



Spatiotemporal clustering of suicides in the US from 1999 to 2016: a spatial epidemiological approach

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Abstract

Purpose This study aims to describe and characterize the spatial and temporal clustering patterns of suicide in the ten states with the greatest suicide burden in the United States from 1999 to 2016.

Methods All suicide deaths from January 1, 1999 to December 31, 2016 in the United States were identified using data from the Wide-ranging Online Data for Epidemiologic Research (WONDER) dataset. The ten states with the highest age-adjusted suicide rates were Montana, Alaska, Wyoming, New Mexico, Nevada, Utah, Idaho, Colorado, Arizona, and Oklahoma. A spatiotemporal scan statistic using a discrete Poisson model was employed to retrospectively detect spatiotemporal suicide clusters.

Results From 1999 to 2016, a total of 649,843 suicides were recorded in the United States. Nineteen statistically significant spatiotemporal suicide mortality clusters were identified in the states with the greatest suicide rates, and 13.53% of the suicide cases within these states clustered spatiotemporally. The risk ratio of the clusters ranged from 1.45 to 3.64 ($p < 0.001$). All states had at least one cluster, with three clusters spanning multiple states, and four clusters were found in Arizona. While there was no clear secular trend in the average size of suicide clusters, the number of clusters increased from 1999 to 2016.

Conclusions Hot spots for suicidal behavior in the United States warrant public health intervention and continued surveillance. As suicide rates in the US continue to increase annually, public health efforts could be maximized by focusing on regions with substantial clustering.

Keywords Suicide · Suicide clusters · Epidemiology · Spatial statistics · Cluster analysis

Introduction

Suicide in the United States has increased by an average of 25.4% from 1999 to 2016 [1], and is currently the tenth leading cause of mortality overall and the second leading cause of mortality among youth and young adults [2]. Risk factors for suicide death include previous suicide attempts,

psychiatric and substance abuse disorders, as well as adverse and stressful life events and family history of psychiatric disorders and suicide [3, 4]. Further, those who complete suicide are more likely to be male, and use a lethal means of death such as a firearm [5]. Additional demographic risk factors include being non-Hispanic White or Native American [3], as well as being an adolescent or older adult [3]. It is also well established that suicide spatially and temporally correlates; however, relatively little has been done to examine suicide clusters across space and time in the United States during the past decade, a time period in which suicide has increased. Previous studies have shown that suicide clusters represent about 5–10% of suicides, and clustering has been found to be especially prevalent among adolescents and young adults [6–8].

Two main kinds of suicide clusters exist: mass and point clusters. More research has been conducted on mass clusters, which are suicide clusters that aggregate in time following, for example, media reports on suicide [9, 10]. Studies have

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demonstrated that suicide rates rise after increased media reporting on suicide, with a corresponding greater increase in suicide rates with increased coverage [11]. Recent evidence indicates an excess of 1841 suicides (9.85%) in the United States in the months following the death of Robin Williams [12]. On the other hand, point clusters, or spatiotemporal clusters that are localized in both geographic space and time, are less understood [13].

To our knowledge, no recent large-scale epidemiological study of spatiotemporal suicide clusters across multiple states or the national population in the United States has been conducted, despite an unprecedented suicide rate increase of over 25%. In the past 6 years, a few smaller studies using spatiotemporal cluster analysis methods have been carried out in the United States; however, all these studies examined clusters within one state. Additionally, most studies have examined spatial clusters, which are geographic areas with high or low suicide rates relative to other regions, as opposed to spatiotemporal clusters which are identified across both space and time [14–16].

Several social and psychological mechanisms have been suggested to contribute to the formation of spatiotemporal clusters of suicide. Environmental factors such as sunlight, seasonality, and altitude contribute to suicidal behavior and underlie some degree of spatial correlation [17–19]; further, clustering of socio-economic deprivation and marginalization [20–22], as well as greater local person-to-person transmission of suicidal thoughts and behavior also contribute [13, 23]. Local transmission can occur through imitation, suggestion, social learning, and assortative relating [23]. More sophisticated statistical methods and techniques, such as methods derived from geographical information systems and spatial epidemiology, may be able to detect significant clustering effects during surveillance. Such detection of suicide clusters is essential for effective prevention of suicides and may allow medical professionals and policymakers to identify potential high-risk areas and intervene to prevent suicide deaths and injuries from attempted suicide.

This study aims to fill in gaps in the current literature on spatiotemporal suicide clustering in the United States. First, we examined demographic and county-level differences in overall suicide mortality in the United States. Second, we examined spatiotemporal clustering from 2009 through 2016 within the ten states in the United States with the highest suicide burden. We focused on states with the highest burden to determine whether suicide cluster analysis would generate additional information that can inform suicide prevention and monitoring efforts. Third, we estimated temporal trends in the number of clusters identified in the United States, as well as the size of (i.e., number of cases within) the clusters. Lastly, we assessed whether there were any significant demographic differences between the counties included and not included in the suicide clusters.

Methods

Data source

Suicide count and age-adjusted suicide rates (from January 2009 to December 2016) were obtained from the Centers for Disease Control and Prevention Wide-ranging Online Data for Epidemiologic Research (CDC WONDER). The CDC WONDER Underlying Cause of Death database contains mortality and population counts for all US counties. The data are based on death certificates for US residents. Each death certificate identifies a single underlying cause of death and demographic data. During the study period, changes to county boundaries may have occurred, including the wholesale addition or deletion of counties. Additional information on the WONDER database is described elsewhere [24]. County-level demographic characteristics were obtained from the US Census Bureau 2010 Profile of General Population and Housing Characteristics (DP-1) [25]. The data are publicly available, de-identified, and were deemed to be without human subjects risk by the Columbia University institutional review board.

Study population and case ascertainment

All suicide deaths identified through WONDER database were included in the study. Suicides in WONDER were selected if they endorsed any one of the following ICD-10 codes: U03 Terrorism intentional—Suicide ($n=4$), X60–X84 Intentional self-harm ($n=648,635$), and Y87.0 Sequelae of intentional self-harm ($n=1204$). In WONDER, deaths in 1999 and beyond were classified using ICD-10; however, category U03 was added in 2001.

Statistical analysis

We began by estimating suicide rates by state across all 50 states. Age-adjusted mortality rates (number of deaths/100,000 persons) were calculated by age, gender, census division, state, and county for the years 1999–2016. We used the midyear 2000 US standard population to estimate age-adjusted rates and account for changes in the age distribution of the population over time. Descriptive statistics were used to characterize the sample demographics and describe trends and variations in suicides across different demographic groups.

We then focused the cluster analysis on the ten states with the highest age-adjusted suicide rates over the study period (Montana, Alaska, Wyoming, New Mexico, Nevada, Utah, Idaho, Colorado, Arizona, and Oklahoma). Spatiotemporal clusters of suicides in these ten states were retrospectively

assessed using cluster analysis with the spatiotemporal SaTScan v.9.4.4 program [26]. Due to WONDER data suppression for sparse cells, we used completed suicide information aggregated every 6 months, which was the most granular time scale without substantial county-level data suppression. The ten states with the highest age-adjusted suicide rate were selected for the cluster analysis, as this also limited the overall amount of data suppression: 65.3% of suicides were not suppressed and remained available for analysis. In contrast, this percentage dropped to 61.9% and 55.2% for the top 15 and 20 states, respectively. SaTScan is commonly used in infectious disease cluster analysis, but has also been previously used in the identification of suicide clusters [27, 28], typically in a specific geographic region or institutional setting within the US, or in population-level analysis in other countries [27, 29–31]. We define a suicide spatiotemporal cluster as a group of suicides that occur (based on place of residence) closer together in space and time than would be expected by chance.

A spatiotemporal scan statistic was obtained using a cylindrical scan window that is moved through space and time, and each window reflects a possible cluster. A discrete Poisson model was used, and the likelihood of each possible cluster was assessed using Monte Carlo simulations. A likelihood ratio test was used to compare the risk of suicides within the scan relative to outside the scan. Statistical significance was reached when a higher than expected risk within the scan window was found compared to the outside. Because the scan window uses longitude and latitude to assess for clusters in the spatial dimension, changes in county boundaries during the study period did not affect the cluster analysis.

Previous research has shown that the time period during which suicide clusters are expected to occur ranges from 24 h to 5 years [32]; our study focus was point clusters more localized in space and time, thus the maximum temporal cluster size was a priori set to 15% of the study period or around 2.5 years, and the minimum temporal cluster size was set to 12 months. The maximum proportion of the population that a cluster could contain was a priori set at 10% of the study population. We then compared the cluster analysis results with state-level, county-level, and group-specific descriptive statistics to evaluate which additional distinct information could be gleaned from the spatiotemporal cluster analysis. Subsequently, we linked these counties with county-level demographic information (age, sex, race, and ethnicity) from the US Census Bureau to determine whether associations exist between these characteristics and suicide clustering.

Further, we stratified the study period into 3-year increments to test whether there was a trend over time in the number and size of suicide clusters, while accounting for overall increasing suicide rates. Because the analysis was

conducted on 3-year segments, the maximum temporal cluster size was set to 1.5 years and the minimum temporal cluster size was set to 6 months. This stratified cluster analysis determines clusters relative to the cases and population size within the 3-year time period of interest and yields different clusters. Therefore, the total number of clusters in the overall time period (1999–2016) is not expected to equal the sum of the clusters within each 3-year period. Additionally, we identified purely spatial clusters in the data, and compared whether these spatial clusters overlapped with the spatiotemporal clusters.

All analyses, except the cluster analysis, were conducted in R version 3.4.1 [33]. The geocode function of the ggmap package of R determined each county's geocoordinates using Google Map APIs. The spatiotemporal cluster analysis was conducted in the spatiotemporal SaTScan v.9.4.4 program [26].

Results

Age-adjusted suicide death rates by gender, age group, and census division

In the United States, 649,843 deaths due to suicide were recorded between 1999 and 2016. The overall age-adjusted suicide rate was 11.74/100,000 persons (95% CI 11.71–11.77). The suicide rate (per 100,000 persons) was highest for those in the 85+ age group (18.07, 95% CI 17.79–18.34), males (19.26, 95% CI 19.20–19.31), and the Mountain State census division (18.06, 95% CI 17.92–18.19) (Table 1). Males were almost four times more likely than females to die from suicide (3.84, 95% CI 3.82–3.87). Furthermore, the rate of suicide mortality increased within the study period; the age-adjusted death rate during 2016 was 1.29 times (95% CI 1.27–1.31) the age-adjusted death rate in 1999.

County-level age-adjusted suicide death rates within age groups and states

When examined within the same age groups and states, there were variations in suicide rates among counties, particularly within the 15–24 and 25–34 age groups. The overall age-adjusted suicide mortality rate (per 100,000 persons) among the 15–24 age group was 10.58 (95% CI 10.51–10.65), which was the second lowest death rate among the age groups. However, a county in Alaska, one in South Dakota, and another in Alaska had age-adjusted death rates (per 100,000 persons) of 254.67 (95% CI 197.36–323.42), 206.08 (95% CI 132.04–306.63), and 196.14 (95% CI 147.76–255.31) among that age group, respectively. These mortality rates

Table 1 Characteristics of suicides in the United States, January 1999 to December 2016

Characteristic	1999–2004		2005–2010		2011–2016		Overall	
	No.	Age-adjusted rate	No.	Age-adjusted rate	No.	Age-adjusted rate	No.	Age-adjusted rate
Total	184,749	10.74	211,843	11.45	253,251	12.88	649,843	11.75
Age group								
5–14 years	1629	0.66	1436	0.59	2277	0.92	5337	0.72
15–24	24,180	9.98	25,810	9.96	30,876	11.74	80,866	10.58
25–34	30,153	12.70	31,608	13.15	39,551	15.25	101,312	13.75
35–44	39,754	14.84	39,814	15.66	40,589	16.69	120,157	15.70
45–54	36,155	15.31	47,879	18.13	52,303	19.95	136,337	17.88
55–64	20,630	13.13	31,519	15.62	43,622	18.30	95,771	16.03
65–74	14,272	12.89	15,859	13.08	23,059	14.95	53,190	13.77
75–84	13,084	17.25	12,617	16.09	14,180	17.35	39,881	16.90
85+	4826	18.50	5255	17.14	6763	18.53	16,844	18.07
Gender								
Female	36,823	4.21	44,543	4.70	56,530	5.66	137,769	4.98
Male	147,926	18.08	167,300	18.83	196,721	20.64	511,947	19.16
Census division								
New England	7023	8.11	8138	9.03	9900	10.60	25,061	9.28
Middle Atlantic	18,612	7.62	20,900	8.27	25,328	9.68	64,840	8.56
East North Central	27,631	10.12	30,896	10.97	35,838	12.49	94,365	11.21
West North Central	13,148	11.26	15,055	12.34	18,495	14.60	46,698	12.78
South Atlantic	37,430	11.54	42,893	11.88	51,020	13.00	131,343	12.20
East South Central	12,803	12.36	15,017	13.64	17,232	14.93	45,052	13.69
West South Central	21,392	11.39	24,385	11.78	30,423	13.33	76,200	12.24
Mountain	18,382	16.67	22,158	17.51	27,359	19.64	67,899	18.05
Pacific	28,222	10.46	32,305	10.93	37,625	11.73	98,152	11.09

were 24.07, 19.48, and 18.54 times the overall age-adjusted death rate among the 15–24 age group.

Furthermore, increased suicide mortality rates in certain counties were observed when county-level rates were examined by state, particularly in Alaska, South Dakota, and North Dakota. Figure 1 presents box and whiskers plots of county age-adjusted suicide rates stratified by state, where counties with suicide rates greater than 1.5 times the interquartile range are outliers (marked as dots). These outliers represent counties with much greater or less than average age-adjusted suicide rates throughout the study period. The age-adjusted death rate (per 100,000 persons) in a county in Alaska (89.86, 95% CI 73.35–106.36), a county in South Dakota (77.83, 95% CI 50.84–114.05), a second county in Alaska (69.52, 95% CI 56.75–82.30), a third county in Alaska (60.08, 95% CI 47.79–74.58), and a county in North Dakota (56.96, 95% CI 41.22–76.72) were 7.65, 6.63, 5.92, 5.12, and 4.85 times the overall age-adjusted death rate of 11.74 (95% CI 11.71–11.77) in the study sample (Fig. 1). There was less variation when county death rates were stratified by year and gender; there were no counties with mortality rates significantly higher relative to the stratified death rate by year and gender group.

Cluster analysis

The Mountain Census Division had the greatest overall age-adjusted suicide rate during 1999–2016, and eight out of the ten states with the highest rate of suicides were included in this Census Division. The states with the greatest overall age-adjusted suicide mortality rates (per 100,000 persons) for all the study years were Montana (21.26, 95% CI 20.57–21.95), Alaska (21.65, 95% CI 20.80–22.49), Wyoming (21.49, 95% CI 20.56–22.42), New Mexico (19.95, 95% CI 19.48–20.42), Nevada (19.32, 95% CI 18.92–19.73), Utah (18.02, 95% CI 17.61–18.43), Idaho (17.73, 95% CI 17.22–18.25), Colorado (17.39, 95% CI 17.11–17.67), Arizona (16.70, 95% CI 16.46–16.95), and Oklahoma (16.33, 95% CI 16.01–16.64). Figure 2 shows the age-adjusted suicide rates by year among these states from 1999 to 2016.

Spatiotemporal analysis was conducted among counties within the states with the highest age-adjusted suicide rates. We found 19 statistically significant spatiotemporal clusters. The rate of suicide in the counties where the clusters occurred was up to 3.64 times higher than in areas outside the clusters, and the risk ratio ranged from 1.45 to 3.64 ($p < 0.001$). Four clusters were found in Arizona, three each

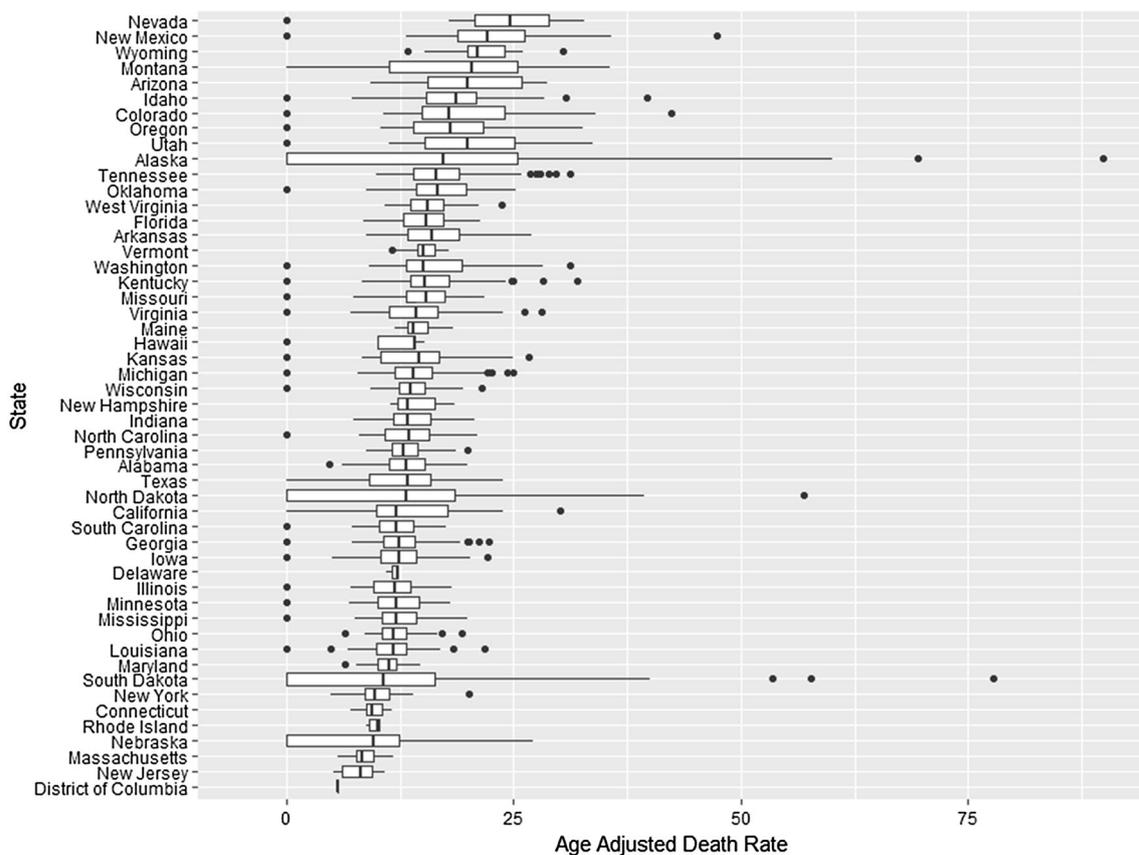


Fig. 1 County-level age-adjusted suicide death rates (per 100,000 persons) by state

in Idaho, Colorado, and Oklahoma, two in Nevada, Wyoming, Utah, and Montana, and one each in New Mexico and Alaska. There were three clusters that spanned multiple states (Table 2).

In total, 7850 suicides (13.53%) within these ten states clustered spatiotemporally, with the lowest percentage of clustered deaths in Montana (2.85%) and the highest in New Mexico (50.50%). Additionally, 66 counties (17.05%) within these ten states were included in the space–time clusters, with the lowest proportion of counties in Alaska (2 out of 32 counties, 6.25%) and the highest in Arizona (11 out of 15 counties, 73.33%). Figure 3 shows the map of the spatiotemporal clusters. We also conducted cluster analysis to detect purely spatial clusters; 15 clusters were found: 2 spatial clusters each were found in Arizona, Colorado, Idaho, Montana, New Mexico, and Oklahoma, and 1 cluster was detected in Kansas, Nevada, and Utah. There were no spatial clusters found in Wyoming, and none of the identified clusters spanned multiple states (Table 3 in Appendix). There was substantial overlap between spatial and spatiotemporal clusters; of 35 counties included in the spatial clusters, only 4 were not in the spatiotemporal clusters. On the other hand, 16 counties were included in the spatiotemporal clusters but

not the spatial clusters. Figure 4 in Appendix provides a map of the clusters.

Association between county-level demographic data and suicide clusters

Significant associations were found between county-level demographic characteristics (age group, race, and ethnicity) and suicide clustering. The percentages of 15–24 years old (13.50% vs 12.37%, $p < 0.01$) and 25–34 years old (12.51% vs 11.75%, $p = 0.026$) among the cluster counties were greater than in the counties not included in the clusters. In contrast, the percentages of older adults 45–54 years old (14.38% vs 15.07%, $p < 0.01$), 75–84 (4.29% vs 4.78%, $p = 0.012$), and above 85 (1.53% vs 1.84%, $p < 0.0001$) within cluster counties were less than in remaining counties. For counties in suicide clusters compared to the non-cluster counties, significant differences were also noted in the percentage of Native Hawaiians and other Pacific Islander (19.22% vs. 10.26%, $p = 0.017$), other race (6.04% vs 4.52%, $p = 0.025$), and Hispanic ethnicity (16.68% vs. 12.59%, $p = 0.032$).

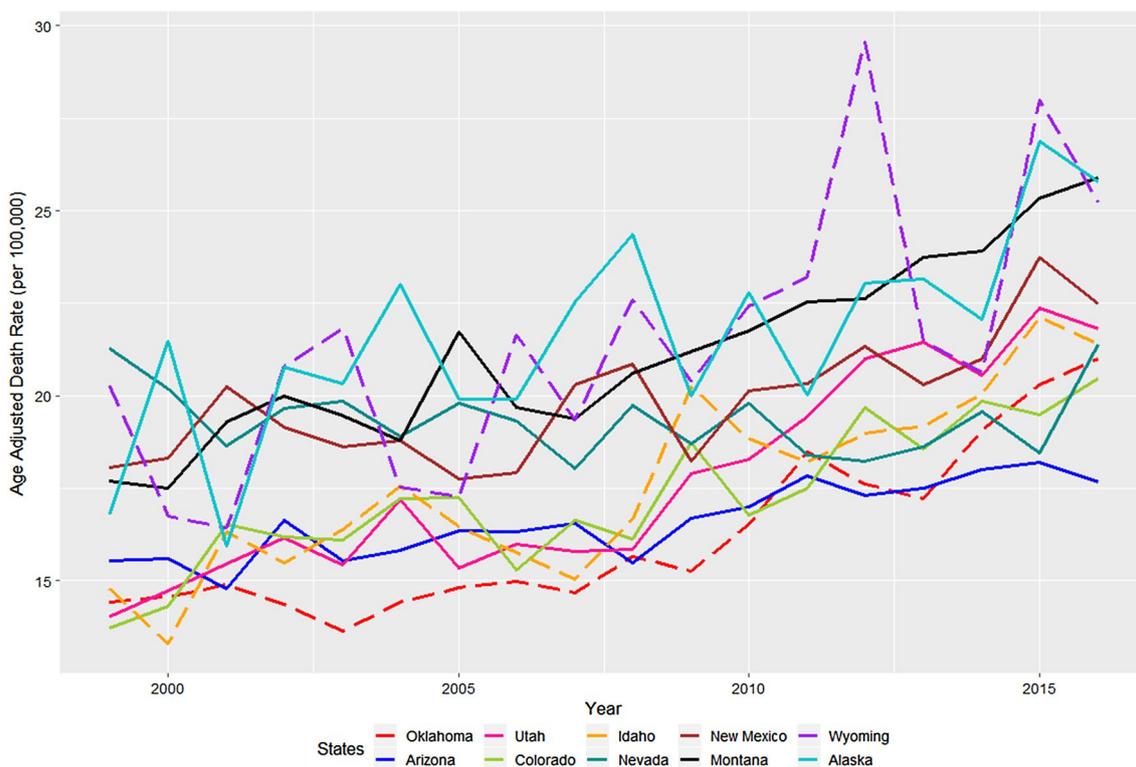


Fig. 2 Suicide rates over time in states with the greatest mortality burden

Table 2 Spatiotemporal suicide cluster characteristics

State	No. of counties	Observed cases	Expected cases	RR	p value	Dates	
1	Arizona, Nevada and Utah	6	1549	911.72	1.72	< 0.001	7/1/2014–12/31/2016
2	Colorado	11	1264	790.33	1.61	< 0.001	7/1/2014–12/31/2016
3	Utah	3	913	537.69	1.71	< 0.001	7/1/2014–12/31/2016
4	New Mexico and Arizona	14	761	487.17	1.57	< 0.001	7/1/2014–12/31/2016
5	Arizona	4	730	502.87	1.46	< 0.001	1/1/2014–6/30/2016
6	Wyoming and Colorado	8	537	350.04	1.54	< 0.001	7/1/2014–12/31/2016
7	Nevada	1	265	142.48	1.86	< 0.001	7/1/2014–12/31/2016
8	Colorado	1	118	46.6	2.54	< 0.001	1/1/2011–6/30/2013
9	Alaska	2	212	112.85	1.88	< 0.001	7/1/2014–12/31/2016
10	Oklahoma	1	321	201.37	1.6	< 0.001	1/1/2014–6/30/2016
11	Oklahoma	2	483	334.34	1.45	< 0.001	7/1/2014–12/31/2016
12	Idaho	2	255	164.7	1.55	< 0.001	1/1/2015–12/31/2016
13	Montana	5	122	68.68	1.78	< 0.001	7/1/2011–6/30/2013
14	Montana	1	95	49.95	1.9	< 0.001	7/1/2014–12/31/2016
15	Idaho	1	75	37.85	1.98	< 0.001	1/1/2014–12/31/2015
16	Wyoming	1	49	20.69	2.37	< 0.001	1/1/2014–12/31/2015
17	Oklahoma	1	23	6.31	3.64	< 0.001	7/1/2015–6/30/2016
18	Arizona	1	22	6.72	3.27	< 0.001	7/1/2014–6/30/2015
19	Idaho	1	56	28.03	2	< 0.001	7/1/2014–6/30/2016

Fig. 3 Graphical representation of 19 suicide spatiotemporal clusters within the 10 states with the greatest age-adjusted suicide rates. Each color represents a different cluster

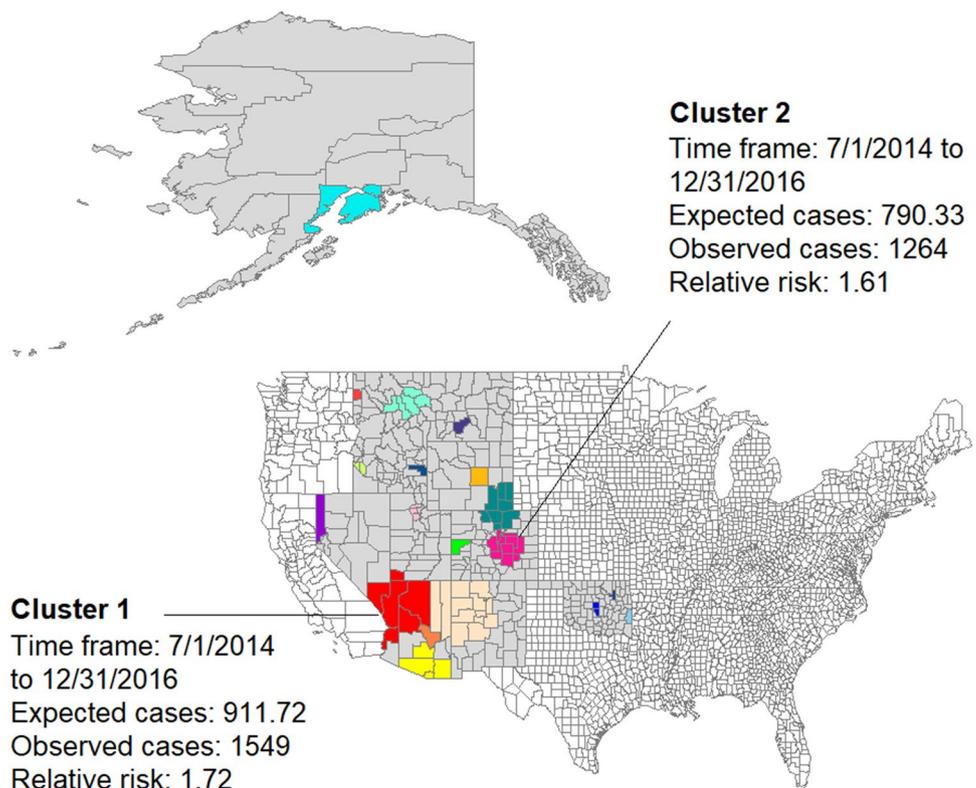


Table 3 Spatial suicide cluster characteristics

	State	No. of counties	Observed cases	Expected cases	RR	<i>p</i> value
1	Arizona	5	8389	5364.28	1.66	< 0.001
2	Colorado	10	6684	5136.49	1.34	< 0.001
3	Utah	3	4663	3440	1.39	< 0.001
4	New Mexico	3	2678	1799.08	1.51	< 0.001
5	Nevada	1	1531	918.26	1.69	< 0.001
6	Arizona	4	4254	3230.98	1.34	< 0.001
7	Colorado	1	575	311.99	1.85	< 0.001
8	Oklahoma	1	1825	1352.36	1.36	< 0.001
9	Oklahoma	1	2019	1618.68	1.26	< 0.001
10	Alaska	1	895	644.8	1.39	< 0.001
11	Idaho	1	1031	849.29	1.22	< 0.001
12	Idaho	1	408	299.77	1.36	< 0.001
13	Montana	1	426	326.72	1.31	< 0.001
14	Montana	1	321	241.37	1.33	< 0.001
15	New Mexico	1	533	450.48	1.18	< 0.001

Temporal trends in number and size of suicide clusters

We also stratified the study period into 3-year increments to test whether there was a trend over time in the number and size of suicide clusters, while accounting for the overall increasing suicide rate. The number of spatiotemporal

clusters varied across the study period; overall, the number of clusters observed across the study period increased by greater than 50% between the first and last 3 years of the study, 1999–2001 and 2014–2016. During the first 3 years of the study, 15 spatiotemporal clusters were found, with an average cluster size of 168.8 people. The number of clusters increased to 20, 21, 20, and 23 in the next five 3-year time

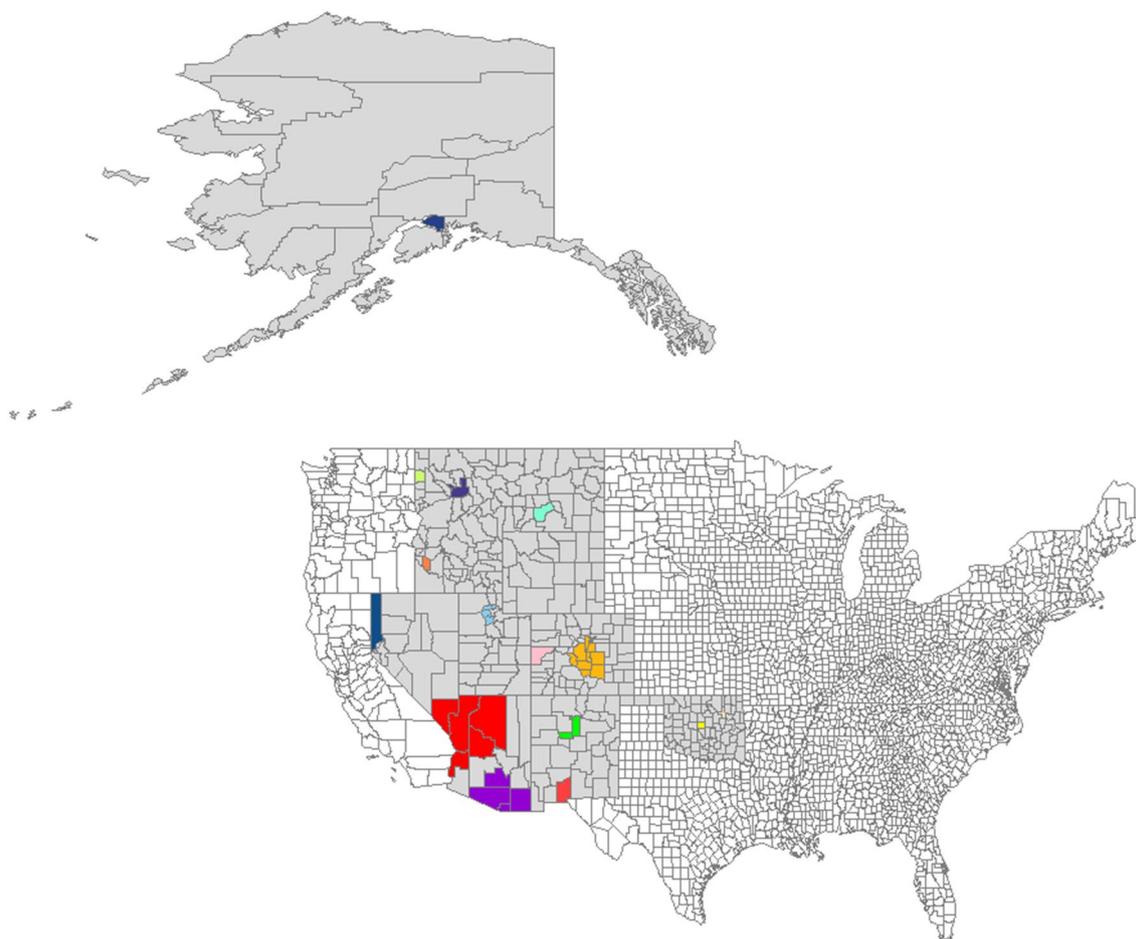


Fig. 4 Graphical representation of 15 suicide spatial clusters within the 10 states with the greatest age-adjusted suicide rate. Each color represents a different cluster

periods and to 24 during the final 3 years. On the other hand, there was no trend in the average size of identified suicide spatiotemporal clusters.

Discussion

We investigated the spatiotemporal distribution of suicides in the United States, with a focus on the ten states with the highest age-adjusted rate of suicides between 1999 and 2016. The results demonstrate that a considerable number of suicides cluster in space and time, and that spatiotemporal cluster analysis contributes valuable information to the surveillance of suicides beyond general state-level trends. Monthly state-level time series plots are unable to account for the spatiotemporal suicide clusters identified in cluster analysis. Based on our findings, suicide prevention efforts that focus on geographic areas with high suicide risk, in addition to individual risk factors, may achieve greater

public health benefits than focusing on individual risk factors alone.

Sub-state-level heterogeneity in suicide death rates and clustering was evident in the WONDER record. These deviations from state-level rates produced some interesting findings. For example, some counties with significantly higher suicide rates, such as several counties in Alaska, did not provide evidence of suicide clustering; whereas other counties with extremely high suicide rates were found in South Dakota and North Dakota, which are states with the 13th and 21st highest state-level suicide rates in the United States (Fig. 1). This finding adds evidence to the inadequacy of aggregated state-level, county-level, or group-specific rates to assess suicide patterns, and points to the need to identify clusters in space and time with few predetermined boundaries. Clusters may represent locations that have the greatest need for targeted interventions and suicide prevention strategies. Additional studies are needed to identify suicide clusters given the finding here that cluster numbers have increased since 1999. The increase in number of suicide

clusters in the United States demonstrates greater spatiotemporal suicide rate heterogeneity in recent years. This trend represents a greater than expected increase in suicide rates in specific counties and should be examined further in future studies.

Our results demonstrate that some county-level demographic features associated with suicide clusters are similar to previously reported individual-level demographic correlates for suicide. The positive associations of Native Hawaiians and other Pacific Islander [34] and other race [34] with suicide have been previously demonstrated; our results indicate that areas with cluster suicides have a higher concentration of individuals with such demographic risk factors. Additionally, previous research has reported disproportionately high rates of suicides among American Indians and Alaskan Natives [35, 36]; suicide clusters in the current study, however, do not have higher concentrations of American Indians and Alaskan Natives. Published research on Hispanic ethnicity and suicide in the United States are inconsistent [37], but our findings identify higher concentrations of Hispanic ethnicity in areas with suicide clusters. Relationships between age groups at risk for completed suicide versus counties with different concentrations of various older and younger adults are also heterogeneous. In sum, while some individual-level risk factors for suicide demonstrate similar aggregate relationships with suicide clusters, many do not. However, area-level factors are a different unit of analysis than individual-level factors, with different potential mechanisms for influencing suicide clusters versus individual risk of suicide. Further studies examining these multi-level ecological and individual-level effects simultaneously would allow for more rigorous disentanglement of individual and ecological risks.

There are several strengths in this study. To our knowledge, this is the first epidemiological study that has applied spatiotemporal scan statistics to county-level data from multiple states in the US in the last decade. The examination of recent suicide trends and clustering is extremely valuable given that suicide rates have increased considerably in the US, and half of states had an increase of more than 30% [1]. Consistent with previous research in similar large-scale studies in other countries, our findings show the potential of suicides to cluster in space and time. We also found considerable overlap between spatial and spatiotemporal clusters; in situations where longitudinal suicide data are unavailable, cross-sectional data could instead be used to detect spatial clusters amenable to geographically targeted interventions. An additional strength of the study is the comparison of the cluster analysis to descriptive statistics, which has not been examined before. This feature allowed explicit examination of the unique information each approach adds to the surveillance of suicide clusters. This study also illustrates

the potential of leveraging publicly available epidemiologic datasets that are conventionally not used for spatial analysis.

Spatiotemporal clusters were also found in recent studies in Australia [29, 30], Brazil [31], United Kingdom [27], and Spain [28]. The inherent propensity of suicides to cluster in both space and time adds further credibility to the notion that the clustering effect is not only due to area or group characteristics, but rather that temporally varying social and psychological mechanisms may underlie the formation of these spatiotemporal clusters. Indeed, by detecting spatiotemporal suicide clusters rather than purely spatial ones, the present study provides greater evidence for complex and poorly understood suicide cluster mechanisms [29]. Such mechanisms might suggest a transmission effect; however, it is important to note that the size of the clusters identified here are much larger than the typical cluster sizes associated with direct transmission of suicidal behaviors. This larger size may be due to the county-level spatial resolution of the WONDER data, which is likely not granular enough to characterize smaller clusters that are socially transmitted. The large geographic area of some counties in the cluster analysis could also have obscured smaller clusters, and future studies with smaller geographic units of analysis are needed. Moreover, both risk (media coverage) and protective (prevention program) factors associated with suicide clustering need to be further examined to link data at geographic scales with individual-level data on suicide decedents. The demographics and topology of the cluster areas identified here suggest a variety of other possible underlying mechanisms may also play a role, including deprivation and isolation, such as might occur in rural areas of Alaska, as well as potential environmental factors, such as altitude and weather. Public health efforts to reduce suicide in these high burden areas should therefore take account of the social and environmental dynamics of the local areas in which high suicide rates are observed.

One of the main limitations of the present study was the data suppression in the WONDER dataset. An aim of the study was to conduct cluster analysis at the county level and at a granular time scale; however, WONDER data from counties with less than 10 reported suicides within a given time interval are suppressed. Consequently, steps were taken to limit the effects of this data suppression. Six months was the most granular time scale employed, and the ten states with the highest age-adjusted suicide mortality rates were selected for analysis as data suppression was less substantial in these states. Moreover, our findings are only generalizable to the ten states included in the study, since clusters in high rate areas may be very different from those in low rate areas. Given that the identification of suicide clusters provides vital information to inform the allocation of resources for suicide prevention and monitoring efforts, future spatiotemporal

cluster analysis needs to be conducted in all states, regardless of suicide rates and without any data suppression.

Another possible limitation is the potential erroneous identification of cause of death. Depending on state law, death investigations are conducted by either a coroner system or a medical examiner system organized at the state, district, or county level. In coroner systems, coroners generally do not need to be physicians, whereas medical examiners are usually required to be physicians and often pathologists [38, 39]. States included in the cluster analysis employ different death investigation systems: Alaska, New Mexico, Utah, Oklahoma, and Montana utilize state medical examiners; in addition, Montana also has county coroners. In contrast, Nevada, Idaho, Wyoming, and Colorado have county-based coroners, and Arizona has county medical examiners [40]. The differences in these systems both at the state and county level could potentially introduce differential bias among states or counties, which could affect the results of the cluster analysis. To the extent that suicide deaths are generally undercounted or misclassified as unintentional deaths, we would anticipate that the number of clusters and cluster sizes would be greater than what is reported in our analysis. Lastly, another possible limitation is that our analysis only examines clustering in space and time, and we are only able to count suicides where the decedent resides, not other areas in which he/she interacts. We were not able to observe cases linked by anything other than time and space such as social networks (e.g., close loved one dying from a different state).

Given that suicides cluster in space and time, coupled with the overall increase in suicides in the US, there is great need for effective interventions to prevent suicide clustering. Previous research on suicide mass clusters found that media portrayals of suicide play a role in the formation of clusters [6], and this finding may also extend to spatiotemporal clusters. The media depiction of suicides—more specifically local sources in the case of spatiotemporal clusters—needs to adhere to appropriate guidelines to minimize subsequent suicides [10, 41]. Further, the advent of social media and new electronic media requires additional research to assess the impact of the transmission of information regarding completed suicides within one's social network [10, 42]. In the wake of a death by suicide of someone in the community or a suicide that receives media attention, additional resources need to be allocated to postvention efforts with vulnerable groups, as it has been suggested that this may contain the effects of suicide contagion [43, 44]. Further research needs to examine the risk factors that lead to the formation of suicide clusters, with the goal of implementing multifactorial suicide prevention and postvention strategies. A 2016 systematic review found that there are numerous effective methods for prevention of suicidal behavior, such as psychotherapy, pharmacological interventions, and community and family-based interventions. A tailored approach

using a combination of evidence-based strategies at both the individual and population level would be most effective for each risk group [45]. Additionally, screening for suicide risk is useful for identifying individuals with high susceptibility to suicide contagion [46].

Our work adds to the body of evidence demonstrating the tendency of suicides to cluster spatiotemporally. Studies utilizing aggregated state-level or group-specific statistics alone are inadequate for examining trends in suicides and miss an essential epidemiological component of suicidality that can assist in the identification of regions of increased suicidal risk. The identification of these clusters is crucial to perform targeted interventions and inform public policy decisions. Future studies need to integrate spatial analysis in the epidemiological investigation of suicides at granular spatial and temporal scales.

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Compliance with ethical standards

Conflict of interest JS and Columbia University disclose partial ownership of SK Analytics. SK has consulted for SK Analytics. The other authors report no conflicts of interest and have no financial relationships with commercial interests.

Appendix

See Table 3 and Fig. 4.

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