1.47)	36.6	7	1.20 (1.01 to 1.42)	46.9
<b>History of cardiovascular diseases</b>									
Yes	6	1.78 (1.29 to 2.45) <sup>*</sup>	50.2	6	1.64 (1.23 to 2.18) <sup>*</sup>	63.5	3	2.31 (1.55 to 3.45)	0
No	4	1.69 (1.25 to 2.28)	0	2	2.74 (1.27 to 5.89)	0	5	1.09 (0.90 to 1.31)	0
<b>NOS score</b>									
Less than 7 stars				1	1.30 (0.80 to 2.30)	0	3	1.65 (1.19 to 2.27)	36.5
7 stars	6	1.56 (1.28 to 1.90)	0	5	1.64 (1.19 to 2.27) <sup>*</sup>	60.8	4	1.17 (0.91 to 1.51)	0
8 stars	4	2.01 (1.20 to 3.38) <sup>*</sup>	63	2	2.44 (1.45 to 4.10)	47.1	1	2.80 (1.20 to 6.20)	

RR = relative risk; CI = confidence interval; HPLC = high-performance liquid chromatography assay; ICD = the International International Classification of Diseases; DSM-III = the Diagnostic and Statistical Manual of Mental Disorders 3rd edition; NINCDSADRDA = National Institute of Neurologic and Cognitive Disorders and Stroke-AD and Related Disorders Association criteria; NINCDS-AIREN = National Institute of Neurological and Communicative Disease and Stroke-Alzheimer's Disease and Related Disorders Association; BMI = body mass index; APOE = apolipoprotein E; MMSE = Mini-mental State Examination; NOS = Newcastle-Ottawa Scale.

\* indicates the pooled result using a random effects model.

Alzheimer-type dementia. Hyperhomocysteinemia is also a risk factor for cognitive disorders, including not only dementia (non-Alzheimer type) but also cognitive impairment/decline. More prospective cohort studies, with large numbers of participants, especially those from developing countries, are needed to provide a more precise assessment of the effects of blood Hcy on non-Alzheimer type dementia and cognitive impairment.

#### Conflict of interest

The authors declare no conflict of interest.

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#### Appendix A. Supplementary data

Supplementary material related to this article can be found, in the online version, at doi:<https://doi.org/10.1016/j.arr.2019.02.006>.

## References

- Bekkering, G.E., Harris, R.J., Thomas, S., Mayer, A.M., Beynon, R., Ness, A.R., Harbord, R.M., Bain, C., Smith, G.D., Sterne, J.A., 2008. How much of the data published in observational studies of the association between diet and prostate or bladder cancer is usable for meta-analysis? *Am. J. Epidemiol.* 167, 1017–1026. <https://doi.org/10.1093/aje/kwn005>.
- Beydoun, M.A., Beydoun, H.A., Gamaldo, A.A., Teel, A., Zonderman, A.B., Wang, Y., 2014. Epidemiologic studies of modifiable factors associated with cognition and dementia: systematic review and meta-analysis. *BMC Public Health* 14 (643). <https://doi.org/10.1186/1471-2458-14-643>.
- Cacciapuoti, F., 2013. Lowering homocysteine levels with folic acid and B-vitamins do not reduce early atherosclerosis, but could interfere with cognitive decline and Alzheimer's disease. *J. Thromb. Thrombolysis* 36, 258–262. <https://doi.org/10.1007/s12399-012-0856-x>.
- Chen, H., Liu, S., Ji, L., Wu, T., Ji, Y., Zhou, Y., Zheng, M., Zhang, M., Xu, W., Huang, G., 2016. Folic acid supplementation mitigates Alzheimer's disease by reducing inflammation: a randomized controlled trial. *Mediators Inflamm.* 2016, 5912146. <https://doi.org/10.1155/2016/5912146>.
- Clarke, R., Smith, A.D., Jobst, K.A., Refsum, H., Sutton, L., Ueland, P.M., 1998. Folate, vitamin B12, and serum total homocysteine levels in confirmed Alzheimer disease. *Arch. Neurol.* 55, 1449–1455.
- Dolgin, E., 2016. How to defeat dementia. *Nature* 539, 156–158. <https://doi.org/10.1038/539156a>.
- Douaud, G., Refsum, H., de Jager, C.A., Jacoby, R., Nichols, T.E., Smith, S.M., Smith, A.D., 2013. Preventing Alzheimer's disease-related gray matter atrophy by B-vitamin treatment. *Proc. Natl. Acad. Sci. U. S. A.* 110, 9523–9528. <https://doi.org/10.1073/pnas.1301816110>.
- Dufouil, C., Alperovitch, A., Ducros, V., Tzourio, C., 2003. Homocysteine, white matter hyperintensities, and cognition in healthy elderly people. *Ann. Neurol.* 53, 214–221. <https://doi.org/10.1002/ana.10440>.
- Elliott, J.O., Jacobson, M.P., Haneef, Z., 2007. Cardiovascular risk factors and homocysteine in epilepsy. *Epilepsy Res.* 76, 113–123. <https://doi.org/10.1016/j.epilepsyres.2007.07.005>.
- Ford, A.H., Flicker, L., Alfonso, H., Hankey, G.J., Norman, P.E., van Bockxmeer, F.M., Almeida, O.P., 2012. Plasma homocysteine and MTHFR C677T polymorphism as risk factors for incident dementia. *J. Neurol. Neurosurg. Psychiatry* 83, 70–75. <https://doi.org/10.1136/jnnp.2011.242446>.
- Greenland, S., Longnecker, M.P., 1992. Methods for trend estimation from summarized dose-response data, with applications to meta-analysis. *Am. J. Epidemiol.* 135, 1301–1309. <https://doi.org/10.1093/oxfordjournals.aje.a116237>.
- Haan, M.N., Miller, J.W., Aiello, A.E., Whitmer, R.A., Jagust, W.J., Mungas, D.M., Allen, L.H., Green, R., 2007. Homocysteine, B vitamins, and the incidence of dementia and cognitive impairment: results from the Sacramento Area Latino Study on Aging. *Am. J. Clin. Nutr.* 85, 511–517. <https://doi.org/10.1093/ajcn/85.2.511>.
- Hendrie, H.C., Baiyewu, O., Lane, K.A., Purnell, C., Gao, S., Hake, A., Ogunniyi, A., Gureje, O., Unverzagt, F.W., Murrell, J., Deeg, M.A., Hall, K., 2013. Homocysteine levels and dementia risk in Yoruba and African Americans. *Int. Psychogeriatr.* 25, 1859–1866. <https://doi.org/10.1017/S1041610213001294>.
- Hooshmand, B., Solomon, A., Kareholt, I., Leiviska, J., Rusanen, M., Ahtiluoto, S., Winblad, B., Laatikainen, T., Soininen, H., Kivipelto, M., 2010. Homocysteine and holotranscobalamin and the risk of Alzheimer disease: a longitudinal study. *Neurology* 75, 1408–1414. <https://doi.org/10.1212/WNL.0b013e3181f88162>.
- Kado, D.M., Karlamangla, A.S., Huang, M.H., Troen, A., Rowe, J.W., Selhub, J., Seeman, T.E., 2005. Homocysteine versus the vitamins folate, B6, and B12 as predictors of cognitive function and decline in older high-functioning adults: MacArthur Studies of Successful Aging. *Am. J. Med.* 118, 161–167. <https://doi.org/10.1016/j.amjmed.2004.08.019>.
- Kalmijn, S., Launer, L.J., Lindemans, J., Bots, M.L., Hofman, A., Breteler, M.M., 1999. Total homocysteine and cognitive decline in a community-based sample of elderly subjects: the Rotterdam Study. *Am. J. Epidemiol.* 150, 283–289. <https://doi.org/10.1023/A:1022137429115>.
- Kivipelto, M., Annerbo, S., Hultdin, J., Backman, L., Viitanen, M., Fratiglioni, L., Lokk, J., 2009. Homocysteine and holo-transcobalamin and the risk of dementia and Alzheimer's disease: a prospective study. *Eur. J. Neurol.* 16, 808–813. <https://doi.org/10.1111/j.1468-1331.2009.02590.x>.
- Lehotsky, J., Tothova, B., Kovalska, M., Dobrota, D., Benova, A., Kalenska, D., Kaplan, P., 2016. Role of homocysteine in the ischemic stroke and development of ischemic tolerance. *Front. Neurosci.* 10, 538. <https://doi.org/10.3389/fnins.2016.00538>.
- Luchsinger, J.A., Tang, M.X., Shea, S., Miller, J., Green, R., Mayeux, R., 2004. Plasma homocysteine levels and risk of Alzheimer disease. *Neurology* 62, 1972–1976. <https://doi.org/10.1212/01.WNL.0000129504.60409.88>.
- McCaddon, A., Davies, G., Hudson, P., Tandy, S., Cattell, H., 1998. Total serum homocysteine in senile dementia of Alzheimer type. *Int. J. Geriatr. Psychiatry* 13, 235–239. [https://doi.org/10.1002/\(sici\)1099-1166\(199804\)13:4<235::aid-gps761>3.0.co;2-8](https://doi.org/10.1002/(sici)1099-1166(199804)13:4<235::aid-gps761>3.0.co;2-8).
- Mendonça, N., Granic, A., Mathers, J.C., Martin-Ruiz, C., Wesnes, K.A., Seal, C.J., Jagger, C., Hill, T.R., 2017. One-carbon metabolism biomarkers and cognitive decline in the very old: the newcastle 85+ study. *J. Am. Med. Dir. Assoc.* 18, 806–819. <https://doi.org/10.1016/j.jamda.2017.05.008>.
- Miwa, K., Tanaka, M., Okazaki, S., Yagita, Y., Sakaguchi, M., Mochizuki, H., Kitagawa, K., 2016. Increased total homocysteine levels predict the risk of incident dementia independent of cerebral small-vessel diseases and vascular risk factors. *J. Alzheimers Dis.* 49, 503–513. <https://doi.org/10.3233/JAD-150458>.
- Nie, T., Lu, T., Xie, L., Huang, P., Lu, Y., Jiang, M., 2014. Hyperhomocysteinemia and risk of cognitive decline: a meta-analysis of prospective cohort studies. *Eur. Neurol.* 72, 241–248. <https://doi.org/10.1159/000363054>.
- Nurk, E., Refsum, H., Tell, G.S., Engedal, K., Vollset, S.E., Ueland, P.M., Nygaard, H.A., Smith, A.D., 2005. Plasma total homocysteine and memory in the elderly: the Hordaland Homocysteine Study. *Ann. Neurol.* 58, 847–857. <https://doi.org/10.1002/ana.20645>.
- Orsini, N., Li, R., Wolk, A., Khudyakov, P., Spiegelman, D., 2012. Meta-analysis for linear and nonlinear dose-response relations: examples, an evaluation of approximations, and software. *Am. J. Epidemiol.* 175, 66–73. <https://doi.org/10.1093/aje/kwr265>.
- Peters, J.L., Sutton, A.J., Jones, D.R., Abrams, K.R., Rushton, L., 2008. Contour-enhanced meta-analysis funnel plots help distinguish publication bias from other causes of asymmetry. *J. Clin. Epidemiol.* 61, 991–996. <https://doi.org/10.1016/j.jclinepi.2007.11.010>.
- Plassman, B.L., Langa, K.M., Fisher, G.G., Heeringa, S.G., Weir, D.R., Ofstedal, M.B., Burke, J.R., Hurd, M.D., Potter, G.G., Rodgers, W.L., Steffens, D.C., McArdle, J.J., Willis, R.J., Wallace, R.B., 2008. Prevalence of cognitive impairment without dementia in the United States. *Ann. Intern. Med.* 148, 427–434. <https://doi.org/10.7326/0003-4819-148-6-200803180-00005>.
- Ravaglia, G., Forti, P., Maioli, F., Martelli, M., Servadei, L., Brunetti, N., Porcellini, E., Licastro, F., 2005. Homocysteine and folate as risk factors for dementia and Alzheimer disease. *Am. J. Clin. Nutr.* 82, 636–643. <https://doi.org/10.1093/ajcn.82.3.636>.
- Ravaglia, G., Forti, P., Maioli, F., Chiappelli, M., Montesi, F., Tumini, E., Mariani, E., Licastro, F., Patterson, C., 2007. Blood inflammatory markers and risk of dementia: the Conselice Study of Brain Aging. *Neurobiol. Aging* 28, 1810–1820. <https://doi.org/10.1016/j.neurobiolaging.2006.08.012>.
- Reitz, C., Tang, M.X., Miller, J., Green, R., Luchsinger, J.A., 2009. Plasma homocysteine and risk of mild cognitive impairment. *Dement. Geriatr. Cogn. Disord.* 27, 11–17. <https://doi.org/10.1159/000182421>.
- Rommer, P.S., Fuchs, D., Leblhuber, F., Schroth, R., Greilberger, M., Tafeit, E., Greilberger, J., 2016. Lowered levels of carbonyl proteins after vitamin B supplementation in patients with mild cognitive impairment and Alzheimer's disease. *Neurodegener. Dis.* 16, 284–289. <https://doi.org/10.1159/000441565>.
- Scarmeas, N., Anastasiou, C.A., Yannakoulia, M., 2018. Nutrition and prevention of cognitive impairment. *Lancet Neurol.* 17, 1006–1015. [https://doi.org/10.1016/S1474-4422\(18\)30338-7](https://doi.org/10.1016/S1474-4422(18)30338-7).
- Seshadri, S., Beiser, A., Selhub, J., Jacques, P.F., Rosenberg, I.H., D'Agostino, R.B., Wilson, P.W., Wolf, P.A., 2002. Plasma homocysteine as a risk factor for dementia and Alzheimer's disease. *N. Engl. J. Med.* 346, 476–483. <https://doi.org/10.1056/NEJMoa011613>.
- Smith, A.D., Refsum, H., 2016. Homocysteine, B vitamins, and cognitive impairment. *Annu. Rev. Nutr.* 36, 211–239. <https://doi.org/10.1146/annurev-nutr-071715-050947>.
- Stroup, D.F., Berlin, J.A., Morton, S.C., Olkin, I., Williamson, G.D., Rennie, D., Moher, D., Becker, B.J., Sipe, T.A., Thacker, S.B., 2000. Meta-analysis of observational studies in epidemiology: a proposal for reporting. Meta-analysis of Observational Studies in Epidemiology (MOOSE) group. *JAMA* 283, 2008–2012. <https://doi.org/10.1001/jama.283.15.2008>.
- Van Dam, F., Van Gool, W.A., 2009. Hyperhomocysteinemia and Alzheimer's disease: a systematic review. *Arch. Gerontol. Geriatr.* 48, 425–430. <https://doi.org/10.1016/j.archger.2008.03.009>.
- Wald, D.S., Kasteurirratne, A., Simmonds, M., 2011. Serum homocysteine and dementia: meta-analysis of eight cohort studies including 8669 participants. *Alzheimers Dement.* 7, 412–417. <https://doi.org/10.1016/j.jalz.2010.08.234>.
- Whalley, L.J., Duthie, S.J., Collins, A.R., Starr, J.M., Deary, I.J., Lemmon, H., Duthie, A.C., Murray, A.D., Staff, R.T., 2014. Homocysteine, antioxidant micronutrients and late onset dementia. *Eur. J. Nutr.* 53, 277–285. <https://doi.org/10.1007/s00394-013-0526-6>.
- Xu, C., Zeng, X.T., Liu, T.Z., Zhang, C., Yang, Z.H., Li, S., Chen, X.Y., 2015. Fruits and vegetables intake and risk of bladder cancer: a PRISMA-compliant systematic review and dose-response meta-analysis of prospective cohort studies. *Med. (Baltimore)* 94, e759. <https://doi.org/10.1097/MD.0000000000000759>.
- Yang, B., Fan, S., Zhi, X., Wang, Y., Wang, Y., Zheng, Q., Sun, G., 2014. Prevalence of hyperhomocysteinemia in China: a systematic review and meta-analysis. *Nutrients* 7, 74–90. <https://doi.org/10.3390/nu7010074>.
- Zhang, J., Yu, K.F., 1998. What's the relative risk? A method of correcting the odds ratio in cohort studies of common outcomes. *JAMA* 280, 1690–1691. <https://doi.org/10.1001/jama.280.19.1690>.
- Zhou, F., Haina, D., 2017. The bridging integrator 1 Gene rs7561528 polymorphism contributes to Alzheimer's disease susceptibility in East Asian and Caucasian populations. *Clin. Chim. Acta* 469, 13–21. <https://doi.org/10.1016/j.cca.2017.03.013>.
- Zylberstein, D.E., Lissner, L., Bjorklund, C., Mehlig, K., Thelle, D.S., Gustafson, D., Ostling, S., Waern, M., Guo, X., Skoog, I., 2011. Midlife homocysteine and late-life dementia in women. A prospective population study. *Neurobiol. Aging* 32, 380–386. <https://doi.org/10.1016/j.neurobiolaging.2009.02.024>.